

Extension to several dimensions (of the non-parametric techniques seen on Friday)

$$x \in \mathbb{R} \rightarrow x \in \mathbb{R}^p$$

E.g., when $p=2$

$$y = f(x_1, x_2) + \epsilon$$

$$\tilde{f}(x) = \underbrace{f(x)}_{\beta_0} + \underbrace{\beta_1(x-x_0)}_{\beta_1}$$

where $f: \mathbb{R}^2 \rightarrow \mathbb{R}$

Let $y_i \in \mathbb{R}$, $x_i = (x_{i1}, x_{i2}) \in \mathbb{R}^2$, $i=1, \dots, n$

The problem is to find the parameters β which minimize

$$\min_{\beta_0, \beta_1} \sum_{i=1}^n \left[y_i - \beta_0 - \beta_1(x_{i1} - x_{01}) - \beta_2(x_{i2} - x_{02}) \right]^2 w_i$$

where w_i has form

$$w_i = \frac{1}{h_1 h_2} K\left(\frac{x_{i1} - x_{01}}{h_1}\right) K\left(\frac{x_{i2} - x_{02}}{h_2}\right)$$

Note that there are two **smoothing** parameters, h_1, h_2 , to take into account the different variability in the two dimensions

Again, we obtain $\hat{\beta}$ through the weighted least-squares

$$\hat{\beta} = (X^T W X)^{-1} X^T W y$$

where $y = (y_1, \dots, y_n)^T$

$$W = \text{diag}(w_1, \dots, w_n)$$

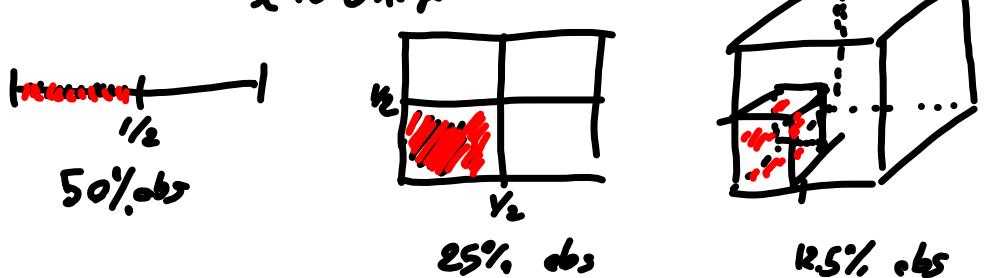
$$X = (1, x_{i1} - x_{01}, x_{i2} - x_{02})$$

While there is no conceptual difference from $p=2$ to a general $p=p$, in practice $p>2$ is almost never used

- difficulties in plotting/visualizing mentally
- hard to interpret the results
- curse of dimensionality

When the number of dimensions increase, the number of observations close to the point of interest is decreases very quickly

E.g.



In order to compensate for the increase dimensionality, we need a number of observations that increases of order n^p (e.g., if we want to use 100 observations in 1 dimension, we need $100^1 = 100$ observations with 5 variables, with 10 variables, we would need $100^5 = 100\,000\,000$ (ten billions) observations)

The problem holds for all non-parametric techniques

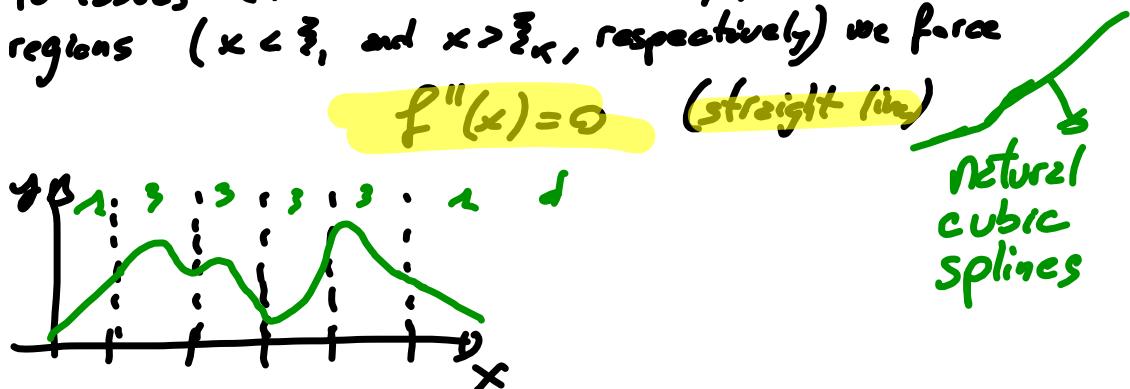
- number of observations
- computational cost

Alternative: use principal components
(information concentrated on a few dimensions)

Splines : piecewise polynomial functions

- split the support of x into several regions (fix K points, $\xi_1 < \xi_2 < \dots < \xi_K$) $\leftarrow \xi_i$ are called knots
- fit a polynomial inside each interval
 - ↳ of the preferred degree d (it is almost always $d=3$)
 - we force the polynomials to have the same value at the knots
 $f(\xi_i^-) = f(\xi_i^+)$ cubic splines
 - same with the first derivative
 $f'(\xi_i^-) = f'(\xi_i^+)$
 - and with the second derivative
 $f''(\xi_i^-) = f''(\xi_i^+)$

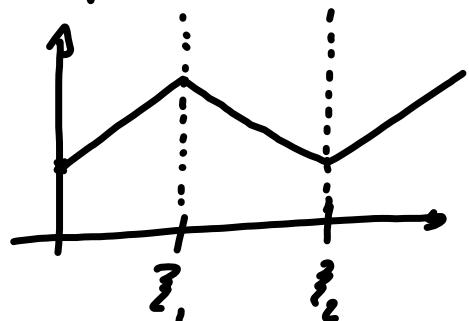
Due to issues related to the variability, in the first and in the last regions ($x < \xi_1$ and $x > \xi_K$, respectively) we force



Use of splines to evaluate the relationship between x and y

$$y = f(x; \beta) + \varepsilon$$

Simplest case: $K=2$, $d=1$

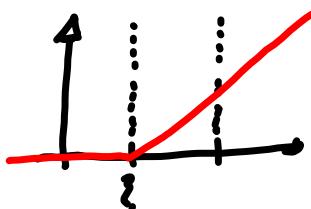


$$f(x; \beta) = \beta_0 + \beta_1 x + \beta_2 (x - \xi_1)_+ + \beta_3 (x - \xi_2)_+$$

$$\underline{h_1(x)} = 1 \quad \underline{h_3(x)} = (x - \xi_1)_+$$

$$\underline{h_2(x)} = x \quad \underline{h_4(x)} = (x - \xi_2)_+$$

where $(x - \xi_i)_+ = \max(0, x - \xi_i)$



$$f(x; \beta) = \beta_0 h_1(x) + \beta_1 h_2(x) + \beta_2 h_3(x) + \beta_3 h_4(x) = \sum_{j=1}^4 \hat{\beta}_j h_j(x)$$

basis

In the case of cubic splines, with a general number of knots k ,

$$f(x) = \sum_{j=1}^{k+1} \hat{\beta}_j h_j(x)$$

where $h_j(x) = x^{j-1}$ for $j = 1, \dots, k$

$$\begin{aligned} h_1(x) &= 1 \\ h_2(x) &= x \\ h_3(x) &= x^2 \\ h_4(x) &= x^3 \end{aligned}$$

$$h_{j+1}(x) = (x - \xi_j)_+^3 \quad \text{for } j = k, \dots, K$$

We need to decide K , the number of knots, and where to place them. K is our complexity parameter: higher values, more complex functions.
↳ find by cross-validation

Once we selected K , we can position the knots

- uniformly along the x : range



- using the quantiles of the empirical distribution of x



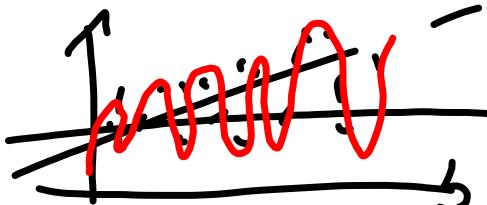
Smoothing splines

Consider the penalized least squares criterion

$$J(f, \lambda) = \sum_{i=1}^n (y_i - f(x_i))^2 + \lambda \int_{-\infty}^{\infty} (f''(t))^2 dt, \quad \lambda > 0$$

λ is the smoothing parameter

The penalty penalizes the "bumpyness" of the curve



$\lambda \rightarrow 0$, more and more

curvature is allowed, until for $\lambda=0 \rightarrow$ interpolation

$\lambda \rightarrow \infty$, curvature is penalized more and more, so for a sufficiently large $\lambda \rightarrow$ straight line

The important result (Green & Silverman, 1994) is that the minimizer of $J(f, \lambda)$ is a natural cubic spline,

which can be rewritten as

$$\hat{f}(x) = \sum_{j=1}^n \hat{\theta}_j N_j(x)$$

$n_0 = \#$ of unique points x_i :

$N_j(x)$ are the basis functions of a natural cubic splines,

$$N_0(x) = 1, \quad N_1(x) = x, \quad N_{K+2}(x) = d_K(x) - d_{K+1}(x)$$

with

$$d_K(x) = \frac{(x - \bar{x}_K)_+^3 - (x - \bar{x}_{K+1})_+^3}{\bar{x}_{K+1} - \bar{x}_K}$$

we will derive this
in STAT-4230

The nice part is that we can rewrite $J(f, \lambda) \approx$

$$J(f, \lambda) = (y - N\theta)^T (y - N\theta) + \lambda \theta^T Q \theta$$

where $\{y\}_{ij} = N_j(x_i)$ and $\{Q\}_{jk} = \int N_j''(t) N_k''(t) dt$

This formula remind us that of the ridge regression, so

$$\hat{\theta} = (N^T N + \lambda Q)^{-1} N y \leftarrow$$

the solution depends on $\lambda \rightarrow Q$ generalized ridge estimator
find by cross-validation