

## STK-MAT3710/4710: Solution to Mandatory Assignment, Fall 2022

**Problem 1.** a) As  $X_n(0) = \sqrt{n}$  for all  $n$ , and  $X(0) = 0$ , the sequence does not converge at 0, and hence it does not converge pointwise.

b) For any  $x > 0$ , we see that  $X_n(x) = 0$  for  $n > \frac{1}{x}$ , and hence  $\lim_{n \rightarrow \infty} X_n(x) = 0$ . This means that convergence only fails on the set  $\{0\}$  which has probability 0, and hence  $\{X_n\}$  converges to 0 a.s.

c) For any  $\epsilon > 0$ , we have

$$\{x \in \omega : |X_n(\omega)| \geq \epsilon\} \subseteq \left[0, \frac{1}{n}\right],$$

and thus

$$P(\{x \in \omega : |X_n(\omega)| \geq \epsilon\}) \leq P\left(\left[0, \frac{1}{n}\right]\right) = \frac{1}{n} \rightarrow 0$$

This shows that the sequence converges to 0 in probability. One may also use b) and the fact that a.s. convergence implies convergence in probability (Proposition 4.5).

d) We have

$$E[|X_n - 0|] = E[X_n] = \sqrt{n} \cdot \frac{1}{n} = \frac{1}{\sqrt{n}} \rightarrow 0$$

which shows that the sequence converges to 0 in  $L^1$ .

e) We have

$$E[|X_n - 0|^2] = E[X_n^2] = (\sqrt{n})^2 \cdot \frac{1}{n} = \frac{n}{n} = 1$$

which shows that the sequence does not converge to 0 in  $L^2$ .

f) The distribution function of the constant random variable  $X = 0$  is

$$F(x) = \begin{cases} 0 & \text{if } x < 0 \\ 1 & \text{if } x \geq 0 \end{cases}$$

The distribution function of  $X_n$  is

$$F_n(x) = \begin{cases} 0 & \text{if } x < 0 \\ 1 - \frac{1}{n} & \text{if } 0 \leq x < \sqrt{n} \\ 1 & \text{if } x \geq \sqrt{n} \end{cases}$$

We see that  $F_n(x) \rightarrow F(x)$  for all  $x$ , and hence  $\{X_n\}$  converges to  $X = 0$  in distribution. (To see this, note that for  $x < 0$ ,  $F_n(x) = F(x) = 0$  and hence the convergence is obvious. For  $x \geq 0$ , we see that when  $n$  gets big,  $x < \sqrt{n}$ , and hence  $F_n(x) = 1 - \frac{1}{n}$ . Consequently,  $\lim_{n \rightarrow \infty} F_n(x) = 1 = F(x)$ .)

**Comment:** Many solves this problem by combining c) and the assertion that convergence in probability implies convergence in distribution. The assertion is correct, but I can't remember that we have covered it.

**Problem 2.** By Lyapounov's second inequality (Corollary 3.23),

$$E[|X - X_n|^p]^{1/p} \leq E[|X - X_n|^q]^{1/q},$$

and hence

$$E[|X - X_n|^p] \leq E[|X - X_n|^q]^{p/q} \rightarrow 0$$

as  $n \rightarrow \infty$ .

**Problem 3.** a) Differentiating, we get

$$\begin{aligned} \phi'(u) &= \frac{2u \cdot (\alpha + u)^2 - (\sigma^2 + u^2) \cdot 2 \cdot (\alpha + u)}{(\alpha + u)^4} \\ &= \frac{2u(\alpha + u) - 2(\sigma^2 + u^2)}{(\alpha + u)^3} = \frac{2u\alpha - 2\sigma^2}{(\alpha + u)^3} \end{aligned}$$

which is 0 for  $u = \frac{\sigma^2}{\alpha}$ . As  $\phi'(u) < 0$  when  $u < \frac{\sigma^2}{\alpha}$ , and  $\phi'(u) > 0$  when  $u > \frac{\sigma^2}{\alpha}$ , we see that  $u = \frac{\sigma^2}{\alpha}$  is the minimum point. The minimum value is

$$\phi\left(\frac{\sigma^2}{\alpha}\right) = \frac{\sigma^2 + \left(\frac{\sigma^2}{\alpha}\right)^2}{\left(\alpha + \frac{\sigma^2}{\alpha}\right)^2} = \frac{\frac{\sigma^2}{\alpha^2}(\alpha^2 + \sigma^2)}{\frac{1}{\alpha^2}(\alpha^2 + \sigma^2)^2} = \frac{\sigma^2}{\sigma^2 + \alpha^2}$$

b) Note that if  $X(\omega) < \alpha$ , then the left hand side of the inequality is zero while the right hand side is nonnegative, so the inequality holds. If, on the other hand,  $X(\omega) \geq \alpha$ , then  $X(\omega) + u \geq \alpha + u > 0$ , and hence  $\frac{(X+u)^2}{(\alpha+u)^2} \geq 1 \geq \mathbf{1}_{\{X \geq \alpha\}}$ .

c) For any  $u > 0$ , we know from b) that

$$P\{\omega : X(\omega) \geq \alpha\} = E[\mathbf{1}_{\{X \geq \alpha\}}] \leq E\left[\frac{(X+u)^2}{(\alpha+u)^2}\right] = \frac{\sigma^2 + u^2}{(\alpha+u)^2}$$

By a), the smallest possible value of the expression on the right is  $\frac{\sigma^2}{\sigma^2 + \alpha^2}$ , and hence

$$P\{\omega : X(\omega) \geq \alpha\} \leq \frac{\sigma^2}{\sigma^2 + \alpha^2}$$

d) Put  $X = Y - \mu$ ; then  $X$  has expectation 0 and variance  $\sigma^2$ , and hence by c)

$$P\{\omega : X(\omega) \geq \alpha\} \leq \frac{\sigma^2}{\sigma^2 + \alpha^2},$$

which is equivalent to

$$P\{\omega : Y(\omega) \geq \alpha + \mu\} \leq \frac{\sigma^2}{\sigma^2 + \alpha^2}.$$

(This inequality is often called *Cantelli's Inequality*.)

**Problem 4.** a) For  $S_n$  to be 0, we need the  $X_k$ 's to take the values 1 as  $-1$  equally many times, and this is impossible if  $n$  is odd. If  $n$  is even, there are  $\binom{n}{n/2}$  ways in which to choose  $n/2$  positive values among  $n$  possible, and each such combination happens with probability  $p^{n/2}(1-p)^{n/2}$ . Hence

$$P[S_n = 0] = \begin{cases} \binom{n}{n/2} p^{n/2} (1-p)^{n/2} & \text{if } n \text{ is even} \\ 0 & \text{if } n \text{ is odd} \end{cases}$$

b) Let

$$A_n = \{\omega : S_{2n}(\omega) = 0\}$$

(note the shift from  $2n$  to  $n$ ). If we can show that  $\sum_{n=0}^{\infty} P(A_n) < \infty$ , the Borel-Cantelli Lemma tells us that the probability that  $S_n$  is 0 infinitely many times equals zero. As

$$P(A_n) = \binom{2n}{n} p^n (1-p)^n$$

by part a), we can use the Ratio Test to check if  $\sum_{n=0}^{\infty} P(A_n)$  converges:

$$\begin{aligned} \lim_{n \rightarrow \infty} \frac{P(A_{n+1})}{P(A_n)} &= \lim_{n \rightarrow \infty} \frac{\binom{2n+2}{n+1} p^{n+1} (1-p)^{n+1}}{\binom{2n}{n} p^n (1-p)^n} \\ &= \lim_{n \rightarrow \infty} \frac{(2n+2)(2n+1)}{(n+1)^2} p(1-p) = 4p(1-p) \end{aligned}$$

A little calculus shows that  $f(p) = p(1-p)$  has its maximal value  $\frac{1}{4}$  at  $p = \frac{1}{2}$ , and since by assumption  $p \neq \frac{1}{2}$ , we have  $p(1-p) < \frac{1}{4}$ . Thus

$$\lim_{n \rightarrow \infty} \frac{P(A_{n+1})}{P(A_n)} < 1$$

which means that the series  $\sum_{n=0}^{\infty} P(A_n)$  converges, and hence the probability that  $S_n$  is 0 infinitely many times is zero. (If  $p = \frac{1}{2}$ , one can show that with probability 1,  $S_n = 0$  infinitely many times.)

**Comment:** It is also possible to use Stirling's formula  $n! \sim \sqrt{2\pi n} \left(\frac{n}{e}\right)^n$  to solve this problem, but one needs to be a little careful with the formulations as  $n! \sim \sqrt{2\pi n} \left(\frac{n}{e}\right)^n$  means that

$$\lim_{n \rightarrow \infty} \frac{n!}{\sqrt{2\pi n} \left(\frac{n}{e}\right)^n} = 1$$

and not that

$$\lim_{n \rightarrow \infty} \left( n! - \sqrt{2\pi n} \left(\frac{n}{e}\right)^n \right) = 0$$

(this last limit is in fact infinite). As long as one sticks to ratio or comparison tests, it is not hard to get this to work.