

0-1-laws

Borel-Cantelli lemma, part II: If $\{\mathbb{F}_n\}$ is independent, then

$P[\limsup \mathbb{F}_n]$ is either 0 or 1

Assume that $\{\mathbb{F}_n\}$ is a sequence of ind. r.v.

First σ -algebras: $\mathbb{F}_n = \sigma(\mathbb{X}_1, \mathbb{X}_2, \dots, \mathbb{X}_n)$

Future σ -algebras: $\mathbb{F}_n^* = \sigma(\mathbb{X}_{n+1}, \mathbb{X}_{n+2}, \dots)$

Note that \mathbb{F}_n and \mathbb{F}_n^* are independent.

Γ -algebra of infinite: $\mathbb{F}_\infty^* = \bigcap_{n \in \mathbb{N}} \mathbb{F}_n^*$

An event in \mathbb{F}_∞^* is called a tail event.

Example: $\limsup \mathbb{F}_n \in \Delta$ is a tail event

$\{\omega: \sum \mathbb{F}_n(\omega) \text{ converges}\} \subset \Delta$

$\{\omega: \sum \mathbb{F}_n(\omega) < \infty\}$ is unconditionally not a tail event.

Kolmogorov/Borel 0-1 law: If $\{\mathbb{F}_n\}$ is an independent

sequence of random variables and Δ is a tail event
(i.e. $\Delta \in \mathbb{F}_\infty^*$), then $P(\Delta)$ is either 0 or 1.

$$\chi^2 = x$$

Proof: Let

$$\mathcal{K} = \{\Gamma: P(\Gamma \cap \Delta) = P(\Gamma)P(\Delta)\}.$$

Note that $\mathbb{F}_n \subseteq \mathcal{K}$ because if $\Gamma \in \mathbb{F}_n$, then Γ is independent of Δ since $\Delta = \mathbb{F}_\infty^* = \bigcap_{n=1}^\infty \mathbb{F}_n^* \subseteq \mathbb{F}_n^*$, and \mathbb{F}_n is independent of \mathbb{F}_n^* . Thus

$$\bigcup_{n \in \mathbb{N}} \mathbb{F}_n \subseteq \mathcal{K}$$

Now $\bigcup_{n \in \mathbb{N}} \mathbb{F}_n$ is an algebra, and hence

$$\sigma(\bigcup_{n \in \mathbb{N}} \mathbb{F}_n) = M(\bigcup_{n \in \mathbb{N}} \mathbb{F}_n). \quad (\text{the smallest monotone class containing } \bigcup_{n \in \mathbb{N}} \mathbb{F}_n)$$

If we can prove that \mathcal{K} is a monotone class, then

$$\sigma(\bigcup_{n \in \mathbb{N}} \mathbb{F}_n) = M(\bigcup_{n \in \mathbb{N}} \mathbb{F}_n) \subseteq \mathcal{K}.$$

Check that \mathcal{K} is a monotone class:

(i) Let $\{\Gamma_n\}$ be an increasing sequence in \mathcal{K} . We must show that $\bigcup_{n \in \mathbb{N}} \Gamma_n \in \mathcal{K}$.

$$\begin{aligned} P\left[\left(\bigcup_{n \in \mathbb{N}} \Gamma_n\right) \cap \Delta\right] &= P\left[\bigcup_{n \in \mathbb{N}} (\Gamma_n \cap \Delta)\right] \stackrel{\text{increasing}}{=} \lim_{n \rightarrow \infty} P[\Gamma_n \cap \Delta] \\ &= \lim_{n \rightarrow \infty} P(\Gamma_n)P(\Delta) \stackrel{\text{cont.}}{=} P\left(\bigcup_{n \in \mathbb{N}} \Gamma_n\right)P(\Delta). \end{aligned}$$

(ii) Let now $\{\Gamma_n\}$ be a decreasing sequence in \mathcal{K} :

$$\begin{aligned} P\left(\bigcap_{n \in \mathbb{N}} \Gamma_n \cap \Delta\right) &= P\left(\bigcap_{n \in \mathbb{N}} (\Gamma_n \cap \Delta)\right) = \lim_{n \rightarrow \infty} P(\Gamma_n \cap \Delta) \\ &= \lim_{n \rightarrow \infty} P(\Gamma_n)P(\Delta) = P\left(\bigcap_{n \in \mathbb{N}} \Gamma_n\right)P(\Delta). \end{aligned}$$

This establishes that M is a monotone class and hence

$$\sigma(\bigcup_{n \in \mathbb{N}} \mathbb{F}_n) \subseteq \mathcal{K}.$$

Since $\Delta \in \mathbb{F}_\infty^* = \bigcap_{n \in \mathbb{N}} \mathbb{F}_n^* \subseteq \sigma(\bigcup_{n \in \mathbb{N}} \mathbb{F}_n) \subseteq \mathcal{K}$.

Thus Δ is independent of itself and

$$P(\Delta) = P(\Delta \cap \Delta) = P(\Delta)P(\Delta) = P(\Delta)^2$$

Complex-valued random variables

$Z = X + iY$, where X and Y are real-valued random variables.

Z is integrable when both X and Y are, and we then define the expectation by

$$E[Z] = E[X] + iE[Y]$$

Prop: Z is integrable iff $|Z|$ is integrable.

Prop: If $\alpha, \beta \in \mathbb{C}$ and Z, W are integrable complex-valued random variables, then

$$E[\alpha Z + \beta W] = \alpha E[Z] + \beta E[W]$$

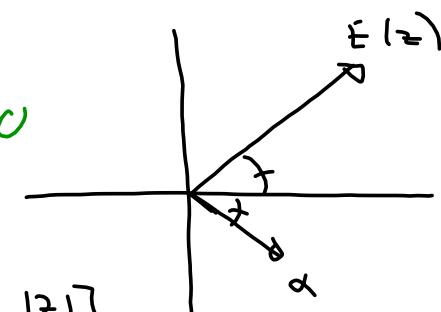
Prop: If Z is integrable, then

$$|E[Z]| \leq E[|Z|]$$

Proof: Since $E[Z]$ is a complex number, there is $\alpha \in \mathbb{C}$ such that $|\alpha|=1$ and $\underline{\alpha E(Z)} = |E(Z)|$

Then

$$\begin{aligned} |E(Z)| &= \alpha E(Z) = E(\underline{\alpha Z}) \\ &= E[\underline{u} + i\underline{v}] = E[\underline{u}] + iE[\underline{v}] \\ &\leq E[|u|] \leq E[|\alpha z|] = E[|\alpha| |z|] \\ &= E[|z|] \end{aligned}$$



Characteristic function

If X is a random variable, we define its characteristic function

$$\varphi(t) = E[e^{itX}] = E[\cos tX + i \sin tX] \\ = E[\cos tX] + i E[\sin tX]$$

Compare the moment generating function

$$M(t) = E[e^{tX}]$$

integrability problems.

Prop: $\varphi(0) = 1$, $|\varphi(t)| \leq 1$.

Proof: $\varphi(0) = E[e^{i0X}] = E[e^0] = 1$.

$$|\varphi(t)| = |E[e^{itX}]| \leq E\left[1_{e^{itX}}\right] = E[1] = 1$$

Prop: φ is uniformly continuous.

$$\begin{aligned} \underline{|\varphi(t+s) - \varphi(t)|} &= |\underline{E[e^{i(t+s)X} - e^{itX}]}| \\ &\leq E\left[\underbrace{|e^{itX}(e^{isX} - 1)|}_{\text{abs val. 1}}\right] = \frac{E[|e^{isX} - 1|]}{2} = \eta(s) \end{aligned}$$

If we can prove that $\eta(s) \rightarrow 0$ as $s \rightarrow 0$, then φ must be uniformly continuous. Why? Given an

$\epsilon > 0$, we can choose S so small that $|\eta(s)| < \epsilon$ for all $|s| < S$. But this means that if $|s| < S$, then for all t , $|\varphi(t+s) - \varphi(t)| < \eta(s) < \epsilon$ when $|s| < S$.

So how do we prove that $\eta(s) \rightarrow 0$? Assume for contradiction that $\eta(s) \not\rightarrow 0$. Then there is a sequence $\{s_n\} \rightarrow 0$ such that $\eta(s_n) \not\rightarrow 0$. But

$$\eta(s_n) = E\left[\overbrace{|e^{is_n X} - 1|}^{\approx 1}\right] \xrightarrow{\text{DCT}} E[0] = 0.$$

Contradiction, $\approx \frac{1}{2}$

$\eta(s) \rightarrow 0$. QED.

Prop: φ is a positive definite function, i.e. for all real numbers t_1, t_2, \dots, t_n and all complex numbers z_1, z_2, \dots, z_n ,

then $\sum_{i,j=1}^n \varphi(t_i - t_j) z_i \bar{z}_j \geq 0.$ $M_{ij} = \varphi(t_i - t_j)$

Prove: $\sum_{i,j=1}^n \varphi(t_i - t_j) z_i \bar{z}_j$ M is a positive matrix

$$= \sum_{i,j=1}^n E[\overline{e^{i(t_i-t_j)\bar{X}}}] z_i \bar{z}_j$$

$$= E\left[\sum_{i,j=1}^n e^{i(t_i-\bar{t}_j)\bar{X}} (\overline{e^{i(t_i-\bar{t}_j)\bar{X}}}) z_i \bar{z}_j\right]$$

$$= E\left[\sum_{i,j=1}^n (z_i e^{i(t_i-\bar{t}_j)\bar{X}})(\overline{z_j e^{i(t_i-\bar{t}_j)\bar{X}}})\right]$$

$$= E\left[(\sum_{i=1}^n z_i e^{i(t_i-\bar{t})\bar{X}})(\sum_{j=1}^n \overline{z_j e^{i(t_i-\bar{t})\bar{X}}})\right]$$

$$E\left[\left| \sum_{i=1}^n z_i e^{i(t_i-\bar{t})\bar{X}} \right|^2 \right] \geq 0.$$

Theorem: Assume that $E[|\bar{X}|^k] < \infty$. Then φ is k times differentiable and

$$\varphi(t) = \sum_{j=0}^k \frac{1}{j!} E[\bar{X}^j] (it)^j + \sigma(t^k)$$

In particular:

$$\varphi^{(j)}(0) = i^j E[\bar{X}^j]$$

Idea behind the proof:

$$\varphi'(t) = \frac{d}{dt} E[e^{it\bar{X}}] \quad \text{circled } \bar{X} \quad E\left[\frac{d}{dt} e^{it\bar{X}}\right]$$

$$= E[i\bar{X} e^{it\bar{X}}] = i E[\bar{X} e^{it\bar{X}}]$$

$$\varphi^{(j)}(t) = \underbrace{i^j E[\bar{X}^j e^{it\bar{X}}]}_{j \leq k} \quad j \leq k.$$

Taylor expansion of order k :

$$\begin{aligned} \varphi(t) &= \sum_{j=0}^k \frac{\varphi^{(j)}(0)}{j!} + \sigma(t^k) \\ &= \sum_{j=0}^k \frac{E[\bar{X}^j]}{j!} (it)^j + \sigma(t^k) \quad \text{formula in known.} \end{aligned}$$

We need to check that we can pull $\frac{d}{dt}$ inside the expectation.

$$\varphi'(t) = \lim_{h \rightarrow 0} \frac{\varphi(t+h) - \varphi(t)}{h} = \lim_{h \rightarrow 0} \frac{1}{h} E\left[e^{i(t+h)\bar{X}} - e^{it\bar{X}}\right]$$

$$= \lim_{h \rightarrow 0} E\left[e^{it\bar{X}} \frac{e^{ih\bar{X}} - 1}{h}\right]$$

$$= \lim_{h \rightarrow 0} E\left[e^{it\bar{X}} \left[\frac{\cosh h\bar{X} - 1}{h\bar{X}} + i \frac{\sinh h\bar{X}}{h\bar{X}} \right]\right] \quad \cos 0 = 1, \sin 0 = 0$$

$$= \lim_{h \rightarrow 0} E\left[e^{it\bar{X}} \left[\underbrace{-\frac{\sinh h\bar{X}}{h\bar{X}}}_{0} \cdot \bar{X} + \underbrace{i \frac{\cosh h\bar{X}}{h\bar{X}} \cdot \bar{X}}_{1} \right]\right]$$

$$\stackrel{\Delta T}{=} E\left[e^{it\bar{X}} i\bar{X}\right] = i E[\bar{X} e^{it\bar{X}}]$$

higher order derivative by induction in a similar manner.