

**UNIVERSITY OF OSLO**  
**Faculty of Mathematics and Natural Sciences**

Solutions of the Examination, STK-MAT3710/4710 - Probability Theory.

**[Problem 1]** (9 points) **Solution:** By the first Borel-Cantelli lemma,  $P(X_n \neq Y_n, i.o.) = 0$ . Hence,  $X_n = Y_n$  for all, just not for finitely many values of  $n$ , almost surely. Off an event of probability zero, the sequences are identical for all large  $n$ . Hence, both converge or diverge together.

**[Problem 2]**

**a)** (9 points) **Solution:**  $E[X] = -5 \cdot \frac{1}{4} + 3 \cdot \frac{1}{4} + 1 \cdot \frac{2}{4} = 0$ .

$$E[X^2] = 25 \cdot \frac{1}{4} + 9 \cdot \frac{1}{4} + 1 \cdot \frac{2}{4} = 9.$$

Then, by Taylor expansion formula of the characteristic function:

$$\phi_X(t) = 1 - \frac{9}{2}t^2 + o(t^2).$$

If you have written :

$$\phi_X(t) = \sum_{k=0}^n \frac{1}{k!} E[X^k] (it)^k + o(t^n) = \sum_{k=0}^n \frac{(-5)^k + 3^k + 2}{4k!} (it)^k + o(t^n)$$

is also accepted.

**b)** (9 points) **Solution:**

$$\begin{aligned} \phi_{S_n}(t) &= E[e^{itS_n}] = E[e^{it(\frac{X_1+X_2+\dots+X_n}{\sqrt{n}})}] \\ &= E[e^{\frac{it}{\sqrt{n}}X_1} \cdot e^{\frac{it}{\sqrt{n}}X_2} \dots e^{\frac{it}{\sqrt{n}}X_n}] \\ &= E[e^{\frac{it}{\sqrt{n}}X_1}] \cdot E[e^{\frac{it}{\sqrt{n}}X_2}] \dots E[e^{\frac{it}{\sqrt{n}}X_n}], \quad (X_n \text{'s are independent}) \\ &= (E[e^{\frac{it}{\sqrt{n}}X_1}])^n \quad (X_n \text{'s are identical distributed}) \\ &= (\phi_{X_1}(\frac{t}{\sqrt{n}}))^n = (\phi_X(\frac{t}{\sqrt{n}}))^n \\ &= (1 - \frac{9}{2} \frac{t^2}{n} + o(\frac{t^2}{n}))^n. \end{aligned}$$

Last step runs since  $X_n$ 's and  $X$  have the same distribution (so they have same characteristic function).

**c)** (9 points) **Solution:** Let us define  $z_n = \frac{-9}{2}t^2 + o(t^2)$ , hence by the Lemma 6.34,

$$\lim_{n \rightarrow \infty} (1 + \frac{z_n}{n})^n = e^{\frac{-9}{2}t^2}, \quad (z_n \rightarrow \frac{-9}{2}t^2, (n \rightarrow \infty)).$$

Since  $e^{-\frac{9}{2}t^2}$  is continuous at  $t = 0$  and  $\phi_{S_n}(t) \rightarrow e^{-\frac{9}{2}t^2}$  ( $n \rightarrow \infty$ ), which is the characteristic function of a random variable with Normal distribution, by Levy continuity theorem,  $S_n$  converges in distribution to a normal distribution with  $N(0, 9)$ .

**[Problem 3]**

**a) (9 points) Solution:**

$$\begin{aligned} E[X_j X_n] &= E[E[X_j X_n | \mathcal{F}_n]] \quad (n < j, X_n \text{ is } \mathcal{F}_n\text{-measurable}) \\ &= E[X_n E[X_j | \mathcal{F}_n]] = E[X_n^2] \quad (X_n \text{ is a martingale wrt } \mathcal{F}_n). \end{aligned}$$

First step runs via the rule of conditional expectation of a random variable is equal to its expectation.

**b) (5 points) Solution:**

$$E[(X_m - X_r)X_n] = E[X_m X_n] - E[X_r X_n] = E[X_n^2] - E[X_n^2] = 0 \quad \text{by (a) and expectation is linear}$$

**c) (9 points) Solution:**

$$\begin{aligned} E[(X_m - X_r)^2 | \mathcal{F}_n] &= E[X_m^2 | \mathcal{F}_n] - 2E[X_m X_r | \mathcal{F}_n] + E[X_r^2 | \mathcal{F}_n] \quad \text{conditional expectation is linear} \\ &= E[X_m^2 | \mathcal{F}_n] - E[X_r^2 | \mathcal{F}_n], \quad (\text{by 1}). \end{aligned}$$

and

$$\begin{aligned} E[X_m X_r | \mathcal{F}_n] &= E[E[X_m X_r | \mathcal{F}_r] | \mathcal{F}_n] \quad \text{by tower property, } n < r, \text{ and } r < m \\ &= E[X_r E[X_m | \mathcal{F}_r] | \mathcal{F}_n] \quad (X_r \text{ is } \mathcal{F}_r\text{-measurable}) \\ &= E[X_r X_r | \mathcal{F}_n] = E[X_r^2 | \mathcal{F}_n] \quad (X_n \text{ is a martingale wrt } \mathcal{F}_n) \quad (1) \end{aligned}$$

**d) (9 points) Solution:**

By taking expectation of both sides of the equation obtained in part (c) and by the linearity of conditional expectation, we get:

$$\begin{aligned} 0 &\leq E[E[(X_m - X_r)^2 | \mathcal{F}_n]] = E[E[X_m^2 | \mathcal{F}_n] - E[X_r^2 | \mathcal{F}_n]] = E[E[(X_m^2 - X_r^2) | \mathcal{F}_n]] \\ 0 &\leq E[(X_m - X_r)^2] = E[X_m^2 - X_r^2], \quad r < m. \end{aligned}$$

Observe that since  $E[X_m^2 - X_r^2] \geq 0$  for  $r < m$ ,  $E[X_m^2] \geq E[X_r^2]$ , for all  $r < m$ . Hence,  $\{E(X_n^2) : n \geq 1\}$  is nondecreasing and by hypothesis it is bounded. Therefore,  $\{E(X_n^2) : n \geq 1\}$  is convergent. Then,

$$\lim_{m \rightarrow \infty} E[X_m^2] = \lim_{r \rightarrow \infty} E[X_r^2], \quad (\text{If limit exists, it is unique.})$$

By Squeeze Theorem,  $\{X_n\}_{n \in \mathbb{N}_0}$  is Cauchy convergent.

**[Problem 4]**

**a) (9 points) Solution:**

Both  $S$  and  $T$  are bounded First Hitting Times, hence they are stopping times. The natural filtration generated by  $X_n$  can be considered as a sub-filtration of  $\mathcal{F}_n$ , hence both  $S$  and  $T$  are  $\mathcal{F}_n$ -measurable and both  $\{T \leq n\} \in \mathcal{F}_n$  and  $\{S \leq n\} \in \mathcal{F}_n$ . Then,

$$\{\gamma = \max(T, S) \leq n\} = \{T \leq n\} \cap \{S \leq n\} \in \mathcal{F}_n,$$

$\sigma$ -algebras are closed under finite intersection. Since, both  $S$  and  $T$  are bounded, maximum of their upper bound is an upper bound for  $\gamma$ .

**b) (10 points) Solution:**

Let us call  $\max(X_n, K) = X_n^+ = \Phi(X_n)$ , since  $\Phi$  is a convex function and  $X_n$  is an  $\mathcal{F}_n$ -martingale; hence by Jensen's inequality:

$$E[X_{n+1}^+ | \mathcal{F}_n] = E[\Phi(X_{n+1}) | \mathcal{F}_n] \geq \Phi(E[X_{n+1} | \mathcal{F}_n]) = \Phi(X_n) = X_n^+.$$

Hence,  $X_n^+$  is an  $\mathcal{F}_n$ -submartingale.

Observe that since  $K < \lambda_2$ , replacing  $X_n^+$  with  $X_n$  in the definitions of  $S$  or  $T$  does not matter. By hypothesis,  $\gamma \geq S$  and both are bounded stopping-times, hence, by Theorem 9.9, since  $X_n^+$  is an  $\mathcal{F}_n$ -submartingale,  $(X_\gamma^+, X_S^+)$  is an  $(\mathcal{F}_S, \mathcal{F}_\gamma)$  submartingale. Then, by applying expectation to both sides of the inequality:

$$\begin{aligned} E[X_\gamma^+ | \mathcal{F}_S] &\geq X_S^+ \\ E[E[X_\gamma^+ | \mathcal{F}_S]] &\geq E[X_S^+] \\ E[X_\gamma^+] &\geq E[X_S^+] = L. \end{aligned}$$

**c) (13 points) Solution:**

Since  $K = 0$ ,  $U_m = E[X_{m+n}^+ | \mathcal{F}_n]$ ,  $m \geq 0$  is a nonnegative sequence of random variables. (Remember that conditional expectation of a random variable is a random variable as well.) Moreover, by Tower property:

$$U_{m+1} = E[X_{m+n+1}^+ | \mathcal{F}_n] = E[E[X_{m+n+1}^+ | \mathcal{F}_{n+m}] | \mathcal{F}_n] \geq E[X_{n+m}^+ | \mathcal{F}_n] = U_m,$$

since  $X_n^+$  is an  $\mathcal{F}_n$ -submartingale. We see that for all  $m \geq 0$ ,  $U_{m+1} \geq U_m$ , then  $U_m$  is nondecreasing (or increasing). So we are allowed to use Monotone Convergence Theorem for conditional expectation. Then,

$$\begin{aligned} E[M_{n+1} | \mathcal{F}_n] &= E[\lim_{m \rightarrow \infty} E[X_{m+n+1}^+ | \mathcal{F}_{n+1}] | \mathcal{F}_n] \\ &= \lim_{m \rightarrow \infty} E[E[X_{m+n+1}^+ | \mathcal{F}_{n+1}] | \mathcal{F}_n] \\ &= \lim_{m \rightarrow \infty} E[X_{m+n+1}^+ | \mathcal{F}_n] = M_n, \quad \text{by Tower property.} \end{aligned}$$

Note that last step runs also by the fact that if limit exists, it has to be unique and all subsequences approach to same point.