

STK-MAT3710/4710: Solution to Exam 2022

Problem 1 a) We have

$$\begin{aligned}\phi_a(t) &= E[e^{itY_a}] = e^{it(1+a)} \cdot \frac{1}{2} + e^{it(-1+a)} \cdot \frac{1}{2} \\ &= e^{iat} \frac{e^{it} + e^{-it}}{2} = e^{ita} \cos t\end{aligned}$$

b) Using the independence of $X_1^{(n)}, X_2^{(n)}, \dots, X_n^{(n)}$, we get

$$\begin{aligned}\phi_{S_n}(t) &= E[e^{itS_n}] = E\left[e^{it \frac{X_1^{(n)} + X_2^{(n)} + \dots + X_n^{(n)}}{\sqrt{n}}}\right] \\ &= E\left[e^{i \frac{t}{\sqrt{n}} X_1^{(n)}}\right] \cdot E\left[e^{i \frac{t}{\sqrt{n}} X_2^{(n)}}\right] \cdot \dots \cdot E\left[e^{i \frac{t}{\sqrt{n}} X_n^{(n)}}\right] \\ &= \left[\phi_{\frac{1}{\sqrt{n}}}\left(\frac{t}{\sqrt{n}}\right)\right]^n = \left[e^{i \frac{t}{\sqrt{n}} \cdot \frac{1}{\sqrt{n}}} \cos \frac{t}{\sqrt{n}}\right]^n = e^{it} \left[\cos \frac{t}{\sqrt{n}}\right]^n\end{aligned}$$

c) Using the Taylor expansion $\cos x = 1 - \frac{x^2}{2} + o(x^2)$, we get

$$\lim_{n \rightarrow \infty} \left[\cos \frac{t}{\sqrt{n}}\right]^n = \lim_{n \rightarrow \infty} \left[1 - \frac{t^2}{2n} + o\left(\frac{t^2}{n}\right)\right]^n = e^{-\frac{t^2}{2}}$$

by Lemma 6.34. This means that

$$\lim_{n \rightarrow \infty} \phi_{S_n}(t) = e^{it - \frac{t^2}{2}}$$

which is the characteristic function of an $\mathcal{N}(1, 1)$ random variable. By Lévy's Continuity Theorem, S_n converges in distribution to a normal distribution with mean 1 and variance 1.

Problem 2 Let A_n be the event that you win on day n . Then the A_n 's are independent (we must assume) and $P(A_n) = \frac{1}{99+n}$. The series

$$\sum_{n=1}^{\infty} P(A_n) = \sum_{n=1}^{\infty} \frac{1}{99+n}$$

diverges (it is the tail of the divergent harmonic series $\sum_{n=1}^{\infty} \frac{1}{n}$), and thus the probability of winning infinitely many times is 1 by the Borel-Cantelli Lemma.

Problem 3 Let $\mathcal{M} = \{\Lambda : \int_{\Lambda} X dP = \int_{\Lambda} Y dP\}$; we must prove that \mathcal{M} is closed under increasing unions and decreasing intersections.

Assume first that $\{\Lambda_n\}$ is an increasing sequence of sets in \mathcal{M} with union Λ . Since the sequence $\{\mathbf{1}_{\Lambda_n} X\}$ is dominated by the integrable function $|X|$, the Dominated Convergence Theorem tells us that

$$\int_{\Lambda} X dP = \int \mathbf{1}_{\Lambda} X dP = \int \lim_{n \rightarrow \infty} \mathbf{1}_{\Lambda_n} X dP \stackrel{DCT}{=} \lim_{n \rightarrow \infty} \int \mathbf{1}_{\Lambda_n} X dP = \lim_{n \rightarrow \infty} \int_{\Lambda_n} X dP$$

and similarly,

$$\int_{\Lambda} Y dP = \lim_{n \rightarrow \infty} \int_{\Lambda_n} Y dP.$$

This shows that $\int_{\Lambda} X dP = \int_{\Lambda} Y dP$, and hence $\Lambda \in \mathcal{M}$.

Assume now that $\{\Lambda_n\}$ is a decreasing sequence of sets in \mathcal{M} with intersection Λ . As above, the sequence $\{\mathbf{1}_{\Lambda_n} X\}$ is dominated by the integrable function $|X|$, and the Dominated Convergence Theorem gives

$$\int_{\Lambda} X dP = \int \mathbf{1}_{\Lambda} X dP = \int \lim_{n \rightarrow \infty} \mathbf{1}_{\Lambda_n} X dP \stackrel{DCT}{=} \lim_{n \rightarrow \infty} \int \mathbf{1}_{\Lambda_n} X dP = \lim_{n \rightarrow \infty} \int_{\Lambda_n} X dP$$

and similarly,

$$\int_{\Lambda} Y dP = \lim_{n \rightarrow \infty} \int_{\Lambda_n} Y dP.$$

This shows that $\int_{\Lambda} X dP = \int_{\Lambda} Y dP$, and hence $\Lambda \in \mathcal{M}$.

Problem 4 Let $Y_n = \mathbf{1}_{A_n}$. Then $E[Y_n] = P(A_n)$, and the random variables $Z_n = Y_n - P(A_n)$ have mean 0. Since the Z_n 's are independent and obviously have bounded fourth moments, Cantelli's Strong Law of Large Numbers tells us that $\frac{Z_1 + Z_2 + \dots + Z_n}{n}$ converges to 0 a.s. But $X_n = Y_1 + Y_2 + \dots + Y_n$ and $P_n = P(A_1) + P(A_2) + \dots + P(A_n)$, and thus

$$\frac{X_n - P_n}{n} = \frac{Z_1 + Z_2 + \dots + Z_n}{n} \rightarrow 0 \quad \text{a.s.}$$

Problem 5

a) To show that

$$Z(\omega) = \sum_{n=1}^{\infty} \mathbf{1}_{[Y=n]}(\omega)(X_1(\omega) + X_2(\omega) + \dots + X_n(\omega)),$$

we first observe that if $Y(\omega) = 0$, then both sides are 0. Next we observe that if $Y(\omega) = n$ for $n > 0$, then both sides equal $X_1(\omega) + X_2(\omega) + \dots + X_n(\omega)$, and hence the expressions are equal for all ω .

Taking expectations on both sides of the equation above, we get

$$E[Z] = E \left[\sum_{n=1}^{\infty} \mathbf{1}_{[Y=n]}(X_1 + X_2 + \dots + X_n) \right]$$

Since all terms are positive, the Monotone Convergence Theorem gives

$$\begin{aligned} E[Z] &= E \left[\lim_{m \rightarrow \infty} \sum_{n=1}^m \mathbf{1}_{[Y=n]}(X_1 + X_2 + \dots + X_n) \right] \\ &= \lim_{m \rightarrow \infty} E \left[\sum_{n=1}^m \mathbf{1}_{[Y=n]}(X_1 + X_2 + \dots + X_n) \right] \end{aligned}$$

Using the independence of Y and the X_i 's, we see that this equals

$$\lim_{m \rightarrow \infty} \left[\sum_{n=1}^m E[\mathbf{1}_{[Y=n]}] E[X_1 + X_2 + \dots + X_n] \right]$$

$$\begin{aligned}
&= \lim_{m \rightarrow \infty} \left[\sum_{n=1}^m P[Y = n](\mu_1 + \mu_2 + \cdots + \mu_n) \right] \\
&= \sum_{n=1}^{\infty} P[Y = n](\mu_1 + \mu_2 + \cdots + \mu_n)
\end{aligned}$$

and hence

$$E[Z] = \sum_{n=1}^{\infty} (\mu_1 + \mu_2 + \cdots + \mu_n) P[Y = n]$$

b) Using the Monotone Convergence Theorem for conditional expextations, we get

$$\begin{aligned}
E[Z|\mathcal{G}] &= E \left[\sum_{n=1}^{\infty} \mathbf{1}_{[Y=n]}(X_1 + X_2 + \cdots + X_n) | \mathcal{G} \right] \\
&= \lim_{m \rightarrow \infty} E \left[\sum_{n=1}^m \mathbf{1}_{[Y=n]}(X_1 + X_2 + \cdots + X_n) | \mathcal{G} \right] \\
&= \lim_{m \rightarrow \infty} \sum_{n=1}^m E[\mathbf{1}_{[Y=n]}(X_1 + X_2 + \cdots + X_n) | \mathcal{G}]
\end{aligned}$$

Since Y is \mathcal{G} -measurable and the $\{X_n\}$'s are independent of \mathcal{G} , we have

$$\begin{aligned}
E[\mathbf{1}_{[Y=n]}(X_1 + X_2 + \cdots + X_n) | \mathcal{G}] &= \mathbf{1}_{[Y=n]} E[X_1 + X_2 + \cdots + X_n | \mathcal{G}] \\
&= \mathbf{1}_{[Y=n]} (\mu_1 + \mu_2 + \cdots + \mu_n)
\end{aligned}$$

and hence

$$\begin{aligned}
E[Z|\mathcal{G}] &= \lim_{m \rightarrow \infty} \sum_{n=1}^m \mathbf{1}_{[Y=n]} (\mu_1 + \mu_2 + \cdots + \mu_n) \\
&= \sum_{n=1}^{\infty} \mathbf{1}_{[Y=n]} (\mu_1 + \mu_2 + \cdots + \mu_n).
\end{aligned}$$

c) Since Z_k is \mathcal{F}_k -measurable by induction, and $X_i^{(k+1)}$ is independent of \mathcal{F}_k by assumption, we get from b) that

$$E[Z_{k+1} | \mathcal{F}_k] = \sum_{n=1}^{\infty} \mathbf{1}_{[Z_k=n]} (\mu + \mu + \cdots + \mu) = \mu \sum_{n=1}^{\infty} n \mathbf{1}_{[Z_k=n]} = \mu Z_k.$$

Since $Z_k \geq 0$, we see that $E[Z_{k+1} | \mathcal{F}_k] \geq Z_k$ when $\mu > 1$, $E[Z_{k+1} | \mathcal{F}_k] = Z_k$ for $\mu = 1$, and $E[Z_{k+1} | \mathcal{F}_k] \leq Z_k$ when $\mu < 1$. This means that Z_k is a submartingale, a martingale, and a supermartingale according to whether $\mu \geq 1$, $\mu = 1$, or $\mu \leq 1$.

d) Using c), we see that

$$E[Y_{k+1} | \mathcal{F}_k] = \frac{1}{\mu^{k+1}} E[Z_{k+1} | \mathcal{F}_k] = \frac{\mu}{\mu^{k+1}} Z_k = Y_k$$

which shows that $Y_k = \frac{Z_k}{\mu^k}$ is a martingale.

e) We use Theorem 6.31: Since $\{Y_k\}$ is a uniformly integrable martingale, Y_k converges a.s. and in L^1 to an integrable random variable Y_∞ such that the augmented process $Y_0, Y_1, \dots, Y_\infty$ is also a martingale. Hence $E[Y_\infty] = E[Y_0] = E[Z_0] = Z_0 > 0$, which proves that Y_∞ is positive on a set Ω_0 of positive measure. Since $Z_k = \mu^k Y_k$, it follows that Z_k goes to infinity on Ω_0 .