

05.11.2015 Exercises 29, 27, then exercise 23

$Y_{n \times 1} \sim N_n(X\beta, \sigma^2 V_0)$ *fixed, and known.*
 $A_{n \times (n-p)} \quad A^T X = 0, \quad p < n$
 $A^T Y \sim N_{n-p}(\vec{0}, \Sigma_{\text{new}})$
 $n \times (n-p) \times 1 \rightarrow n-p \times 1$

$E\{A^T Y\} = A^T E\{Y\} = A^T E\{X\beta\} = A^T X\beta = \vec{0}$
 $\text{COV}(A^T Y) = A^T \text{COV}(Y) A = A^T \sigma^2 V_0 A = \sigma^2 \underbrace{A^T V_0 A}_{V_1}$
 V_1 $n-p \times n-p$

$Z = A^T Y \sim N_{n-p}(\vec{0}, \sigma^2 V_1)$
 $f(A^T Y) = \frac{1}{(2\pi)^{\frac{n-p}{2}} |\sigma^2 V_1|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} (A^T Y - 0)^T (\sigma^2 V_1)^{-1} (A^T Y - 0)\right)$
 $= \frac{1}{(2\pi)^{\frac{n-p}{2}} |\sigma^2 V_1|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} Y^T A V_1^{-1} A^T Y\right)$

$\log f(A^T Y) = \text{const} - \frac{1}{2} \log |\sigma^2 V_1| - \frac{1}{2\sigma^2} Y^T A V_1^{-1} A^T Y = \text{const} - \frac{1}{2} \log |\sigma^2 V_1| - \frac{1}{2\sigma^2} Y^T A V_1^{-1} A^T Y$
 $= \text{const} - \frac{n-p}{2} \log \sigma^2 - \frac{1}{2\sigma^2} Y^T A V_1^{-1} A^T Y$
 $\frac{\partial \log f(\cdot)}{\partial \sigma^2} = -\frac{n-p}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} Y^T A V_1^{-1} A^T Y = 0$
 $-\frac{(n-p)}{2} + \frac{1}{2\sigma^2} Y^T A V_1^{-1} A^T Y = 0 \Leftrightarrow$

$\hat{\sigma}_{\text{REML}}^2 = \left(\frac{1}{n-p} \cdot Y^T A V_1^{-1} A^T Y \right)$

$c) \quad Y^T A (A^T V_0 A)^{-1} A^T Y = \text{tr} \left((A^T V_0 A)^{-1} A^T Y (A^T Y)^T \right)$
 $n \times n$ $n \times p$ $n \times n$ $n \times p$ $n \times n$
 $Y^T A (A^T V_0 A)^{-1} A^T Y = \left[\begin{matrix} 1 \times n-p \\ n \times n-p \end{matrix} \right] \times \left[\begin{matrix} n-p \times n-p \\ n \times n-p \end{matrix} \right] \times \left[\begin{matrix} n-p \times 1 \\ n \times 1 \end{matrix} \right]$
 $= \left[1 \times 1 - \text{constant} \right] = \left[\text{tr}(\text{scalar}) = \text{scalar} \right] = \text{tr} \left(\underbrace{Y^T A}_{M_1} \underbrace{(A^T V_0 A)^{-1}}_{M_2} A^T Y \right) = \left[\text{tr}(M_1 \times M_2) = \text{tr}(M_2 \times M_1) \right]$

$= \text{tr} \left((A^T V_0 A)^{-1} A^T Y \underbrace{Y^T A}_{(A^T Y)^T} \right) = \text{tr} \left((A^T V_0 A)^{-1} \right) \times \left(A^T Y \right)^T \left(A^T Y \right)$

$d) \quad E \left\{ A^T Y (A^T Y)^T \right\} = \sigma^2 A^T V_0 A$

$E \left\{ (A^T Y - 0) (A^T Y - 0)^T \right\} =$

$\text{COV}(A^T Y) = \sigma^2 A^T V_0 A$

$E \left\{ \hat{\sigma}_{\text{REML}}^2 \right\} = E \left\{ \frac{1}{n-p} \cdot Y^T A (A^T V_0 A)^{-1} A^T Y \right\} =$

$= \frac{1}{n-p} E \left\{ \text{tr} \left((A^T V_0 A)^{-1} A^T Y (A^T Y)^T \right) \right\} =$

$= \frac{1}{n-p} \text{tr} \left\{ E \left\{ (A^T V_0 A)^{-1} A^T Y (A^T Y)^T \right\} \right\} =$

$= \frac{1}{n-p} \text{tr} \left\{ (A^T V_0 A)^{-1} E \left\{ A^T Y (A^T Y)^T \right\} \right\} =$

$= \frac{1}{n-p} \text{tr} \left\{ (A^T V_0 A)^{-1} \cdot \sigma^2 A^T V_0 A \right\} =$

$= \frac{1}{n-p} \text{tr} \left(\sigma^2 \mathbf{I}_{n-p} \right) = \frac{1}{n-p} \sum_{i=1}^{n-p} \sigma^2 \cdot 1 = \frac{\sigma^2 (n-p)}{n-p} =$

$= \sigma^2$

$$\begin{aligned}
 e) \quad \tilde{A} &= A(B) \\
 \hat{\sigma}_{\text{REML}}^2 &= \frac{1}{n-p} Y^T A I (A^T V_0 A)^{-1} A^T Y = \\
 &= \frac{1}{n-p} Y^T A (B B^{-1}) (A^T V_0 A)^{-1} (B^T)^{-1} B^T A^T Y = \\
 &= \frac{1}{n-p} Y^T A B (B^T A^T V_0 A B)^{-1} B^T A^T Y = \\
 &= \frac{1}{n-p} Y^T A B (A B)^T V_0 A B (A B)^T Y = \\
 &= [A = A B] = \frac{1}{n-p} Y^T \tilde{A} (\tilde{A}^T V_0 \tilde{A})^{-1} \tilde{A}^T Y
 \end{aligned}$$

$$e) \quad X_{n \times 1} = \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix} \quad \sigma^2 V_0 = \sigma^2 I_{n \times n} = \begin{pmatrix} 1 & & & 0 \\ & \ddots & & \\ & & 1 & \\ 0 & & & \ddots & \\ & & & & 1 \end{pmatrix}$$

$$A^T = \begin{pmatrix} 1 & 0 & 0 & \dots & 0 & -1 \\ 0 & 1 & 0 & \dots & 0 & -1 \\ \dots & \dots & \dots & \dots & \dots & \dots \\ 0 & \dots & \dots & 1 & -1 \end{pmatrix}$$

$$A^T X = \begin{pmatrix} 1 & 0 & 0 & \dots & 0 & -1 \\ 0 & 1 & 0 & \dots & 0 & -1 \\ \dots & \dots & \dots & \dots & \dots & \dots \\ 0 & \dots & \dots & 1 & -1 \end{pmatrix} \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix}_{n \times 1} = \begin{pmatrix} 1-1 \\ 1-1 \\ \dots \\ 1-1 \end{pmatrix}_{n-1} = \begin{pmatrix} 0 \\ \vdots \\ 0 \end{pmatrix}_{n-1}$$

$$\begin{aligned}
 \hat{\sigma}_{\text{REML}}^2 &= \frac{1}{n-1} Y^T A (A^T V_0 A)^{-1} A^T Y = \\
 &= \frac{1}{n-1} Y^T A (A^T A)^{-1} A^T Y = [A^T A = I + \mathbb{1}\mathbb{1}^T] = \\
 &= \frac{1}{n-1} Y^T A (I + \mathbb{1}\mathbb{1}^T)^{-1} A^T Y = [\text{second hint}] = \\
 &= \frac{1}{n-1} Y^T A (I - K \mathbb{1}\mathbb{1}^T) A^T Y = [K = \frac{1}{\dim(\mathbb{1}\mathbb{1}^T)} = \frac{1}{n-1}] = \\
 &= \frac{1}{n-1} \left(\begin{matrix} Y_1 - Y_n, Y_2 - Y_n, \dots, Y_n - Y_n \end{matrix} \right)_{1 \times n} \begin{pmatrix} 1 - \frac{1}{n} & & & -\frac{1}{n} \\ & \ddots & & \\ & & 1 - \frac{1}{n} & \\ & & & \ddots & \\ & & & & 1 - \frac{1}{n} \end{pmatrix}_{n \times n}
 \end{aligned}$$

$$\begin{aligned}
 [Y^T A = (A^T Y)^T] \quad & \times A^T Y = \left[\sum_i (Y_i - Y_n) \left(1 - \frac{1}{n}\right) + \sum_{j \neq i} (Y_j - Y_n) \left(-\frac{1}{n}\right) \right]_{n \times 1} \\
 & = \left[Y_i - Y_n - \sum_j \frac{Y_j - Y_n}{n} = Y_i - \bar{Y} - \bar{Y} + Y_n = Y_i - \bar{Y} \right] =
 \end{aligned}$$

$$\frac{1}{n-1} \begin{pmatrix} Y_1 - \bar{Y} & \dots & Y_n - \bar{Y} \end{pmatrix} \begin{pmatrix} Y_1 - Y_n \\ \vdots \\ Y_n - Y_n \end{pmatrix} =$$

$$\begin{aligned}
 &= \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y}) (Y_i - \bar{Y} + \bar{Y} - Y_n) = \\
 &= \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2 + \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y}) (\bar{Y} - Y_n) = \\
 &= \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2
 \end{aligned}$$

$Y_{ij} \sim \mathcal{N}(\mu_i, \sigma^2)$, $i=1, \dots, n$, $j=1, \dots, m$
independent.

$$L(\mu_i, \sigma^2) = \prod_{i=1}^n \prod_{j=1}^m \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2} (Y_{ij} - \mu_i)^2\right)$$

$$\begin{aligned} \ell(\mu_i, \sigma^2) &= \sum_{i=1}^n \sum_{j=1}^m \left(\log \frac{1}{\sqrt{2\pi}} - \frac{1}{2} \log(\sigma^2) - \right. \\ &= \left. -\frac{1}{2\sigma^2} (Y_{ij} - \mu_i)^2 \right) \end{aligned}$$

$$\frac{\partial \ell(\cdot)}{\partial \mu_i} = \sum_{j=1}^m -\frac{1}{\sigma^2} (Y_{ij} - \mu_i) = 0 \quad (=)$$

$$(\Rightarrow) \sum_{j=1}^m Y_{ij} - m\mu_i = 0$$

$$\hat{\mu}_i = \frac{1}{m} \sum_{j=1}^m Y_{ij} = \bar{Y}_{i\cdot} \quad \forall i \in \overline{1, n}$$

$$\frac{\partial \ell(\cdot)}{\partial \sigma^2} = \sum_{i=1}^n \sum_{j=1}^m \left(-\frac{1}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} (Y_{ij} - \mu_i)^2 \right) = 0$$

$$\Rightarrow \sum_{i=1}^n \sum_{j=1}^m \left(-\frac{1}{2} + \frac{1}{2\sigma^2} (Y_{ij} - \mu_i)^2 \right) = 0$$

$$nm\sigma^2 = \sum_{i=1}^n \sum_{j=1}^m (Y_{ij} - \mu_i)^2 = \sum_{i=1}^n (Y_{i\cdot} - \hat{\mu}_i)^2$$

$$\hat{\sigma}_{ML}^2 = \frac{1}{nm} \sum_{i=1}^n \sum_{j=1}^m (Y_{ij} - \hat{\mu}_i)^2$$

$$\begin{aligned} E\{\hat{\sigma}_{ML}^2\} &= E\left\{ \frac{1}{nm} \sum_{i=1}^n \sum_{j=1}^m (Y_{ij} - \mu_i)^2 \right\} = \\ &= \frac{1}{nm} \sum_{i=1}^n \sum_{j=1}^m E\left\{ (Y_{ij} - \mu_i)^2 \right\} = \left[z_{ij} = Y_{ij} - \mu_i \right] \\ &= \left[z_{ij} \sim \mathcal{N}(0, \sigma^2) \right] = \end{aligned}$$

$$= \frac{1}{nm} \sum_{i,j} E\{z^2\} = \frac{1}{nm} \sum_{i,j} (\text{Var}(z) + E^2(z)) =$$

$$= \frac{1}{nm} \sum_{i,j} \sigma^2 = \sigma^2, \text{ providing that we } \mu_i$$

now address the case when μ_i are estimated.

$$E(\hat{\sigma}^2) = \frac{1}{nm} \sum_{i,j} (E(Y_{ij}^2) - E(\hat{\mu}_i^2)) = \frac{1}{nm} \sum_{i,j} \left(\sigma^2 + \mu_i^2 - \frac{\sigma^2}{m} - \mu_i^2 \right)$$

$$= \frac{1}{nm} \sum_{i,j} \sigma^2 \left(1 - \frac{1}{m} \right), \text{ so if } \mu_i \text{ are unknown.}$$

and are estimated we get biased estimators for $\hat{\sigma}^2$