

$\pi = (\pi_0, \dots, \pi_k)$ long run proportions
 $\pi < \pi P, \sum \pi_j = 1$

Example Hardy-Weinberg law and Markov chains
 Consider a large population
 Pool of genes of types a and A
 Each individual has two genes.
 Consider a first generation

Type of combinations	proportions	
AA	r_0	$r_0 + q_0 + r_0 = 1$
aA	q_0	
Aa	r_0	

Random mating: Individuals are selected randomly and contribute one gene each to new individual and this gene picked randomly

In the next generation the proportion of genes are
 $P(A) = P(A|AA)P(AA) + P(A|aA)P(aA) + P(A|Aa)P(Aa) + P(A|aa)P(aa)$
 $= \frac{1}{2} r_0 + 0 \cdot q_0 + \frac{1}{2} r_0 = r_0 + r_0/2$
 $P(a) = q_0 + r_0/2$

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The proportion of gene combinations are now
 $p = P(AA) = P(A)P(A) = (r_0 + \frac{r_0}{2})^2$
 $q = P(aa) = P(a)P(a) = (q_0 + \frac{r_0}{2})^2$
 $r = P(Aa) \cdot P(aA) = 2P(A)P(a) = 2(r_0 + \frac{r_0}{2})(q_0 + \frac{r_0}{2})$

because mating is random
 In a third generation the proportion of A and a is $p + \frac{r}{2}$ and $q + \frac{r}{2}$
 But $p + \frac{r}{2} = (r_0 + \frac{r_0}{2})^2 + (r_0 + \frac{r_0}{2})(q_0 + \frac{r_0}{2})$
 $= (r_0 + \frac{r_0}{2}) [r_0 + \frac{r_0}{2} + q_0 + \frac{r_0}{2}] = (r_0 + \frac{r_0}{2})$
 So in the proportion of genes of type A and a are the same as in the previous generation
 Hence with random mating the proportion of individuals with combinations AA, aA, Aa will be fixed as p, q and r .

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This is the famous Hardy-Weinberg law.
 What about Markov Chains?
 Consider a situation where individuals have 1 offspring and let X_n be the state at n -th generation. X_n is a Markov chain with transition matrix

	AA	aA	AA
$P =$	$\begin{pmatrix} p + \frac{r}{2} & 0 & q + \frac{r}{2} \\ 0 & q + \frac{r}{2} & p + \frac{r}{2} \\ \frac{1}{2}(p + \frac{r}{2}) & \frac{1}{2}(q + \frac{r}{2}) & \frac{p}{2} + \frac{q}{2} + \frac{r}{2} \end{pmatrix}$		

Is $\pi = (p, q, r)$ the long-run proportions?
 $(p, q, r) \begin{bmatrix} p_{00} \\ p_{10} \\ p_{20} \end{bmatrix} = p(p + \frac{r}{2}) + r(\frac{p}{2} + \frac{q}{2}) = p + \frac{r}{2} p$
 which we have shown is equal to $P(AA) = p$.

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Similarly $(p, q, r) \begin{bmatrix} p_{01} \\ p_{11} \\ p_{21} \end{bmatrix} = q(q + \frac{r}{2}) + r(\frac{q}{2} + \frac{r}{2}) = q + \frac{r}{2} q$
 and $(p, q, r) \begin{bmatrix} p_{02} \\ p_{12} \\ p_{22} \end{bmatrix} = p(q + \frac{r}{2}) + q(p + \frac{r}{2}) + r(p + q + r) = 2pq + \frac{r}{2}p + \frac{r}{2}q + \frac{r}{2}p + \frac{r}{2}q + \frac{r}{2}r = 2pq + r + q + r = 2pq + r + 2(\frac{p}{2} + \frac{q}{2}) = 2pq + r + p + q = p + q + r = 1$
 $P(Aa) = r$
 Hence we have shown $\pi = \pi P$
 where $\pi = (p, q, r)$ is Hardy-Weinberg law.
 The long run proportions $\pi_j, j > 0$ is often called the stationary probabilities and satisfy the following. If $\pi_j = P(X_0 = j)$ then $\pi_j = P(X_n = j)$ for all $n \geq 0$ and $j \geq 0$

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To show the result one can use induction $\pi = \pi P$
 OK for $n=0$, so suppose OK for $n-1$, i.e. $\pi_j = P(X_{n-1} = j) \forall j$. Then
 $P(X_n = j) = \sum_i P(X_n = j, X_{n-1} = i) = \sum_i P(X_{n-1} = i) P(X_n = j | X_{n-1} = i) = \sum_i \pi_i P_{ij} = \pi_j$

Proposition 4.6 Let $\{X_n, n \geq 0\}$ be an irreducible Markov chain with stationary probabilities $\pi_j, j \geq 0$ and let r be a bounded function on the state space, then
 $\lim_{N \rightarrow \infty} \frac{\sum_{n=1}^N r(X_n)}{N} = \sum_{j=0}^{\infty} r(j) \pi_j$

Proof Let $a_j(N)$ be the amount of time X_n is in j for $n \leq N$
 Then $\sum_{n=1}^N r(X_n) = \sum_{j=0}^{\infty} a_j(N) r(j)$. But $\frac{a_j(N)}{N} \rightarrow \pi_j$
 Hence $\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N r(X_n) = \lim_{N \rightarrow \infty} \sum_{j=0}^{\infty} \frac{a_j(N)}{N} r(j) \rightarrow \sum_{j=0}^{\infty} \pi_j r(j)$

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Example Bonus/malus, auto insurance

		1	2	3	4	states {1,2,3,4}
$P =$	1	a_0	a_1	a_2	a_3	premiums
	2	a_0	0	a_1	$1 - a_0 - a_1$	200, 250, 400, 600
	3	0	a_0	0	$1 - a_0$	
	4	0	0	a_0	$1 - a_0$	

 $R = 1, 2, 3$ stationary probability solves
 $\pi_1 = a_0 \pi_1 + a_0 \pi_2$
 $\pi_2 = a_1 \pi_1 + a_0 \pi_2$
 $\pi_3 = a_2 \pi_1 + a_1 \pi_2 + a_0 \pi_3$
 $\pi_1 + \pi_2 + \pi_3 + \pi_4 = 1$

Average annual premium
 $AAP = 200 \cdot \pi_1 + 250 \cdot \pi_2 + 400 \cdot \pi_3 + 600 \cdot \pi_4$
 when $L = 1/2, a_0 = 0.60, a_1 = 0.20, a_2 = 0.075, a_3 = 0.125$
 $\pi_1 = 0.3692, \pi_2 = 0.2875, \pi_3 = 0.2103, \pi_4 = 0.1330$
 Then AAP = 326.375

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4.4.1 Limiting probabilities

If $P = \begin{pmatrix} 0.7 & 0.3 \\ 0.4 & 0.6 \end{pmatrix}$, $P^2 = \begin{pmatrix} 0.574 & 0.426 \\ 0.568 & 0.432 \end{pmatrix}$

$P = \pi$

$$P^2 = \begin{pmatrix} 0.572 & 0.428 \\ 0.570 & 0.430 \end{pmatrix}; \pi_1 = \frac{4}{7} = 0.571$$

$$\pi_2 = \frac{3}{7} = 0.429$$

so will $P^n \rightarrow$ matrix with equal rows
and with stationary probabilities
rows given by stationary probabilities

That is not always true.

If $P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$, X_n alternates between 0 and 1
and the long run proportions are $\pi_0 = 1/2, \pi_1 = 1/2$

Hence $P_{00}^n = \begin{cases} 1 & n \text{ even} \\ 0 & n \text{ odd} \end{cases}$

Definition A chain that can only return to a state
in multiple of $d > 1$ is called periodic
and does not have limiting probabilities
If the chain is not periodic it is aperiodic

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For aperiodic chains the limiting probabilities
when they exist are equal to the long run
proportions and are independent of initial states.

Let $d_j = \lim_{n \rightarrow \infty} P(X_n = j)$

But $P(X_{n+1} = j) = \sum_{i=0}^2 P(X_{n+1} = j | X_n = i) \cdot P(X_n = i)$

$$= \sum_{i=0}^2 P_{ij} \cdot P(X_n = i)$$

and also $d_j = \sum_{i=0}^2 P(X_{n+1} = i)$

Let $n \rightarrow \infty$, then $d_j = \sum_{i=0}^2 d_i P_{ij}$
so $\{d_j\}$ solves the equations for $\{\pi_j\}$.
where they are unique solution for $\pi = \pi P$

Hence $d_j = \pi_j$. A sufficient condition
for uniqueness is that the chain is
irreducible and positive recurrent

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