

Limit theorems for expectations

Recall:

convergence in prob.  
↓Lebesgue's Dominated Convergence Theorem: If  $X_n$  converges a.s.to  $X$  and there is an integrable r.v.  $\forall$  such that $|X_n| \leq \forall$  a.s. for all  $n$ , then

$$\lim_{n \rightarrow \infty} E[X_n] = E[X]$$

Lemma: Assume that  $\{x_n\}$  is a bounded sequence of real numbers with a subsequence  $\{y_n\}$  converging to a number  $a$ . If  $\{x_n\}$  does not converge to  $a$ , then  $\{x_n\}$  has another subsequence  $\{z_n\}$  converging to a number  $b \neq a$ .

Proof: Assume that  $\{x_n\}$  does not converge to  $a$ . Then there is a  $\epsilon > 0$  such that  $|x_n - a| \geq \epsilon$  for arbitrarily large  $n$ . Hence there is a subsequence  $\{y_n\}$  of  $\{x_n\}$  such that  $|y_n - a| \geq \epsilon$  for all  $n$ . Since  $\{y_n\}$  is bounded it has a convergent subsequence  $\{z_n\}$  converging to a number  $b$ . Since  $|z_n - a| \geq \epsilon$  for all  $n$ ,  $b \neq a$ .

DCT for convergence in prob.: Assume  $\{X_n\}$  is converging to  $X$  in probability and that there is an integrable r.v.  $\forall$  such that  $|X_n| \leq \forall$ . Then

$$\lim_{n \rightarrow \infty} E[X_n] = E[X]$$

Proof: We know that  $\{X_n\}$  has a subsequence  $\{Y_n\}$  converging to  $X$  a.s. Hence

$$\lim_{n \rightarrow \infty} E[Y_n] = E[X]$$

Assume for contradiction that  $E[X_n]$  does not converge to  $E[X]$ . By the lemma, there is another subsequence  $\{E[Z_n]\}$  converging to a  $b \neq E[X]$ .

Now since  $Z_n$  converges to  $X$  in prob.,  $Z_n$  has a subsequence  $\{u_n\}$  converging to  $X$  a.s. By the original form of the DCT, we get  $E[u_n] \rightarrow E[X]$ .

But  $\{E[u_n]\}$  is a subsequence of  $E[Z_n] \rightarrow b$ , and  $E[u_n] \rightarrow b \neq E[X]$ . This is a contradiction as  $E[u_n]$  could both converge to  $E[X]$  and  $b \neq E[X]$ .

Hence  $E[X_n] \rightarrow E[X]$ 

Q.E.D.

Observation: As we are working with prob. spaces, all constant functions are integrable, and we may often use a constant  $\gamma$  in DCT. This is often called the Bounded C.T.

Corollary: Assume that  $\{X_n\}$  converges to  $X$  either a.s. or in probability. Assume also that there is an integrable  $\gamma$  such that  $|X_n| \leq \gamma$  a.s. Then  $X_n \rightarrow X$  in  $L^1$ -norm i.e.

$$E[|X - X_n|] \rightarrow 0.$$

Proof: Note that  $|X - X_n| \leq |X| + |X_n| \leq 2\gamma$  a.s.

Applying the DCT, we get

$$\lim_{n \rightarrow \infty} E[\underbrace{|X - X_n|}_{\leq 2\gamma}] = E[0] = 0$$

## Laws of large number

Idea: If  $\{X_n\}$  is a sequence of independent, identically distributed r.v. with mean  $\mu$ , then

$$\frac{X_1 + X_2 + \dots + X_n}{n}$$

$\xrightarrow{\mu}$   
 $\nearrow$   
 in what sense

$X_n \sim \mu$

Assume that  $E[X_i] = 0$ :

### Weak laws of large numbers

Conditions  $\Rightarrow \frac{X_1 + \dots + X_n}{n} \rightarrow 0$  in prob.

### Strong laws of large numbers

Conditions  $\Rightarrow \frac{X_1 + \dots + X_n}{n} \rightarrow 0$  a.s.

(A) Weak law of large numbers: Assume that  $\{X_n\}$  is an independent sequence of random variables with  $E[X_n] = 0$  and  $E[X_n^2] \leq \sigma^2$  for all  $n$ . Then

$$\frac{X_1 + \dots + X_n}{n} \rightarrow 0 \text{ in prob.}$$

Proof: Let  $S_n = X_1 + X_2 + \dots + X_n$ , must prove  $\frac{S_n}{n} \rightarrow 0$  in prob.

By Chebyshev's inequality with  $p=2$ :

$$P\left[\left|\frac{S_n}{n}\right| \geq \varepsilon\right] \leq \frac{1}{\varepsilon^2} E\left[\left(\frac{S_n}{n}\right)^2\right]$$

$$= \frac{1}{n^2 \varepsilon^2} E\left[(X_1 + X_2 + \dots + X_n)^2\right]$$

$$= \frac{1}{n^2 \varepsilon^2} E\left[X_1^2 + X_2^2 + \dots + X_n^2\right]$$

$$\leq \frac{1}{n^2 \varepsilon^2} n \sigma^2 = \frac{\sigma^2}{n \varepsilon^2} \rightarrow 0 \text{ as } n \rightarrow \infty.$$

Hence  $\frac{S_n}{n} \rightarrow 0$  in prob.

Strong law of large numbers

Observation: Assume  $E[X^4] < \infty$ . Then  $E[X^2] \leq E[X^4]^{1/2}$

since by Schwarz' inequality

$$E[X^2] = E[1 \cdot X^2] \leq \underbrace{(E[1^2])^{1/2}}_1 E[(X^2)^2]^{1/2} = E[X^4]^{1/2}$$

Cauchy's strong law of large numbers: Assume that  $\{X_n\}$  is an independent sequence of v.v. with mean 0 and  $E[X_n^4] \leq M$  for all  $n$ . Then

$$\frac{S_n}{n} = \frac{X_1 + \dots + X_n}{n} \rightarrow 0 \text{ a.s.}$$

Proof:  $E[S_n^4] = E[(X_1 + X_2 + \dots + X_n)^4]$

$$\begin{aligned} E\left[\sum_{i,j,k,l} X_i X_j X_k X_l\right] &= \sum_{i,j,k,l} E[X_i X_j X_k X_l] \quad \text{X, diff.} \\ &= \sum_{i=1}^n E[X_i^4] + 3 \sum_{i \neq j} E[X_i^2 X_j^2] \\ &= \sum_{i=1}^n E[X_i^4] + 3 \sum_{i \neq j} \underbrace{E[X_i^2]}_{M^{1/2}} \cdot \underbrace{E[X_j^2]}_{M^{1/2}} = M \\ &\leq nM + 3n(n-1)M \\ &= nM + 3n^2M - 3nM \leq 3n^2M \end{aligned}$$

$$\begin{aligned} E[X_i X_j X_k X_l] &= E[X_i] E[X_j X_k X_l] \\ &= 0 \end{aligned}$$

$$\begin{aligned} E[X_i^2] &\leq E[X_i^4]^{1/2} \\ &\leq M^{1/2} \end{aligned}$$

Hence:  $E[S_n^4] \leq 3n^2M$ .

Chebyshev's inequality with  $p=4$  and  $\lambda = \frac{1}{n^{1/8}}$ :

$$\begin{aligned} P\left[\left|\frac{S_n}{n}\right| \geq \frac{1}{n^{1/8}}\right] &\leq \frac{1}{\left(\frac{1}{n^{1/8}}\right)^4} E\left[\left|\frac{S_n}{n}\right|^4\right] \\ &= n^{1/2} \frac{1}{n^4} E[S_n^4] \leq \frac{n^{1/2}}{n^4} 3n^2M \\ &= \frac{3M}{n^{3/2}} \end{aligned}$$

$$P[|X| \geq \lambda] \leq \frac{1}{\lambda^p} E[|X|^p]$$

$$\left(2 + \frac{1}{2}\right) - 4 = -\frac{3}{2}$$

Calculus:  $\sum \frac{1}{n^{3/2}} < \infty$ .

Borel-Cantelli now tells us that

$$P\left[\limsup_{n \rightarrow \infty} \left\{\omega: \left|\frac{S_n(\omega)}{n}\right| \geq \frac{1}{n^{1/8}}\right\}\right] = 0$$

For all  $\omega$  outside this set of prob. 0, we can only have  $\left|\frac{S_n(\omega)}{n}\right| \geq \frac{1}{n^{1/8}}$  for finitely many  $n$ . Hence

for sufficiently large  $\left|\frac{S_n(\omega)}{n}\right| < \frac{1}{n^{1/8}} \rightarrow 0$ .

Thus  $\left|\frac{S_n(\omega)}{n}\right| \rightarrow 0$  on a set of prob 1, hence

Sec 5.4:  $\sigma$ -algebras as information

Prob space:  $(\Omega, \mathcal{F}, P)$

Assume that  $\mathcal{G}$  is another  $\sigma$ -algebra on  $\Omega$  (usually  $\mathcal{G} \subseteq \mathcal{F}$ )

A function  $X: \Omega \rightarrow \mathbb{R}$  is called  $\mathcal{G}$ -measurable if

$$\{\omega: X(\omega) \leq x\} \in \mathcal{G} \text{ for all } x.$$

Prop: If  $X$  is  $\mathcal{G}$ -measurable, then  $X^{-1}(B) \in \mathcal{G}$  for all Borel sets  $B$ .

Definition: If  $X: \Omega \rightarrow \mathbb{R}$ , we define the  $\sigma$ -algebra  $\sigma(X)$  generated by  $X$  to be the smallest  $\sigma$ -algebra such that  $\{\omega: X(\omega) \leq x\} \in \sigma(X)$  (in other words,  $\sigma(X)$  is the  $\sigma$ -algebra generated by the family  $\{\{\omega: X(\omega) \leq x\}\}_{x \in \mathbb{R}}$ .)

Prop:  $X$  is measurable w.r.t  $\sigma(X)$ .

Definition: If  $\{X_i\}_{i \in I}$  is a family of functions  $X_i: \Omega \rightarrow \mathbb{R}$ , then the  $\sigma$ -algebra  $\sigma(\{X_i\}_{i \in I})$  generated by  $\{X_i\}_{i \in I}$  is the smallest  $\sigma$ -algebra containing all sets

$$\{\omega: X_i(\omega) \leq x\}$$

for all  $i \in I$  and all  $x \in \mathbb{R}$ .