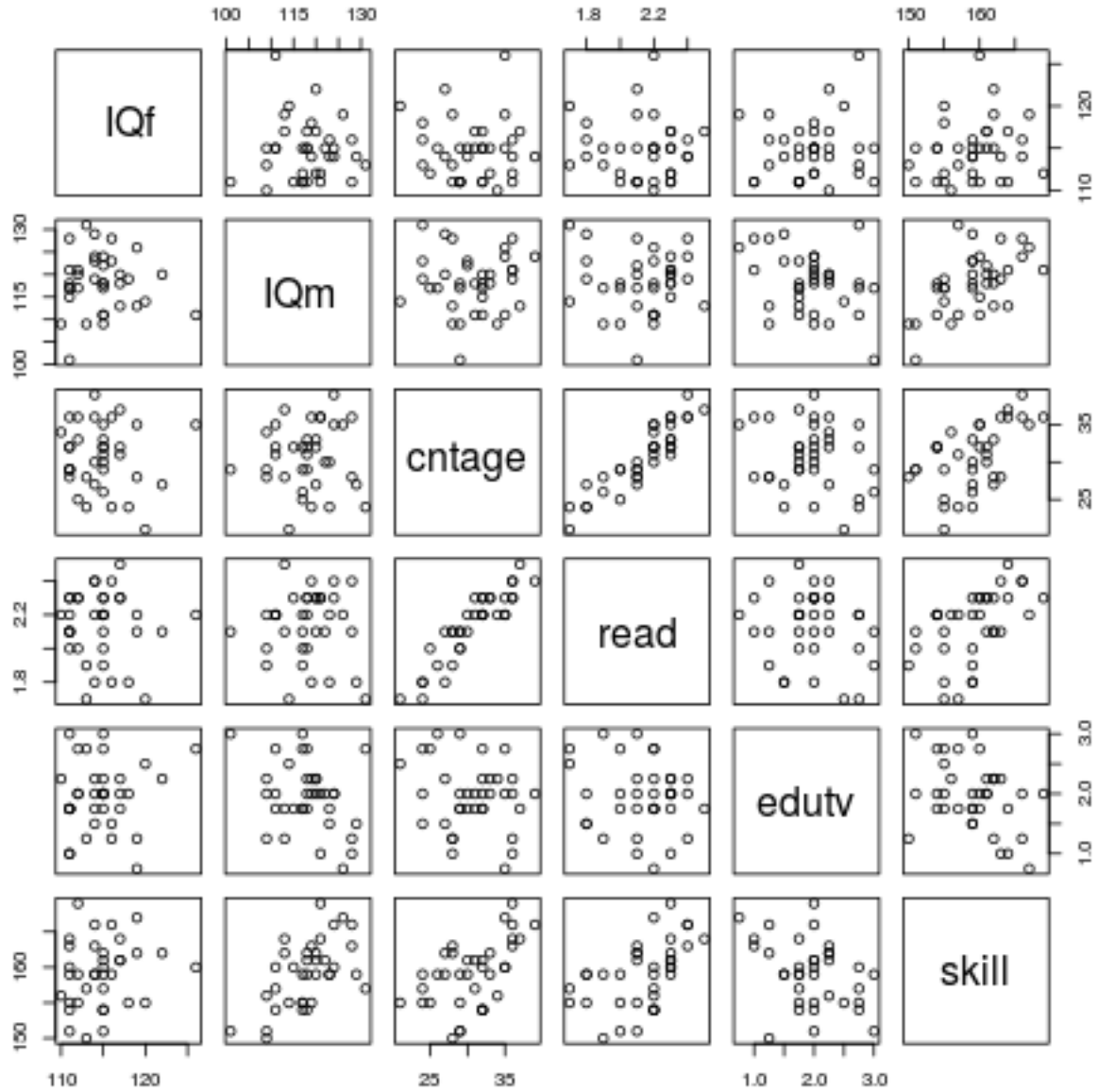


1  
a



```
> summary(skills)
```

IQf		IQm		cntage		read	
Min.	:110.0	Min.	:101.0	Min.	:21.00	Min.	:1.700
1st Qu.:	:112.0	1st Qu.:	:113.8	1st Qu.:	:28.00	1st Qu.:	:2.000
Median	:115.0	Median	:118.0	Median	:31.00	Median	:2.200
Mean	:114.8	Mean	:118.2	Mean	:30.69	Mean	:2.136
3rd Qu.:	:116.2	3rd Qu.:	:122.2	3rd Qu.:	:34.25	3rd Qu.:	:2.300
Max.	:126.0	Max.	:131.0	Max.	:39.00	Max.	:2.500

	edutv	skill
Min.	:0.750	Min. :150.0
1st Qu.:	1.750	1st Qu.:155.0
Median	:2.000	Median :159.0
Mean	:1.958	Mean :159.1
3rd Qu.:	2.250	3rd Qu.:162.0
Max.	:3.000	Max. :169.0

IQm, cntage and read seems highly correlated with skill. And cntage and read is also highly correlated.  
**b**

Forwards selection works by first finding the best model with one variable, then finding the best model found by including 1 more variable etc. Mallows Cp works by finding the number of parameters where Mallows Cp is closest to the number of parameters (m) + 1.

```
library(leaps)
skills = read.table(
  "http://www.uio.no/studier/emner/matnat/math/STK2120/v16/skills.txt",
  header=T)
summary(skills)
png("Oblig2-1summary.png")
plot(skills)
dev.off()
```

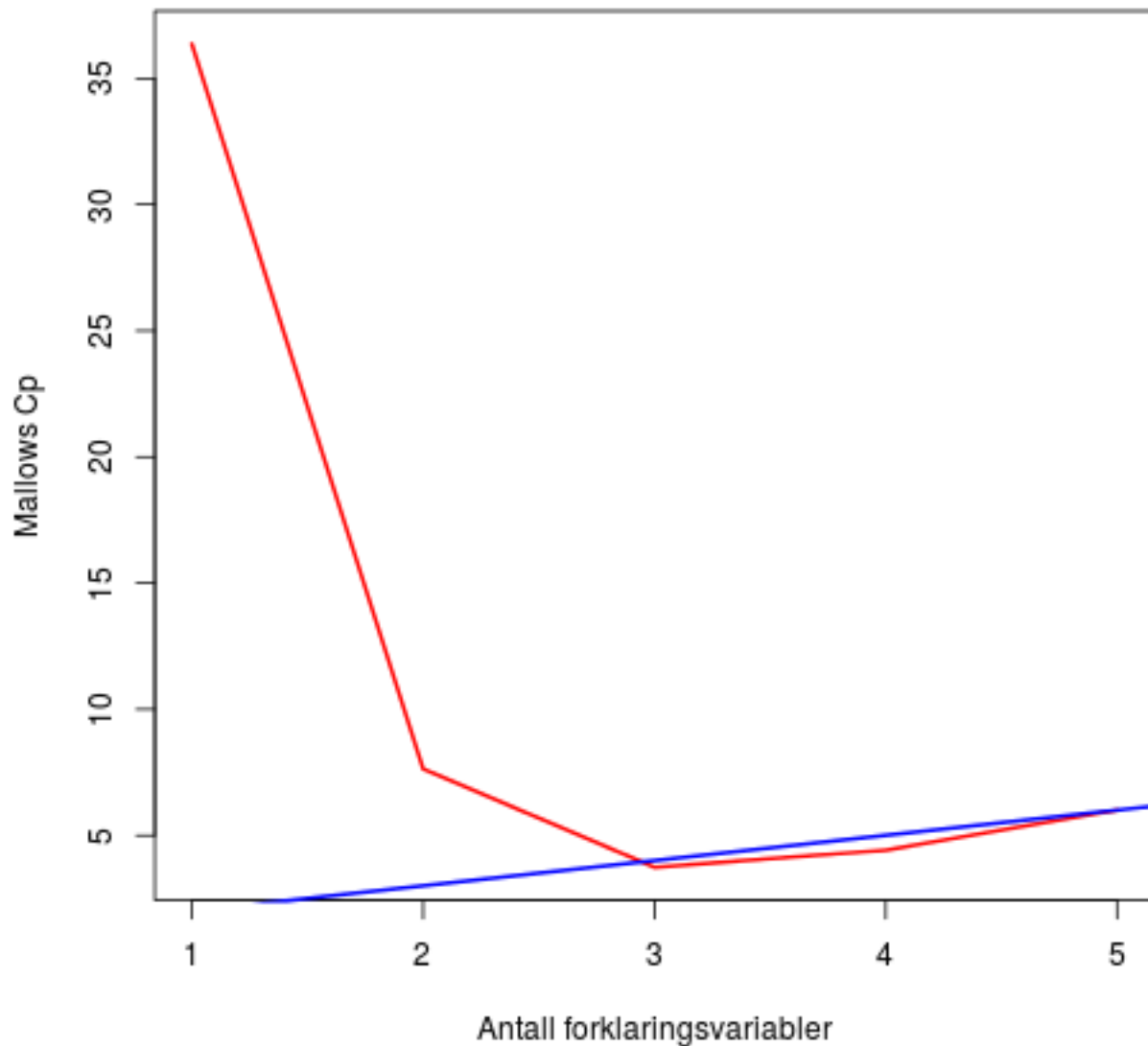
```
fit.forward = regsubsets(skill~., data=skills, nvmax=10, method="forward")
summary.forward = summary(fit.forward)
print(summary.forward)
png("Oblig2-1mallows.png")
```

```
plot(1:5, summary.forward$cp,
     xlab="Antall forklaringsvariabler", ylab="Mallows_Cp",
     type='l', col='red', lwd=2)
abline(1,1, lwd=2, col='blue')
dev.off()
```

```
fit.backward = regsubsets(skill~., data=skills, nvmax=10, method="backward")
summary.backward = summary(fit.backward)
print(summary.backward)
```

```
fit.exhaustive = regsubsets(skill~., data=skills, nvmax=10, method="exhaustive")
summary.exhaustive = summary(fit.exhaustive)
print(summary.exhaustive)
```

```
Selection Algorithm: forward
      IQf IQm cntage read edutv
1 ( 1 ) " " " " " " " "
2 ( 1 ) " " " " " " " "
3 ( 1 ) " " " " " " " "
4 ( 1 ) " " " " " " " "
5 ( 1 ) " " " " " " " "
```



Since  $C_p$  crosses  $m + 1$  at  $m = 3$ , we use a model with 3 parameters. By forward selection this is: IQf, IQm and read. Cntage is highly correlated with skill, but is not included since it is also correlated with read, so including both does not improve the model much.

**c**

Backwards works like forwards, except starting with all variables, and iteratively removing 1 variable. Exhaustive works by trying every combination of variables.

Selection Algorithm: backward

		IQf	IQm	cntage	read	edutv
1	( 1 )	"_"	"*"	"_"	"_"	"_"
2	( 1 )	"_"	"*"	"_"	"*"	"_"
3	( 1 )	"*"	"*"	"_"	"*"	"_"

```

4 ( 1 ) "*" "*" "_" "*" "*"
5 ( 1 ) "*" "*" "*" "*" "*"

```

Selection Algorithm: exhaustive

```

      IQf IQm cntage read edutv
1 ( 1 ) "_" "*" "_" "_" "_"
2 ( 1 ) "_" "*" "_" "*" "*"
3 ( 1 ) "*" "*" "_" "*" "*"
4 ( 1 ) "*" "*" "_" "*" "*"
5 ( 1 ) "*" "*" "*" "*" "*"

```

All methods yields the same models.

## 2

a

$$\begin{aligned}
L(\beta_0, \beta_1) &= \prod_i [p(x_i)^{y_i} (1 - p(x_i))^{1-y_i}] \\
L(\beta_0, \beta_1) &= \prod_i \left[ \left( \frac{e^{\beta_0 + \beta_1 x_i}}{1 + e^{\beta_0 + \beta_1 x_i}} \right)^{y_i} \left( 1 - \frac{e^{\beta_0 + \beta_1 x_i}}{1 + e^{\beta_0 + \beta_1 x_i}} \right)^{1-y_i} \right] \\
L(\beta_0, \beta_1) &= \prod_i \left[ \left( \frac{e^{\beta_0 + \beta_1 x_i}}{1 + e^{\beta_0 + \beta_1 x_i}} \right)^{y_i} \left( \frac{1}{1 + e^{\beta_0 + \beta_1 x_i}} \right)^{1-y_i} \right] \\
L(\beta_0, \beta_1) &= \prod_i \left[ \left( \frac{e^{\beta_0 + \beta_1 x_i}}{1 + e^{\beta_0 + \beta_1 x_i}} \right)^{y_i} \frac{(1 + e^{\beta_0 + \beta_1 x_i})^{y_i}}{1 + e^{\beta_0 + \beta_1 x_i}} \right] \\
L(\beta_0, \beta_1) &= \prod_i \left[ \frac{e^{\beta_0 y_i + \beta_1 x_i y_i}}{1 + e^{\beta_0 + \beta_1 x_i}} \right]
\end{aligned}$$

b

$$\begin{aligned}
s_1(\beta_0, \beta_1) &= \frac{\partial l(\beta_0, \beta_1)}{\partial \beta_0} \\
l(\beta_0, \beta_1) &= \sum \log \frac{e^{\beta_0 y_i + \beta_1 x_i y_i}}{1 + e^{\beta_0 + \beta_1 x_i}} \\
l(\beta_0, \beta_1) &= \sum [\beta_0 y_i + \beta_1 x_i y_i - \log(1 + e^{\beta_0 + \beta_1 x_i})] \\
s_1(\beta_0, \beta_1) &= \frac{\partial}{\partial \beta_0} \sum [\beta_0 y_i + \beta_1 x_i y_i - \log(1 + e^{\beta_0 + \beta_1 x_i})] \\
s_1(\beta_0, \beta_1) &= \sum \left[ y_i - \frac{e^{\beta_0 + \beta_1 x_i}}{1 + e^{\beta_0 + \beta_1 x_i}} \right] \\
s_2(\beta_0, \beta_1) &= \frac{\partial}{\partial \beta_1} \sum [\beta_0 y_i + \beta_1 x_i y_i - \log(1 + e^{\beta_0 + \beta_1 x_i})] \\
s_2(\beta_0, \beta_1) &= \sum \left[ x_i y_i - \frac{e^{\beta_0 + \beta_1 x_i} x_i}{1 + e^{\beta_0 + \beta_1 x_i}} \right]
\end{aligned}$$

c

$$\begin{aligned}
J_{11}(\beta_0, \beta_1) &= -\frac{\partial s_1(\beta_0, \beta_1)}{\partial \beta_0} \\
J_{11}(\beta_0, \beta_1) &= -\frac{\partial}{\partial \beta_0} \sum \left[ y_i - \frac{e^{\beta_0 + \beta_1 x_i}}{1 + e^{\beta_0 + \beta_1 x_i}} \right] \\
J_{11}(\beta_0, \beta_1) &= -\sum \left[ -\frac{e^{\beta_0 + \beta_1 x_i} (1 + e^{\beta_0 + \beta_1 x_i}) - e^{\beta_0 + \beta_1 x_i} e^{\beta_0 + \beta_1 x_i}}{(1 + e^{\beta_0 + \beta_1 x_i})^2} \right] \\
&= \sum \left[ \frac{e^{\beta_0 + \beta_1 x_i} (1 + e^{\beta_0 + \beta_1 x_i} - e^{\beta_0 + \beta_1 x_i})}{(1 + e^{\beta_0 + \beta_1 x_i})^2} \right] \\
&= \sum \frac{e^{\beta_0 + \beta_1 x_i}}{(1 + e^{\beta_0 + \beta_1 x_i})^2}
\end{aligned}$$

$$\begin{aligned}
J_{22}(\beta_0, \beta_1) &= -\frac{\partial s_2(\beta_0, \beta_1)}{\partial \beta_1} \\
&= -\frac{\partial}{\partial \beta_1} \sum \left[ x_i y_i - \frac{e^{\beta_0 + \beta_1 x_i} x_i}{1 + e^{\beta_0 + \beta_1 x_i}} \right] \\
&= \sum \left[ \frac{e^{\beta_0 + \beta_1 x_i} x_i^2 (1 + e^{\beta_0 + \beta_1 x_i}) - e^{\beta_0 + \beta_1 x_i} x_i e^{\beta_0 + \beta_1 x_i} x_i}{(1 + e^{\beta_0 + \beta_1 x_i})^2} \right] \\
&= \sum \left[ \frac{e^{\beta_0 + \beta_1 x_i} x_i^2 (1 + e^{\beta_0 + \beta_1 x_i} - e^{\beta_0 + \beta_1 x_i})}{(1 + e^{\beta_0 + \beta_1 x_i})^2} \right] \\
&= \sum \frac{e^{\beta_0 + \beta_1 x_i} x_i^2}{(1 + e^{\beta_0 + \beta_1 x_i})^2}
\end{aligned}$$

$$\begin{aligned}
J_{12}(\beta_0, \beta_1) = J_{21}(\beta_0, \beta_1) &= -\frac{\partial s_2(\beta_0, \beta_1)}{\partial \beta_0} \\
&= -\frac{\partial}{\partial \beta_0} \sum \left[ x_i y_i - \frac{e^{\beta_0 + \beta_1 x_i} x_i}{1 + e^{\beta_0 + \beta_1 x_i}} \right] \\
&= \sum \left[ \frac{e^{\beta_0 + \beta_1 x_i} x_i (1 + e^{\beta_0 + \beta_1 x_i}) - e^{\beta_0 + \beta_1 x_i} x_i e^{\beta_0 + \beta_1 x_i}}{1 + e^{\beta_0 + \beta_1 x_i}} \right] \\
&= \sum \left[ \frac{e^{\beta_0 + \beta_1 x_i} x_i (1 + e^{\beta_0 + \beta_1 x_i} - e^{\beta_0 + \beta_1 x_i})}{1 + e^{\beta_0 + \beta_1 x_i}} \right] \\
&= \sum \frac{e^{\beta_0 + \beta_1 x_i} x_i}{1 + e^{\beta_0 + \beta_1 x_i}}
\end{aligned}$$

d

$$\begin{aligned}
E(\hat{\beta}_0 + \hat{\beta}_1 x) &= \beta_0 + \beta_1 x \\
V(\hat{\beta}_0 + \hat{\beta}_1 x) &= I^{11}(\beta_0, \beta_1) + x^2 I^2(\beta_0, \beta_1) + 2x I^{12}(\beta_0, \beta_1) \\
&= \sum_i \left( \frac{(1 + e^{\beta_0 + \beta_1 x_i})^2}{e^{\beta_0 + \beta_1 x_i}} + x^2 \sum_i \frac{(1 + e^{\beta_0 + \beta_1 x_i})^2}{x_i^2 e^{\beta_0 + \beta_1 x_i}} + 2x \sum_i \frac{(1 + e^{\beta_0 + \beta_1 x_i})^2}{x_i e^{\beta_0 + \beta_1 x_i}} \right) \\
&= \sum_i \left[ \frac{(1 + e^{\beta_0 + \beta_1 x_i})^2}{e^{\beta_0 + \beta_1 x_i}} + x^2 \frac{(1 + e^{\beta_0 + \beta_1 x_i})^2}{x_i^2 e^{\beta_0 + \beta_1 x_i}} + 2x \frac{(1 + e^{\beta_0 + \beta_1 x_i})^2}{x_i e^{\beta_0 + \beta_1 x_i}} \right] \\
&= \sum_i \left[ (x^2 + 2x_i x + x_i^2) \frac{(1 + e^{\beta_0 + \beta_1 x_i})^2}{x_i^2 e^{\beta_0 + \beta_1 x_i}} \right] \\
&= \sum_i \left[ (x + x_i)^2 \frac{(1 + e^{\beta_0 + \beta_1 x_i})^2}{x_i^2 e^{\beta_0 + \beta_1 x_i}} \right]
\end{aligned}$$

CI for  $\beta_0 + \beta_1 x$ :

$$\hat{\beta}_0 + \hat{\beta}_1 x \pm z_{.025} \sqrt{\sum_i \left[ (x + x_i)^2 \frac{(1 + e^{\beta_0 + \beta_1 x_i})^2}{x_i^2 e^{\beta_0 + \beta_1 x_i}} \right]}$$

Since the logit function is monotone, the CI for  $\text{logit}(\beta_0 + \beta_1 x)$  is just the logit of the CI for  $\beta_0 + \beta_1 x$ , so:

$$\left( \frac{e^{\hat{\beta}_0 + \hat{\beta}_1 x - z_{.025} \sqrt{\sum_i \left[ (x + x_i)^2 \frac{(1 + e^{\beta_0 + \beta_1 x_i})^2}{x_i^2 e^{\beta_0 + \beta_1 x_i}} \right]}}}{1 + e^{\hat{\beta}_0 + \hat{\beta}_1 x - z_{.025} \sqrt{\sum_i \left[ (x + x_i)^2 \frac{(1 + e^{\beta_0 + \beta_1 x_i})^2}{x_i^2 e^{\beta_0 + \beta_1 x_i}} \right]}}}, \frac{e^{\hat{\beta}_0 + \hat{\beta}_1 x + z_{.025} \sqrt{\sum_i \left[ (x + x_i)^2 \frac{(1 + e^{\beta_0 + \beta_1 x_i})^2}{x_i^2 e^{\beta_0 + \beta_1 x_i}} \right]}}}{1 + e^{\hat{\beta}_0 + \hat{\beta}_1 x + z_{.025} \sqrt{\sum_i \left[ (x + x_i)^2 \frac{(1 + e^{\beta_0 + \beta_1 x_i})^2}{x_i^2 e^{\beta_0 + \beta_1 x_i}} \right]}}} \right)$$

3

a

```

s1 <- function(sy, ebetas)
{
  return(sy - sum(ebetas/(1+ebetas)))
}

s2 <- function(sxy, ebetas, x)
{
  return(sxy - sum(x*ebetas/(1+ebetas)))
}

J11 <- function(ebetas)
{
  return(sum(ebetas/(1+ebetas)^2))
}

```

```

J12 <- function(x, ebetas)
{
  return(sum(x*ebetas/(1+ebetas)^2))
}

J22 <- function(x, ebetas)
{
  return(sum(x^2*ebetas/(1+ebetas)^2))
}

scorefunction <- function(x, sy, sxy, ebetas)
{
  return(c(s1(sy, ebetas), s2(sxy, ebetas, x)))
}

obsinfo <- function(ebetas, x)
{
  j12 = J12(x, ebetas)
  return (matrix(c(
    J11(ebetas), j12,
    j12, J22(x, ebetas)
  ), ncol=2))
}

logit <- function(x)
{
  return (exp(x)/(1+exp(x)))
}

logisticNewR <- function(x, y, theta0, eps=0.000001)
{
  n = length(x);
  sy = sum(y)
  sxy = sum(x*y)

  diff = 1; h = 0.0000001;
  theta = theta0
  while(diff>eps)
  {
    theta.old = theta
    ebetas = exp(theta[1] + theta[2]*x)
    s = scorefunction(x, sy, sxy, ebetas)
    Jbar = obsinfo(ebetas, x)
    theta = theta + solve(Jbar, s)
    diff = sum(abs(theta-theta.old))
  }
  return(list(theta=theta, J=Jbar))
}

exmp12.14=read.table(
  "http://www.uio.no/studier/emner/matnat/math/STK2120/v16/exmp12-14.txt",
  header=T)

```

```

x = exmp12.14$temp-70
y = exmp12.14$fail

res = logisticNewR(x, y, c(0,0))

beta0 = res$theta[1]
beta1 = res$theta[2]
J = res$J

Jinv = solve(J)
print(J)
print(Jinv)
SE0 = sqrt(Jinv[1,1])
SE1 = sqrt(Jinv[2,2])

pred = beta0 + (31-70)*beta1

print(beta0)
print(beta1)
print(varBeta0)
print(varBeta1)
b

exmp12.14=read.table(
+   "http://www.uio.no/studier/emner/matnat/math/STK2120/v16/exmp12-14.txt",
+   header=T)
>
> x = exmp12.14$temp-70
> y = exmp12.14$fail
>
> res = logisticNewR(x, y, c(0,0))
>
> beta0 = res$theta[1]
> beta1 = res$theta[2]
> Jinv = solve(res$J)
> SE0 = sqrt(Jinv[1,1])
> SE1 = sqrt(Jinv[2,2])
> beta0
[1] -1.20849
> beta1
[1] -0.2321627
> SE0
[1] 0.5952546
> SE1
[1] 0.1082365
c

> x = 31-70
> pred = beta0 + x*beta1
> phat = logit(pred)
> phat
[1] 0.9996088

```



```
> s = sqrt(Jinv[1,1]+t^2*Jinv[2,2] + 2*t* Jinv[1,2])
> z = qnorm(0.975)
> L = logit(pred-s*z)
> L
[1] 0.4816089
> U = logit(pred+s*z)
> U
[1] 0.9999999
```

So the 95% CI for  $p(x)$  is given by (0.4816, 0.9999). The probability of error seems to be at least almost half, which is high. But we do not have any data points around  $t = 31$  so our model might not be valid there.