MAT4701 - Stochastic Analysis with Applications

Exercises about Conditional Expectations, Stopping Times, Martingales

Exercise 1

Let $(X_n)_{n\in\mathbb{N}}$ be a submartingale w.r.t. the filtration $(\mathcal{F}_n)_{n\in\mathbb{N}}$ such that

$$\forall n \in \mathbb{N}, \exists k \geq n : \mathbb{E}(X_k) \leq \mathbb{E}(X_n).$$

Show that $(X_n)_{n\in\mathbb{N}}$ is a martingale!

Exercise 2

Let $X_1, X_2, ...$ be i.i.d., integrable random variables with $\mathbb{E}(X_1) = 1$. Show that $M_n := \prod_{k=1}^n X_k$ is a martingale with respect to the filtration $\mathcal{F}_n := \sigma(X_1, ..., X_n)$

Exercise 3

Let $X_1, X_2,...$ be i.i.d. with $P(X_1 = -\frac{1}{2}) = P(X_1 = \frac{1}{2}) = \frac{1}{2}$ and

$$M_n := \prod_{i=1}^n (1 + X_k).$$

- (i) Show hat $(M_n)_{n\in\mathbb{N}}$ is a martingale w.r.t. an adequate filtration.
- (ii) Show that for $a \in \mathbb{R}$

$$\tau := \min\{k \ge 1 : X_k = a\}$$

is a stopping time w.r.t. the natural filtration.

Exercise 4

Let X_1, X_2, \ldots be i.i.d. random variables which are square-integrable, but not constant. Let

$$M_n := \sum_{k=1}^n X_k X_{k+1}.$$

Under which conditions is $(M_n)_{n\in\mathbb{N}}$ w.r.t. the filtration $\mathcal{F}_n := \sigma(X_1,\ldots,X_{n+1})$ a martingale?

Exercise 5

Let $(X_n)_{n\in\mathbb{N}}$ be a sequence of $\mathcal{N}(-a,1)$ random variables and a>0.

- (i) For which value of $h \in \mathbb{R}$ is the process $Y_n = \exp(h \sum_{i=1}^n X_i)$ a martingale with respect to the natural filtration $\mathcal{F} := \sigma(X_1, \dots, X_n)$?
- (ii) For which values of h is it a sub- or supermartingale?
- (iii) Assume h = 2a, let x > 0. Define $S_n := \sum_{i=1}^n X_i$. Show that

$$P(\sup_{n\geq 1} S_n > x) \leq e^{-2ax}.$$

Exercise 6

Let $(X_n)_{n\geq 1}$ i.i.d. random variables with density

$$f(x) = \frac{1}{2}e^{-|x|}, \ x \in \mathbb{R}.$$

Find $A(\theta)$ such that $\forall |\theta| < 1$ the sequence

$$M_n = M_n(\theta) = \exp\left[\theta \sum_{i=1}^n X_k - nA(\theta)\right]$$

is a martingale w.r.t. the natural filtration $\mathcal{F}_n := \sigma(X_1, \dots, X_n)$.

Exercise 7

Let (Ω, \mathcal{F}, P) a probability space, $X \in L^2(\Omega, \mathcal{F}, P)$ a random variable and $\mathcal{A} \subset \mathcal{F}$ a sub- σ -algebra. The expected variance of X given \mathcal{A} is defined by:

$$\operatorname{Var}[X|\mathcal{A}] := \mathbb{E}[(X - \mathbb{E}[X|\mathcal{A}])^2|\mathcal{A}].$$

Show that

$$Var[X] = \mathbb{E}[Var[X|\mathcal{A}]] + Var[\mathbb{E}[X|\mathcal{A}]]$$

Exercise 8

Let $(X_k)_{k\geq 1}$ be i.i.d. random variables with $P(X_k=k)=\frac{1}{k}$ and $P(X_k=0)=\frac{k-1}{k}$. Let $M_n=\prod_{k=1}^n X_k,\,M_0=0$. Is $(M_n)_{n\in\mathbb{N}}$ a martingale?

Exercise 9

Let ξ_1, ξ_2, \ldots i.di.d with $\mathbb{E}[\xi_j^2] = \sigma^2 < \infty$ and $\mathbb{E}[\xi_j] = 0$. Then, $X_n := \sum_{i=1}^n \xi_i$ is a martingale and X_n^2 is a submartingale. Find the increasing process A_n in the Doob decomposition of X_n^2 , i.e. find A_n s.t. $X_n^2 - A_n$ is a martingale.

Solution of Exercise 1

We consider the Doob composition $X_n = M_n + A_n$ with M_n a martingale and $A_n \ge 0$ an increasing predictable process. It follows

$$\mathbb{E}[X_n] = \underbrace{\mathbb{E}[M_n]}_{=\mathbb{E}[M_{n+1}]} + \underbrace{\mathbb{E}[A_n]}_{\geq \mathbb{E}[A_{n-1}]} \geq \mathbb{E}[X_{n-1}]. \tag{1}$$

Hence, the sequence of $(\mathbb{E}[X_n])_{n\in\mathbb{N}}$ is an increasing sequence of real numbers. On the other hand, by

$$\forall n \in \mathbb{N}, \exists k \geq n : \mathbb{E}(X_k) \leq \mathbb{E}(X_n)$$

we can a extract a decreasing a subsequence $(\mathbb{E}[X_{n_k}])_{k\in\mathbb{N}}$. Hence, the whole sequence $(\mathbb{E}[X_n])_{n\in\mathbb{N}}$ must be constant. By (1), this is only possible if $\mathbb{E}[A_n] = 0$ for all n. Since $A_n \geq 0$, it follows $A_n = 0$. Hence, $X_n = M_n$ is a martingale.

Solution of Exercise 2

By definition, M_n is adapted to \mathcal{F}_n .

Furthermore by independece, we have

$$\mathbb{E}[|M_n|] = \mathbb{E}[|\prod_{i=1}^n X_i|] = \prod_{i=1}^n \mathbb{E}[|X_i|] < \infty.$$

It's left to check the Martingale property:

$$\mathbb{E}[M_n|\mathcal{F}_{n-1}] = \mathbb{E}[\underbrace{\prod_{i=1}^{n-1} X_i \cdot X_n|\mathcal{F}_{n-1}]}_{\mathcal{F}_{n-1}-\text{meas.}} = \underbrace{\prod_{i=1}^{n-1} X_i \cdot \underbrace{\mathbb{E}[X_n]}_{\text{ind. of } \mathcal{F}_{n-1}}}_{\mathcal{F}_{n-1}} \mathcal{F}_{n-1}] = \prod_{i=1}^{n-1} X_i = M_{n-1}.$$

Solution of Exercise 3(i)

 $\mathcal{F}_n := \sigma(X_1, \dots, X_n)$ is a filtration which (M_n) is adapted to. As a finite random variable M_n is also integrable. Furthermore, we have

$$\mathbb{E}[M_{n+1}|\mathcal{F}_n] = \mathbb{E}[M_n(1+X_{n+1})|\mathcal{F}_n] = M_n\mathbb{E}[1+X_{n+1}|\mathcal{F}_n]M_n\mathbb{E}[1+X_{n+1}]$$
$$= M_n(\frac{1}{2} \cdot \frac{1}{2} + \frac{3}{2} \cdot \frac{1}{2}) = M_n.$$

Solution of Exercise 3(ii)

$$\{\tau = n\} = \{X_1, \dots, X_{n-1} \neq a, X_n = a\}$$
$$= \{M_1 - 1 \neq a, \frac{M_2}{M_1} - 1 \neq a, \dots, \frac{M_n}{M_{n-1}} - 1 \neq a\} \in \sigma(M_1, \dots, M_n)$$

Solution of Exercise 4

 M_n is obviously adapted.

By Cauchy-Schwarz, each summand $X_k X_{k+1}$ is in L^1 and hence M_n is integrable. (This can also be seen by exploiting independence of X_k .) Hence, M_n is and integral if and only if

$$M_n \stackrel{!}{=} \mathbb{E}[\underbrace{M_n}_{\mathcal{F}_n - \text{meas.}} + \underbrace{X_{n+1}}_{\mathcal{F}_n - \text{meas.}} X_{n+2} | \mathcal{F}_n] = M_n + X_{n+1} \mathbb{E}[X_{n+2} | \mathcal{F}_n] = M_n + X_{n+1} \mathbb{E}[X_{n+2} | \mathcal{F}_n]$$

Since X_i are not constant, it follows that $\mathbb{E}[X_i] = 0$ must hold.

Solution of Exercise 5

i) We have $e^{hS_n} = \prod_{1 \le i \le n} e^{hX_i}$ which is adapted to the given filtration. Because of the independence of X_i , we have

$$\mathbb{E}[e^{hS_n}] = \prod_{1 \le i \le n} \mathbb{E}[e^{hX_i}]$$

and it suffices to show the integrability of e^{hX_1} :

$$\mathbb{E}[e^{hX_1}] = \int e^{hx} \frac{1}{\sqrt{2\pi}} e^{-(x+a)^2/2} dx$$

$$= e^{-ha} e^{h^2/2} \underbrace{\int \frac{1}{\sqrt{2\pi}} e^{-(x-h)^2/2} dx}_{=1, \text{ (density)}}$$

$$= e^{-ha+h^2/2} < \infty. \tag{2}$$

Hence, it is left to prove that

$$\mathbb{E}[e^{hS_n}|\mathcal{F}_{n-1}] = e^{hS_{n-1}} \cdot \underbrace{\mathbb{E}[e^{hX_n}|\mathcal{F}_{n-1}]}_{=\mathbb{E}[e^{hX_n}]} \stackrel{!}{=} e^{hS_{n-1}}$$
(3)

holds, i.e. $\mathbb{E}[e^{hX_n}] \stackrel{\text{i.i.d.}}{=} \mathbb{E}[e^{hX_1}] \stackrel{!}{=} 1$. From (2), we see that this is the case if and only if h = 2a or h = 0.

ii) For $h \in [0, 2a]$ we get a supermartingale, for $h \leq 0$ or $h \geq 2a$ we get a submartingale (which we can see by plugging in these values in the equation (3) above):

$$Y_n$$
 supermartingale $\Leftrightarrow \mathbb{E}[e^{hX_1}] \leq 1 \Leftrightarrow h \in [0,2]$

and

$$Y_n$$
 submartingale $\Leftrightarrow \mathbb{E}[e^{hX_1}] \ge 1 \Leftrightarrow h \in (-\infty, 0] \cup [2, \infty)$.

iii) We apply Doob's martingale inequality:

$$P(\sup_{n\geq 1} S_n > x) \stackrel{\text{MCT}}{=} \lim_{k\to\infty} P(\sup_{n\leq k} X_n > x) = \lim_{k\to\infty} P(\sup_{n\leq \infty} e^{hS_n} < e^{hx})$$
$$\leq \lim_{k\to\infty} \frac{\mathbb{E}[e^{hS_n}]}{e^{hx}} = \frac{1}{e^{hx}} = e^{-2ax}.$$

Solution of Exercise 6

Measurability is obvious.

From

$$M_n = \exp(\theta \sum_{k=1}^n X_k - nA(\theta)) = \exp(\sum_{k=1}^n (\theta X_k - A(\theta))) = \prod_{k=1}^n \underbrace{\exp((\theta X_k - A(\theta)))}_{=:Y_k}$$
(4)

we obtain

$$\mathbb{E}[M_n] := \prod_{k=1}^n \int_0^\infty \frac{1}{2} e^{-|x|} \exp(\theta x - A(\theta)) dx$$

$$= \prod_{k=1}^n \frac{1}{2} \left(\int_0^\infty e^{-x} \exp(\theta x - A(\theta)) dx + \int_0^\infty \exp(-(\theta + 1)x - A(\theta)) dx \right)$$

$$= \prod_{k=1}^n \frac{1}{2} \left(\int_0^\infty \exp[(\theta - 1)x - A(\theta)] ds + \int_0^\infty \exp[-(\theta + 1)x - A(\theta)] ds \right). \tag{5}$$

Since $\theta - 1$, $-\theta - 1 < 0$, this is integrable for every choice of $A(\theta)$. By independence, we get

$$\mathbb{E}[M_{n+1}|\underbrace{\sigma(M_1,\dots,M_n)}_{=\sigma(X_1,\dots,X_n)}] = M_n \cdot \mathbb{E}[\exp[\theta X_{n+1} - A(\theta))]$$
$$= M_n \cdot \mathbb{E}[\exp[(\theta X_1 - A(\theta))]].$$

So, we have to chose $A(\theta)$ such that

$$1 = \mathbb{E}[\exp(\theta X_1 - A(\theta))]$$

$$= \frac{1}{2} \left(\int_0^\infty \exp[(\theta - 1)x - A(\theta)] dx + \int_0^\infty \exp[-(\theta + 1)x - A(\theta)] dx \right)$$

$$= \frac{1}{2} \left(\frac{1}{\theta - 1} \exp((\theta - 1)x - A(\theta)) \Big|_0^\infty - \frac{1}{\theta + 1} \exp(-(\theta + 1)x - A(\theta)) \Big|_0^\infty \right)$$

$$= \frac{\exp(-A(\theta))}{(1 - \theta^2)}$$

holds. Hence, we have to chose $A(\theta) = -\ln(1 - \theta^2)$.

Solution of Exercise 7

We define $\bar{X} := \mathbb{E}[X|\mathcal{A}]$:

$$\mathbb{E}[\operatorname{Var}[X|]] + \operatorname{Var}[\bar{X}] = \mathbb{E}[\mathbb{E}[(X - \bar{X})^{2}|\mathcal{A}]] + \operatorname{Var}[\bar{X}]$$

$$= \mathbb{E}[\mathbb{E}[(X - \bar{X})^{2}|\mathcal{A}]] + \mathbb{E}[\bar{X}^{2}] + \operatorname{Var}[\bar{X}]$$

$$= \mathbb{E}[X^{2}] - 2\mathbb{E}[\mathbb{E}[X \underbrace{\bar{X}}_{A - \text{meas.}} |\mathcal{A}]] + \mathbb{E}[\bar{X}^{2}] + \operatorname{Var}[\bar{X}]$$

$$= \mathbb{E}[X^{2}] - 2\mathbb{E}[\bar{X} \cdot \bar{X}] + \mathbb{E}[\bar{X}^{2}] + \mathbb{E}[\bar{X}^{2}] - \mathbb{E}[\bar{X}]^{2}$$

$$= \mathbb{E}[X^{2}] - 2\mathbb{E}[\bar{X}^{2}] + 2\mathbb{E}[\bar{X}^{2}] - \mathbb{E}[X^{2}]^{2}$$

$$= \mathbb{E}[X^{2}] - \mathbb{E}[X]^{2} = \operatorname{Var}[X]$$
(6)

Solution of Exercise 8

Yes, M_n is a martingale with respect to $\mathcal{F}_n := \sigma(M_0, \dots, M_n)$. It is adapted by definition and integrable because of

$$\mathbb{E}[|M_n|] = \mathbb{E}[M_n] = \prod_{i=1}^n \mathbb{E}[X_i] = 1.$$

Finally, we have:

$$\mathbb{E}[M_{n+1}|\mathcal{F}_n] = \prod_{i=1}^n \mathbb{E}[X_i|\mathcal{F}_n] \cdot \mathbb{E}[X_{n+1}|\mathcal{F}_n] = \prod_{i=1}^n X_i \mathbb{E}[X_{n+1}] = M_n \cdot 1 = M_n.$$

Solution of Exercise 9

Claim: $A_n = n\sigma^2$

Proof. A_n is positive, increasing and as a constant random variable \mathcal{F}_{n-1} measurable. Furthermore,

$$\mathbb{E}[X_n^2 - n\sigma^2 | \mathcal{F}_{n-1}] = \sum_{1 \le i, j \le n-1} \underbrace{\mathbb{E}[\xi_i \xi_j | \mathcal{F}_{n-1}]}_{=\xi_i \xi_j} + \underbrace{\mathbb{E}[\xi_n^2]}_{=\sigma^2} + 2 \sum_{j=1}^{n-1} \underbrace{\mathbb{E}[\xi_j \xi_n | \mathcal{F}_{n-1}]}_{=\mathbb{E}[\xi_j] \mathbb{E}[\xi_n] = 0} - n\sigma^2$$
$$= X_{n-1}^2 - (n-1) \cdot \sigma^2$$