8. The Cox-Ross-Rubinstein Model

S. Ortiz-Latorre STK-MAT 3700/4700 An Introduction to Mathematical Finance November 15, 2021

Department of Mathematics University of Oslo Bernoulli process and related processes

The Cox-Ross-Rubinstein model

Pricing European options in the CRR model

Hedging European options in the CRR model

Introduction

- The Cox-Ross-Rubinstein (**CRR**) market model, also known as the binomial model, is an example of a multi-period market model.
- At each point in time, the stock price is assumed to either go 'up' by a fixed factor *u* or go 'down' by a fixed factor *d*.

$$S(t) \xrightarrow{p} S(t+1) = S(t)u$$

$$1 - p \xrightarrow{S(t+1)} S(t+1) = S(t)d$$

- Only four parameters are needed to specify the binomial asset pricing model: u > 1 > d > 0, r > -1 and S(0) > 0.
- The real-world probability of an 'up' movement is assumed to be 0 for each period and is assumed to be independent of all previous stock price movements.

Bernoulli process and related processes

The Bernoulli process

Definition 1

A stochastic process $X = \{X(t)\}_{t \in \{1,...,T\}}$ defined on some probability space (Ω, \mathcal{F}, P) is said to be a (truncated) **Bernoulli process** with parameter 0 (and timehorizon*T*) if the random variables <math>X(1), X(2), ..., X(T) are independent and have the following common probability distribution

$$P(X(t) = 1) = 1 - P(X(t) = 0) = p, \quad t \in \mathbb{N}.$$

- We can think of a Bernoulli process as the random experiment of flipping sequentially *T* coins.
- The sample space Ω is the set of vectors of zero's and one's of length *T*. Obviously, $\#\Omega = 2^T$.

The Bernoulli process

- $X(t, \omega)$ takes the value 1 or 0 as ω_t , the *t*-th component of $\omega \in \Omega$, is 1 or 0, that is, $X(t, \omega) = \omega_t$.
- \mathcal{F}_t^X is the algebra corresponding to the observation of the first *t* coin flips.
- $\mathcal{F}_t^X = \mathfrak{a}(\pi_t)$ where π_t is a partition with 2^t elements, one for each possible sequence of t coin flips.
- The probability measure P is given by

$$P(\omega) = p^n \left(1 - p\right)^{T - n},$$

where ω is any elementary outcome corresponding to n "heads" and T - n "tails".

• Setting this probability measure on Ω is equivalent to say that the random variables X(1), ..., X(T) are independent and identically distributed.

The Bernoulli process

Example

• Consider T = 3. Let

$$\begin{split} A_0 &= \{ (0,0,0), (0,0,1), (0,1,0), (0,1,1) \}, \\ A_1 &= \{ (1,0,0), (1,0,1), (1,1,0), (1,1,1) \}, \\ A_{0,0} &= \{ (0,0,0), (0,0,1) \}, \quad A_{0,1} &= \{ (0,1,0), (0,1,1) \}, \\ A_{1,0} &= \{ (1,0,0), (1,0,1) \}, \quad A_{1,1} &= \{ (1,1,0), (1,1,1) \}. \end{split}$$

• We have that $\pi_0 = \{\Omega\}$, $\pi_1 = \{A_0, A_1\}$,

$$\pi_2 = \{A_{0,0}, A_{0,1}, A_{1,0}, A_{1,1}\},\$$

$$\pi_{3} = \{\{\omega\}\}_{\omega \in \Omega}.$$

• $\mathcal{F}_{t} = \mathfrak{a}(\pi_{t}), t = 0, ..., 3.$ In particular, $\mathcal{F}_{3} = \mathcal{P}(\Omega).$

Definition 2

The **Bernoulli counting process** $N = \{N(t)\}_{t \in \{0,...,T\}}$ is defined in terms of the Bernoulli process X by setting N(0) = 0 and

$$N(t,\omega) = X(1,\omega) + \dots + X(t,\omega), \qquad t \in \{1,\dots,T\}, \quad \omega \in \Omega.$$

- The Bernoulli counting process is an example of *additive random walk*.
- The random variable N(t) should be thought as the number of heads in the first t coin flips.

The Bernoulli counting process

• Since $\mathbb{E} [X(t)] = p$, Var [X(t)] = p (1 - p) and the random variables X(t) are independent, we have

 $\mathbb{E}\left[N\left(t\right)\right] = tp, \qquad \operatorname{Var}\left[N\left(t\right)\right] = tp\left(1-p\right).$

• Moreover, for all $t \in \{1, ..., T\}$ one has

$$P(N(t) = n) = {t \choose n} p^n (1-p)^{t-n}, \quad n = 0, ..., t,$$

that is, $N(t) \sim Binomial(t, p)$.

The Cox-Ross-Rubinstein model

- The bank account process is given by $B = \left\{ B(t) = (1+r)^t \right\}_{t=0,\dots,T}.$
- The binomial security price model features 4 parameters: p, d, u and S(0), where 0 and <math>S(0) > 0.
- The time t price of the security is given by

$$S(t) = S(0) u^{N(t)} d^{t-N(t)}, \quad t = 1, ..., T.$$

• The underlying Bernoulli process X governs the up and down movements of the stock. The stock price moves up at time t if $X(t, \omega) = 1$ and moves down if $X(t, \omega) = 0$.

- The Bernoulli counting process N counts the up movements. Before and including time t, the stock price moves up N(t) times and down t N(t) times.
- The dynamics of the stock price can be seen as an example of a *multiplicative or geometric random walk*.
- The price process has the following probability distribution

$$P(S(t) = S(0) u^{n} d^{t-n}) = {\binom{t}{n}} p^{n} (1-p)^{t-n}, \quad n = 0, ..., t.$$

Lattice representation



- The event $\{S(t) = S(0) u^n d^{t-n}\}$ occurs if and only if exactly n out of the first t moves are up. The order of these t moves does not matter.
- At time *t*, there are 2^{*t*} possible sample paths of length *t*.
- At time t, the price process S(t) can only take one of t + 1 possible values.
- This reduction, from exponential to linear in time, in the number of relevant nodes in the lattice is crucial in numerical implementations.

Example

Consider T = 2. Let

$$\Omega = \{ (d,d), (d,u), (u,d), (u,u) \}$$

$$A_d = \{ (d,d), (d,u) \}, \quad A_u = \{ (u,d), (u,u) \}.$$

We have that $\pi_0 = \{\Omega\}, \pi_1 = \{A_d, A_u\}, \pi_2 = \{\{(d, d)\}, \{(d, u)\}, \{(u, d)\}, \{(u, u)\}\}, \text{ and } \mathcal{F}_t = \mathfrak{a}(\pi_t), t = 0, ..., 3.$ Note that

$$\{S(2) = S(0) ud\} = \{(d, u), (u, d)\} \notin \pi_2.$$

Hence, the lattice representation is NOT the information tree of the model.

Theorem 3

There exists a unique martingale measure in the CRR market model if and only if d < 1 + r < u, and is given by

$$Q\left(\omega\right)=q^{n}\left(1-q\right)^{T-n},$$

where ω is any elementary outcome corresponding to n up movements and T - n down movement of the stock and

$$q = \frac{1+r-d}{u-d}$$

Corollary 4

If d < 1 + r < u, then the CRR model is arbitrage free and complete.

Lemma 5

Let Z be a r.v. defined on some prob. space (Ω, \mathcal{F}, P) , with P(Z = a) + P(Z = b) = 1 for $a, b \in \mathbb{R}$. Let $\mathcal{G} \subset \mathcal{F}$ be an algebra on Ω . If $\mathbb{E}[Z|\mathcal{G}]$ is constant then Z is independent of \mathcal{G} . (Note that the constant must be equal to $\mathbb{E}[Z]$).

Proof of Lemma 5.

Let
$$A = \{Z = a\}$$
 and $A^c = \{Z = b\}$. Then for any $B \in \mathcal{G}$

$$\mathbb{E}\left[Z\mathbf{1}_{B}\right] = \mathbb{E}\left[\left(a\mathbf{1}_{A} + b\mathbf{1}_{A^{c}}\right)\mathbf{1}_{B}\right] = aP\left(A \cap B\right) + bP\left(A^{c} \cap B\right),$$

and

$$\mathbb{E}\left[\mathbb{E}\left[Z\right]\mathbf{1}_{B}\right] = \mathbb{E}\left[\left(aP\left(A\right) + bP\left(B\right)\right)\mathbf{1}_{B}\right] = aP\left(A\right)P\left(B\right) + bP\left(A^{c}\right)P\left(B\right).$$

By the definition of cond. expect. we have that $\mathbb{E}[Z\mathbf{1}_B] = \mathbb{E}[\mathbb{E}[Z]\mathbf{1}_B]$. Using that $P(A^c) = 1 - P(A)$ and $P(A^c \cap B) = P(B) - P(A \cap B)$, we get that $P(A \cap B) = P(A) P(B)$ and $P(A^c \cap B) = P(A^c) P(B)$, which yields that $\mathfrak{a}(Z)$ is independent of \mathcal{G} .

Proof of Theorem 3.

Note that
$$S^{*}(t) = S(t)(1+r)^{-t}$$
, $t = 0, ...T$. Moreover

$$\frac{S(t+1)}{S(t)} = \frac{S(0) u^{N(t+1)} d^{t+1-N(t+1)}}{S(0) u^{N(t)} d^{t-N(t)}}$$
$$= u^{N(t+1)-N(t)} d^{1-(N(t+1)-N(t))}$$
$$= u^{X(t+1)} d^{1-X(t+1)}, \quad t = 0, ..., T-1$$

Let Q be another probability measure on Ω . We impose the martingale condition under Q $T = \begin{bmatrix} C^*(t+1) & T \end{bmatrix} = C^*(t) \leftrightarrow T = \begin{bmatrix} X(t+1) & 1 - X(t+1) \\ T \end{bmatrix} = 1 + 1$

$$\mathbb{E}_{Q}\left[S^{*}\left(t+1\right)|\mathcal{F}_{t}\right] = S^{*}\left(t\right) \Leftrightarrow \mathbb{E}_{Q}\left[u^{X\left(t+1\right)}d^{1-X\left(t+1\right)}\middle|\mathcal{F}_{t}\right] = 1+r.$$

Proof of Theorem 3.

This gives

$$(1+r) = \mathbb{E}_{Q} \left[u^{X(t+1)} d^{1-X(t+1)} \middle| \mathcal{F}_{t} \right]$$

= $uQ \left(X (t+1) = 1 \middle| \mathcal{F}_{t} \right) + dQ \left(X (t+1) = 0 \middle| \mathcal{F}_{t} \right).$

In addition,

$$1 = Q(X(t+1) = 1 | \mathcal{F}_t) + Q(X(t+1) = 0 | \mathcal{F}_t).$$

Solving the previous equations we get the unique solution

$$Q(X(t+1) = 1 | \mathcal{F}_t) = \frac{1+r-d}{u-d} = q,$$

$$Q(X(t+1) = 0 | \mathcal{F}_t) = \frac{u-(1+r)}{u-d} = 1-q.$$
17/33

Proof of Theorem 3.

Note that the r.v. $u^{X(t+1)}d^{1-X(t+1)}$ satisfies the hypothesis of Lemma 5 and, therefore, $u^{X(t+1)}d^{1-X(t+1)}$ is independent (under Q) of \mathcal{F}_t .

This means that

$$\begin{split} (1+r) &= \mathbb{E}_{Q} \left[\left. u^{X(t+1)} d^{1-X(t+1)} \right| \, \mathcal{F}_{t} \right] \\ &= \mathbb{E}_{Q} \left[u^{X(t+1)} d^{1-X(t+1)} \right] \\ &= uQ \left(X \left(t+1 \right) = 1 \right) + dQ \left(X \left(t+1 \right) = 0 \right), \end{split}$$

and we get that

$$Q (X (t+1) = 1) = Q (X (t+1) = 1 | \mathcal{F}_t),$$

$$Q (X (t+1) = 0) = Q (X (t+1) = 0 | \mathcal{F}_t).$$
18/33

Proof of Theorem 3.

As the previous unconditional probabilities does not depend on t we obtain that the random variables X(1), ...X(T) are identically distributed under Q, i.e. X(i) = Bernoulli(q). Moreover, for $a \in \{0, 1\}^T$ we have that

$$Q\left(\bigcap_{t=1}^{T} \{X(t) = a_t\}\right) = \mathbb{E}_Q\left[\prod_{t=1}^{T} \mathbf{1}_{\{X(t)=a_t\}}\right]$$
$$= \mathbb{E}_Q\left[\prod_{t=1}^{T-1} \mathbf{1}_{\{X(t)=a_t\}} \mathbb{E}_Q\left[\mathbf{1}_{\{X(T)=a_T\}} \middle| \mathcal{F}_{T-1}\right]\right]$$
$$= \mathbb{E}_Q\left[\prod_{t=1}^{T-1} \mathbf{1}_{\{X(t)=a_t\}} Q\left(X(T) = a_T \middle| \mathcal{F}_{T-1}\right)\right]$$
$$= \mathbb{E}_Q\left[\prod_{t=1}^{T-1} \mathbf{1}_{\{X(t)=a_t\}}\right] Q\left(X(T) = a_T\right)$$
$$= Q\left(\bigcap_{t=1}^{T-1} \{X(t) = a_t\}\right) Q\left(X(T) = a_T\right).$$

19/33

Proof of Theorem 3.

Iterating this procedure we get that

$$Q\left(\bigcap_{t=1}^{T} \left\{ X\left(t\right) = a_{t} \right\} \right) = \prod_{t=1}^{T} Q\left(X\left(t\right) = a_{t} \right),$$

and we can conclude that X(1), ...X(T) are also independent under Q.

Therefore, under Q, we obtain the same probabilistic model as under P but with p = q, that is,

$$Q(\omega) = q^n (1-q)^{T-n}, \qquad n = \sum_{t=1}^T \omega_t.$$

The conditions for q are equivalent to $Q(\omega) > 0$, which yields that Q is the unique martingale measure.

• By the general theory developed for multiperiod markets we have the following result.

Proposition 6 (Risk Neutral Pricing Principle)

The arbitrage free price process of a European contingent claim *X* in the CRR model is given by

$$P_{X}(t) = B(t) \mathbb{E}_{Q} \left[\frac{X}{B(T)} \middle| \mathcal{F}_{t} \right]$$
$$= (1+r)^{-(T-t)} \mathbb{E}_{Q} \left[X \middle| \mathcal{F}_{t} \right], \quad t = 0, ..., T,$$

where Q is the unique martingale measure characterized by $q = \frac{1+r-d}{u-d}$.

• Given g, a non-negative function, define

$$F_{p,g}(t,x) := \sum_{n=0}^{t} \begin{pmatrix} t \\ n \end{pmatrix} p^n \left(1-p\right)^{t-n} g\left(x u^n d^{t-n}\right).$$

Proposition 7

Consider a European contingent claim of the form X = g(S(T)). Then, the arbitrage free price process $P_X(t)$ is given by

$$P_X(t) = (1+r)^{-(T-t)} F_{q,g}(T-t, S(t)), \qquad t = 0, ..., T,$$

where $q = \frac{1+r-d}{u-d}$.

Proof of Proposition 7.

Recall that

$$S(t) = S(0) u^{N(t)} d^{t-N(t)} = S(0) \prod_{j=1}^{t} u^{X_j} d^{1-X_j}, \quad t = 1, ..., T.$$

By Proposition 6 we have that

$$(1+r)^{(T-t)} P_X(t) = \mathbb{E}_Q \left[g\left(S\left(T\right) \right) | \mathcal{F}_t \right] = \mathbb{E}_Q \left[g\left(S\left(t\right) \prod_{j=t+1}^T u^{X_j} d^{1-X_j} \right) \middle| \mathcal{F}_t \right]$$
$$= \mathbb{E}_Q \left[g\left(s \prod_{j=t+1}^T u^{X_j} d^{1-X_j} \right) \right] \bigg|_{s=S(t)} = F_{q,g}\left(T - t, S\left(t\right) \right),$$

where in the last equality we have used that S(t) is \mathcal{F}_t -measurable and $X_{t+1}, ..., X_T$ are independent of \mathcal{F}_t .

Note that if X is \mathcal{G} -measurable and Y is independent of \mathcal{G} then

$$\mathbb{E}\left[f\left(X,Y\right)|\mathcal{G}\right] = \mathbb{E}\left[f\left(x,Y\right)\right]|_{x=X}.$$

23/33

Corollary 8

Consider a European call option with expiry time T and strike price K writen on the stock S. The arbitrage free price $P_C(t)$ of the call option is given by

$$P_{C}(t) = S(t) \sum_{n=\hat{n}}^{T-t} {\binom{T-t}{n}} \hat{q}^{n} (1-\hat{q})^{T-t-n} - \frac{K}{(1+r)^{T-t}} \sum_{n=\hat{n}}^{T-t} {\binom{T-t}{n}} q^{n} (1-q)^{T-t-n},$$

where

$$\hat{n} = \inf \left\{ n \in \mathbb{N} : n > \log \left(K / (S(t) d^{T-t}) \right) / \log (u/d) \right\},$$

and $\hat{q} = \frac{qu}{1+r} \in (0,1).$

24/33

Proof of Corollary 8.

First note that

$$S(t) u^n d^{T-t-n} - K > 0 \iff n > \log\left(K/(S(t) d^{T-t})\right) / \log(u/d).$$

Let $g(x) = (x - K)^+$. If $\hat{n} > T - t$ then $F_{q,g}(T - t, S(t)) = 0$. If $\hat{n} \le T - t$, then the formula in Proposition 7 yields

$$\begin{aligned} &(1+r)^{T-t} P_C(t) \\ &= F_{q,g}(T-t,S(t)) \\ &= \sum_{n=0}^{T-t} {T-t \choose n} q^n (1-q)^{T-t-n} \left(S(t) u^n d^{T-t-n} - K\right)^+ \\ &= \sum_{n=0}^{\hat{n}} {T-t \choose n} q^n (1-q)^{T-t-n} 0 \\ &+ \sum_{n=\hat{n}}^{T-t} {T-t \choose n} q^n (1-q)^{T-t-n} \left(S(t) u^n d^{T-t-n} - K\right) \end{aligned}$$

Proof of Corollary 8.

$$=\sum_{n=\hat{n}}^{T-t} {\binom{T-t}{n}} q^n (1-q)^{T-t-n} S(t) u^n d^{T-t-n}$$
$$-\sum_{n=\hat{n}}^{T-t} {\binom{T-t}{n}} q^n (1-q)^{T-t-n} K$$
$$= S(t) \sum_{n=\hat{n}}^{T-t} {\binom{T-t}{n}} (qu)^n ((1-q) d)^{T-t-n}$$
$$-K \sum_{n=\hat{n}}^{T-t} {\binom{T-t}{n}} q^n (1-q)^{T-t-n} .$$

The result follows by defining $\hat{q} = \frac{qu}{1+r}$ and noting that

$$1 - \hat{q} = \frac{1 + r - qu}{1 + r} = \frac{qu + (1 - q)d - qu}{1 + r} = \frac{(1 - q)d}{1 + r},$$

where we have used $qu + (1-q)d = \mathbb{E}_Q\left[u^{X(t+1)}d^{1-X(t+1)}\right] = 1 + r.$ \Box 26/33

- Let X be a contingent claim and $P_X = \{P_X(t)\}_{t=0,\dots,T}$ be its price process (assumed to be computed/known).
- As the CRR model is complete we can find a self-financing trading strategy

$$H = \{H(t)\}_{t=1,...,T} = \{(H_0(t), H_1(t))^T\}_{t=1,...,T} \text{ such that}$$

$$P_X(t) = V(t) = H_0(t)(1+r)^t + H_1(t)S(t), \quad t = 1,...,T,$$
(1)

$$P_{X}(0) = V(0) = H_{0}(1) + H_{1}(1) S(0).$$

- Given t = 1, ..., T we can use the information up to (and including) t 1 to ensure that H is predictable.
- Hence, at time t, we know S(t-1) but we only know that

$$S(t) = S(t-1) u^{X(t)} d^{1-X(t)}.$$

- Using that $u^{X(t)}d^{1-X(t)} \in \{u, d\}$ we can solve equation (1) uniquely for $H_0(t)$ and $H_1(t)$.
- Making the dependence of P_X explicit on S we have the equations

$$P_X(t, S(t-1)u) = H_0(t)(1+r)^t + H_1(t)S(t-1)u,$$

$$P_X(t, S(t-1)d) = H_0(t)(1+r)^t + H_1(t)S(t-1)d.$$

The solution for these equations is

$$H_{0}(t) = \frac{uP_{X}(t, S(t-1)d) - dP_{X}(t, S(t-1)u)}{(1+r)^{t}(u-d)},$$

$$H_{1}(t) = \frac{P_{X}(t, S(t-1)u) - P_{X}(t, S(t-1)d)}{S(t-1)(u-d)}.$$

• The previous formulas only make use of the lattice representation of the model and not the information tree.

Proposition 9

Consider a European contingent claim X = g(S(T)). Then, the replicating trading strategy $H = \{H(t)\}_{t=1,...,T} = \{(H_0(t), H_1(t))^T\}_{t=1,...,T}$ is given by $H_0(t) = \frac{uF_{q,g}(T - t, S(t - 1)d) - dF_{q,g}(T - t, S(t - 1)u)}{(1 + r)^T(u - d)},$ $H_1(t) = \frac{(1 + r)^{T-t} \{F_{q,g}(T - t, S(t - 1)u) - F_{q,g}(T - t, S(t - 1)d)\}}{S(t - 1)(u - d)}.$

• Let

$$C(\tau, x) = \sum_{n=0}^{\tau} \begin{pmatrix} \tau \\ n \end{pmatrix} q^n (1-q)^{\tau-n} \left(x u^n d^{\tau-n} - K \right)^+.$$

• Then,

$$P_{C}(t) = (1+r)^{-(T-t)} C(T-t, S(t)).$$

• In the following theorem we combine the previous formula and Proposition 9 to find the hedging strategy for a European call option.

Proposition 10

The replicating trading strategy $H = \{H(t)\}_{t=1,\dots,T} = \{(H_0(t), H_1(t))^T\}_{t=1,\dots,T} \text{ for a}$ European call option with strike K and expiry time T is given by

$$H_{0}(t) = \frac{uC(T-t,S(t-1)d) - dC(T-t,S(t-1)u)}{(1+r)^{T}(u-d)},$$
$$H_{1}(t) = \frac{(1+r)^{T-t} \{C(T-t,S(t-1)u) - C(T-t,S(t-1)d)\}}{S(t-1)(u-d)}$$

- As $C(\tau, x)$ is increasing in x we have that $H_1(t) \ge 0$, that is, the replicating strategy does not involve short-selling.
- This property extends to any European contingent claim $_{31/33}$ with increasing payoff g.

- We can also use the value of the contingent claim X and backward induction to find its price process P_X and its replicating strategy H simultaneously.
- We have to choose a replicating strategy H(T) based on the information available at time T 1.
- This gives raise to two equations

$$P_X(T, S(T-1)u) = H_0(T)(1+r)^T + H_1(T)S(T-1)u,$$
(2)

$$P_X(T, S(T-1)d) = H_0(T)(1+r)^T + H_1(T)S(T-1)d.$$
(3)

• The solution is

$$H_{0}(T) = \frac{uP_{X}(T, S(T-1)d) - dP_{X}(T, S(T-1)u)}{(1+r)^{T}(u-d)},$$

$$H_{1}(T) = \frac{P_{X}(T, S(T-1)u) - P_{X}(T, S(T-1)d)}{S(T-1)(u-d)}.$$

• Next, using that H is self-financing, we can compute

$$P_X(T-1, S(T-1)) = H_0(T)(1+r)^{T-1} + H_1(T)S(T-1),$$

and repeat the procedure (changing T to T-1 in equations (2) and (3)) to compute H(T-1).