

### The Central Limit Theorem

$\{X_n\}$  i.i.d with mean  $\mu$  and variance  $\sigma^2$

Laws of large numbers:  $\frac{X_1 + X_2 + \dots + X_n}{n} \rightarrow \mu$

Central Limit Theorem:  $\frac{X_1 + X_2 + \dots + X_n - n\mu}{\sqrt{n}} \rightarrow N(0, \sigma^2)$

Idea:  $\varphi_{\frac{X_1 + X_2 + \dots + X_n - n\mu}{\sqrt{n}}} \rightarrow$  charact. fund.  $N(0, \sigma^2)$

The Lin. Cent. Th. will give us convergence in distribution

Recall:  $\varphi_X(t) = 1 + tE[X]it - \frac{1}{2}E[X^2]t^2 + o(t^3)$   
 $= 1 + \mu it - \frac{1}{2}\sigma^2 t^2 + o(t^3)$

Central Limit Theorem (i.i.d case) Assume that  $\{X_n\}$

is an i.i.d. sequence of random variables with mean  $\mu$  and variance  $\sigma^2$ . Let  $S_n = X_1 + X_2 + \dots + X_n$ . Then

$$\frac{S_n - n\mu}{\sqrt{n}\sigma} \Rightarrow N(0, 1)$$

Proof: We may assume that  $\mu = 0$  (otherwise just replace  $X_n$  by  $X'_n = X_n - \mu$ ). Then

$$\varphi_{X_i}(t) = 1 - \frac{\sigma^2 t^2}{2} + o(t^3)$$

Then  $\varphi_{\frac{S_n}{\sqrt{n}\sigma}}(t) = E\left[e^{it \frac{S_n}{\sqrt{n}\sigma}}\right] = E\left[e^{i \frac{t}{\sqrt{n}\sigma} (X_1 + X_2 + \dots + X_n)}\right]$

$$= E\left[e^{i \frac{t}{\sqrt{n}\sigma} X_1} e^{i \frac{t}{\sqrt{n}\sigma} X_2} \dots e^{i \frac{t}{\sqrt{n}\sigma} X_n}\right]$$

$$\stackrel{\text{independence}}{=} E\left[e^{i \frac{t}{\sqrt{n}\sigma} X_1}\right] E\left[e^{i \frac{t}{\sqrt{n}\sigma} X_2}\right] \dots E\left[e^{i \frac{t}{\sqrt{n}\sigma} X_n}\right]$$

$$= \varphi_{X_1}\left(\frac{t}{\sqrt{n}\sigma}\right)^n = \left(1 - \frac{\sigma^2 \left(\frac{t}{\sqrt{n}\sigma}\right)^2}{2} + o\left(\left(\frac{t}{\sqrt{n}\sigma}\right)^3\right)\right)^n$$

$$= \left(1 - \frac{t^2}{2n} + o\left(\frac{t^3}{n^{3/2}\sigma^3}\right)\right)^n$$

$$\boxed{\text{Recall } (1 + \frac{z_n}{n})^n \rightarrow e^z \text{ if } z_n \rightarrow z}$$

$\rightarrow e^{-\frac{t^2}{2}}$  which is the ch.f. of  $N(0, 1)$

By Lévy's Continuity Theorem  $\frac{S_n}{\sqrt{n}\sigma} \Rightarrow N(0, 1)$ .

Central Limit Theorem (Lyapunov's version) Assume that  $\{X_n\}$

is a sequence of independent random variables with mean 0 and variance  $\sigma_n^2$ . Assume that we have third moments and put  $\gamma_j = E[|X_j|^j]$ . Let  $S_n = X_1 + X_2 + \dots + X_n$  and

$\Delta_n^2 = \sigma_1^2 + \sigma_2^2 + \dots + \sigma_n^2$ . Assume  $\frac{\sum_{j=1}^n \gamma_j}{\Delta_n^3} \rightarrow 0$  as  $n \rightarrow \infty$ .

Then

$\frac{S_n}{\Delta_n} \Rightarrow N(0, 1)$  as  $n \rightarrow \infty$ .

Remarks: (i) Taylor expansion:  $f(x) = f(0) + f'(0)x + \frac{1}{2}f''(0)x^2$

For random variables with mean 0 and variance  $\sigma^2$   $+ \frac{1}{6}f'''(\xi)x^3$

$$\begin{aligned} \varphi(t) = E[e^{itX}] &= 1 + i E[X]t - \frac{1}{2} E[|X|^2]t^2 \\ &\quad - \frac{i}{6} E[|X|^3 e^{i\xi X}]t^3 \\ &= 1 - \frac{1}{2} \sigma^2 t^2 + \frac{1}{6} t^3 \gamma \frac{\Delta}{\Delta^3} E[|X|^3] \\ &\quad \leq 1 \text{ in absolute value} \end{aligned}$$

(ii) How reasonable is the condition  $\frac{\sum_{j=1}^n \gamma_j}{\Delta_n^3} \rightarrow 0$ ?

Check the i.i.d case:  $\frac{n\gamma}{(\Delta_n^2)^{3/2}} = \frac{n\gamma}{(n\sigma^2)^{3/2}} = \frac{n\gamma}{n^{3/2}\sigma^3} = \frac{\gamma}{n^{1/2}\sigma^3} \rightarrow 0$

(iii)  $\left(\frac{\sigma_j^2}{\Delta_n^2}\right)^{3/2} = \frac{E[|X_j|^3]^{3/2}}{\Delta_n^3} \leq \frac{E[|X_j|^3]}{\Delta_n^3}$

Lyapunov's inequality  
 $|E[X]^p| \leq E[|X|^p]$

$= \frac{\gamma_j}{\Delta_n^3} \leq \frac{\sum_{j=1}^n \gamma_j}{\Delta_n^3} \xrightarrow{\text{L.E.}} 0$

Hence  $\frac{\sigma_j^2}{\Delta_n^2} \rightarrow 0$  as  $n \rightarrow \infty$ .

Proof: We have

$$\varphi_j(t) = E[e^{itJ_j}] = 1 - \frac{t^2 \sigma_j^2}{2} + \frac{1}{6} \Delta_j \gamma_j t^3$$

when  $|\Delta_j| \leq 1$ .

$$\begin{aligned} \varphi_{\frac{S_n}{\Delta_n}}(t) &= E[e^{it \frac{S_n}{\Delta_n}}] = E[e^{i \frac{t}{\Delta_n} (\sum_{j=1}^n J_j)}] \\ &= \varphi_1\left(\frac{t}{\Delta_n}\right) \varphi_2\left(\frac{t}{\Delta_n}\right) \dots \varphi_n\left(\frac{t}{\Delta_n}\right) \\ &= \prod_{j=1}^n \left( 1 - \frac{\left(\frac{t}{\Delta_n}\right)^2 \sigma_j^2}{2} + \frac{1}{6} \Delta_j \gamma_j \left(\frac{t}{\Delta_n}\right)^3 \right) \\ &= \prod_{j=1}^n (1 + \mathcal{D}_{j,n}) \end{aligned}$$

Want to prove  $\varphi_{\frac{S_n}{\Delta_n}}(t) \rightarrow e^{-\frac{t^2}{2}}$  or equivalently

$$\text{Log } \varphi_{\frac{S_n}{\Delta_n}}(t) \rightarrow -\frac{t^2}{2} \quad \text{Log}(1+x) = x - \frac{x^2}{2} + \frac{x^3}{3} - \dots$$

$$\begin{aligned} \text{Log}(\varphi_{\frac{S_n}{\Delta_n}}(t)) &= \text{Log} \prod_{j=1}^n (1 + \mathcal{D}_{j,n}) = \sum_{j=1}^n \text{Log}(1 + \mathcal{D}_{j,n}) \\ &= \sum_{j=1}^n [\mathcal{D}_{j,n} + \varepsilon(\mathcal{D}_{j,n})] \quad \text{where } |\varepsilon(\mathcal{D}_{j,n})| \leq |\mathcal{D}_{j,n}|^2 \leftarrow \text{for big } n \end{aligned}$$

Then we have things to prove:

(i)  $\sum_{j=1}^n \mathcal{D}_{j,n} \rightarrow -\frac{t^2}{2}$

(ii)  $\sum_{j=1}^n \varepsilon(\mathcal{D}_{j,n}) \rightarrow 0$  (suffices to prove that  $\sum_{j=1}^n |\mathcal{D}_{j,n}|^2 \rightarrow 0$ )

We get:

$$(i) \sum_{j=1}^n \mathcal{D}_{j,n} = \sum_{j=1}^n \left( -\frac{\sigma_j^2 t^2}{2 \Delta_n^2} + \frac{1}{6} \Delta_j \gamma_j \left(\frac{t}{\Delta_n}\right)^3 \right)$$

$$\rightarrow -\frac{t^2}{2} + 0 = -\frac{t^2}{2}$$

(ii)  $\sum_{j=1}^n |\mathcal{D}_{j,n}|^2 \leq \sup_{j \leq n} |\mathcal{D}_{j,n}| \sum_{j=1}^n |\mathcal{D}_{j,n}|$

$$\sup_{1 \leq j \leq n} |\mathcal{D}_{j,n}| = \sup_{j \leq n} \left( \frac{\sigma_j^2 t^2}{2 \Delta_n^2} + \frac{1}{6} \Delta_j \gamma_j \left(\frac{t}{\Delta_n}\right)^3 \right) \rightarrow 0$$

need  $\frac{\sigma_j^2 t^2}{2 \Delta_n^2} \rightarrow 0$  as  $n \rightarrow \infty$

hence  $\sum_{j=1}^n |\mathcal{D}_{j,n}|^2 \rightarrow 0$  and we are done.

$$\frac{\sigma_j^2}{\Delta_n^2} \rightarrow 0$$

## Stochastic Processes (7.1) + (7.4)

Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a prob. space. and let  $\mathbb{T} \subseteq \mathbb{R}$ .

A stochastic process on  $(\Omega, \mathcal{F}, \mathbb{P})$  indexed by  $\mathbb{T}$  is just a collection  $\{X_t\}_{t \in \mathbb{T}}$  of random variables on  $(\Omega, \mathcal{F}, \mathbb{P})$ .

Typical cases:

(i)  $\mathbb{T} = \mathbb{N}$

(ii)  $\mathbb{T} = \mathbb{Z}$

(iii)  $\mathbb{T} = \{1, 2, 3, \dots, N\}$

discrete time processes

(iv)  $\mathbb{T} = [0, \infty)$

(v)  $\mathbb{T} = \mathbb{R}$

(vi)  $\mathbb{T} = [a, b]$

continuous time processes.

Fix an  $\omega \in \Omega$  (fix one scenario). Then

$t \mapsto X_t(\omega)$  is called a path of the process.

Example: Toss a coin  $N$  times and call the results

1 for heads and -1 for tails.

Basic event:  $\omega = (\omega_1, \omega_2, \dots, \omega_N)$  where each  $\omega_i = \pm 1$ .

$$\Omega = \{ \omega : \omega = (\omega_1, \omega_2, \dots, \omega_N) \}$$

$$S_n(\omega) = \omega_1 + \omega_2 + \dots + \omega_n \quad \left( \begin{array}{l} \text{cum gain at time } n \\ \text{given that } \omega \text{ is what} \\ \text{happens} \end{array} \right)$$

Fix  $\omega$ , and look at the path.

