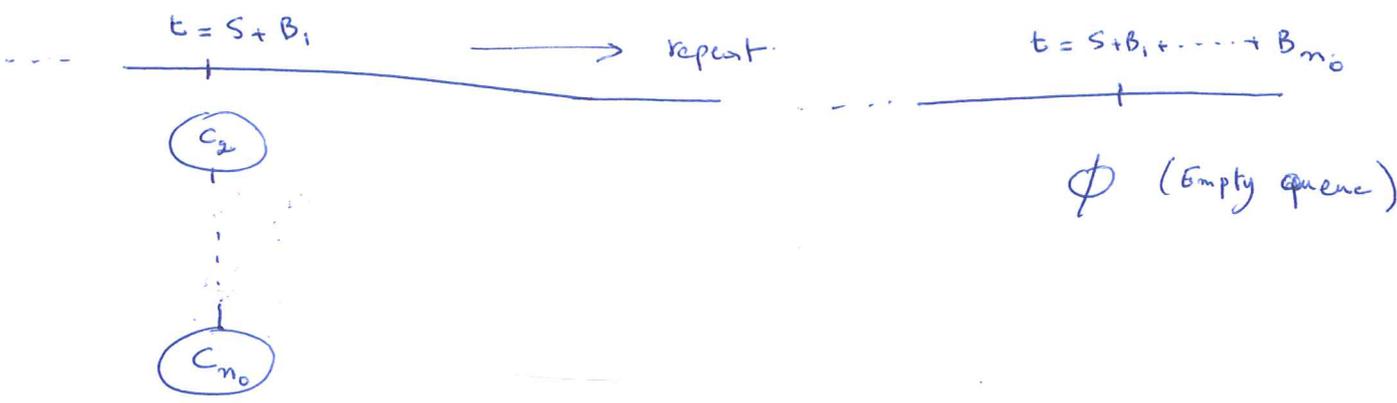
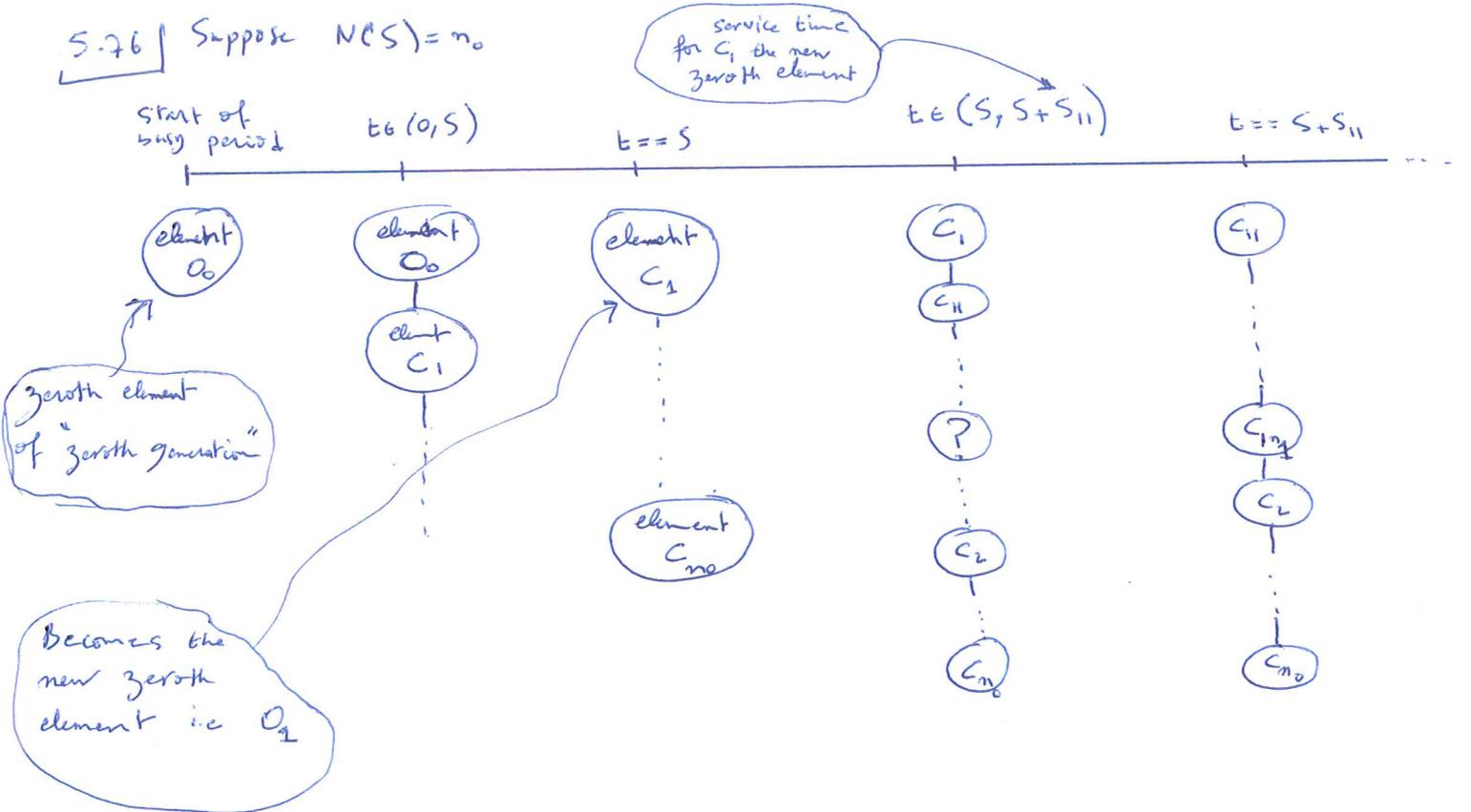


5.76 | Suppose $N(S) = n_0$



Now applying the same logic, if we denote the number of customers during a busy period by M , we have:

$$M = 1 + M_1 + \dots + M_{N(S)} = 1 + \sum_{i=1}^{N(S)} M_i$$

with $M \sim M_i \quad \forall i$

Thus:

$$E[M] = E[E[M|S]] = E[1 + E[\sum_{i=1}^{N(S)} M_i | S]]$$

(eq. 5.24 Ross)

$$= 1 + E[\lambda S E[M]] = 1 + \lambda E[S] E[M]$$

$E[M_i] = E[M]$
 \Rightarrow

$$E[M] = \frac{1}{1 - \lambda E[S]} \quad (\text{if } 1 - \lambda E[S] > 0)$$

Similarly and by the variance formula for compound Poisson processes:

$$Var(M) = E[Var(M|S)] + Var(E[M|S]) = E[S \lambda E[M^2]] + Var(\lambda S E[M])$$

(eq. 5.25 Ross)

$$\begin{aligned}\Rightarrow \text{Var}(M) &= \lambda E[M^2] E[S] + \lambda^2 E[M]^2 \text{Var}(S) \\ &= \lambda E[S] (\text{Var}(M) + E[M]^2) + \lambda^2 E[M]^2 \text{Var}(S)\end{aligned}$$

$$\Rightarrow \text{Var}(M) = \frac{\lambda E[S] E[M]^2 + \lambda^2 \text{Var}(S) E[M]^2}{1 - \lambda E[S]}.$$

Ex 5.78

$$\lambda(t) = \begin{cases} 0 & t \in [0, 8) \\ 4 & t \in [8, 10) \\ 8 & t \in [10, 12) \\ 8 + (t-12) & t \in [12, 14) \\ 10 - 2(t-14) & t \in [14, 17) \\ 0 & t \in [17, 24) \end{cases}$$

$$\int_0^{24} \lambda(t) dt = \left(\int_0^8 + \int_8^{10} + \int_{10}^{12} + \int_{12}^{14} + \int_{14}^{17} + \int_{17}^{24} \right) \lambda(t) dt$$
$$= \dots = 63$$

Thus $N(24) - N(0) \sim \text{Poisson}(63)$.

Ex 6.1

Suppose $(N_1(t), N_2(t)) = (n_1, n_2)$ then:

$$v_{(n_1, n_2)} = \frac{1}{E[T]} \quad \text{where } T \sim \exp(n_1 n_2 \cdot \lambda)$$

This because by definition T is the amount of time the process spends in state (n_1, n_2) , which is exponential with intensity proportional to the intensity of Δ mating multiplied by n_1 and then by n_2 . Moreover

$$P_{(n_1, n_2), (n_1, n_2+1)} = P_{(n_1, n_2), (n_1+1, n_2)} = 1/2$$

Ex 6.2

Let $N_A(t), N_B(t)$ denote the number of organisms in state A, B respectively.

Then if $(N_A(t), N_B(t)) = (n_A, n_B)$, we have:

$$v_{(n_A, n_B)} = [\text{Rate of } T, \text{ the time til the next event}]$$
$$= \alpha n_A + \beta n_B$$

On the other hand:

$$P_{(n_A, n_B) \rightarrow (n_A-1, n_B+1)} = P(\text{A division of type A happens before one of type B})$$
$$= P(T_A < T_B) \quad \text{where } \begin{cases} T_A \sim \exp(\alpha n_A) \\ T_B \sim \exp(\beta n_B) \end{cases}$$
$$= \frac{\alpha n_A}{\alpha n_A + \beta n_B}$$

Similarly:

$$P_{(n_A, n_B) \rightarrow (n_A+2, n_B-1)} = \frac{\beta n_B}{\alpha n_A + \beta n_B} = 1 - P_{(n_A, n_B) \rightarrow (n_A-1, n_B+1)}$$

Ex 6.3 | Note that the failure rate depends on the machine and therefore a Birth-Death model isn't appropriate.

Still we can model the situation using a continuous MC with the following state space:

"~~*~~" := both are working

"~~t₂~~" := both are down but 2 is been serviced.

t₁ := " " " 1 is been serviced.

"*₂" := 1 is working, 2 is down & been serviced.

"*₁" := 2 _____, 1 _____

Now we have:

$$\vartheta_* = \mu_1 + \mu_2 \quad ; \quad \vartheta_{t_2} = \vartheta_{t_1} = \mu \quad ; \quad \vartheta_{*2} = \mu_1 + \mu \quad ; \quad \vartheta_{*1} = \mu_2 + \mu$$

Also:

$$P_{*,*2} = P(1 \text{ fails before } 2) = \frac{\mu_1}{\mu_1 + \mu_2} = 1 - P_{*,*1}$$

$$P_{*2,t_1} = \frac{\mu_2}{\mu_2 + \mu} = 1 - P_{*2,t_2}$$

$$P_{*1,t_2} = \frac{\mu_1}{\mu_1 + \mu} = 1 - P_{*1,t_1}$$

$$P_{t_1,*2} = P_{t_2,*1} = 1 \quad \left(\begin{array}{l} \text{As } t_1 \rightsquigarrow *2 \text{ is the only transition possible} \\ \text{from } t_1 \text{. Similar for } t_2 \end{array} \right)$$

Ex 6.4 | If we think of the people inside the queue as a population, then this population grows with rate $\lambda_m = \alpha_m \cdot \lambda$ and diminishes with rate $\mu_m = \mu$.