

1. Branching processes.  
Models for dynamics of population

$X_0 = 1$

$X_n$ : size of population time  $n$

$Z_i$ : offspring distribution

$\mu = \mathbb{E}(Z)$ ,  $\sigma^2 = \text{Var}(Z)$

$P(Z=j) = P_j$

Have shown: If  $P_0 > 0$  either  $X_n \rightarrow 0$  or  $X_n \rightarrow \infty$

Let  $\pi_0 = \lim_{n \rightarrow \infty} P(X_n = 0 | X_0 = 1)$  i.e. the probability that the population dies out

$\pi_0$  ?

$0 < \mu < 1$ :  $\mu^n = E(X_n) = \sum_{j=0}^n j P(X_n = j | X_0 = 1)$   
 $\geq \sum_{j=1}^n P(X_n = j | X_0 = 1)$   
 $P(X_n \geq 1 | X_0 = 1)$

Since  $\mu^n \rightarrow 0$ , then  
 $\lim_{n \rightarrow \infty} P(X_n = 0 | X_0 = 1) = 1 - \lim_{n \rightarrow \infty} P(X_n \geq 1 | X_0 = 1) = 1$

and  $\lim_{n \rightarrow \infty} P(X_n = 0 | X_0 = 1) \rightarrow \pi_0$

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$\mu > 1$ .  $\pi_0 = P(\text{population dies out} | X_0 = 1)$

$$= \sum_{j=0}^{\infty} P(\text{population dies out} | X_1 = j) \cdot P(X_1 = j | X_0 = 1)$$

$$= \sum_{j=0}^{\infty} \pi_0^j P_j$$

Then  $\pi_0$  is the smallest value that satisfies  $\pi_0 = \sum_{j=0}^{\infty} \pi_0^j P_j$

$\mu = 1$ : Then one can show  $\lim_{n \rightarrow \infty} P(X_n = 0 | X_0 = 1) \rightarrow \pi_0$

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Examples:  $P_0 = 1/2, P_1 = 1/4, P_2 = 1/4$

Then  $\mu = 0 \cdot 1/2 + 1 \cdot 1/4 + 2 \cdot 1/4 = 3/4 < 1$

Example  $P_0 = 1/4, P_1 = 1/4, P_2 = 1/2$

$\mu = 1 \cdot 1/4 + 2 \cdot 1/2 = 5/4 > 1$

$\pi_0$  satisfies  
 $\pi_0 = \frac{1}{4} + \frac{1}{4} \pi_0 + \frac{1}{2} \pi_0^2$

so  $2\pi_0^2 - 3\pi_0 + 1 = 0$  with solution  
 $\frac{3 \pm \sqrt{9-8}}{2 \cdot 2} = \left\{ \frac{1}{2}, \frac{1}{2} \right\}$

so  $\pi_0 = 1/2$

Example if  $X_0 = n$  then  
 $\lim_{n \rightarrow \infty} P(\text{population dies out}) = \pi_0^n$

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4.8 Time-reversible Markov chains

Suppose that  $X_n$  is ergodic (irreducible, aperiodic, positive recurrent)

Let  $P$  be transition matrix of  $X_n$

We know that  $\exists \pi = \pi P$   $\pi_j > 0 \forall j$

Let  $Y_n = X_{n-N}$ ,  $n = 0, 1, \dots, N$  and consider  $Y_0, Y_1, \dots, Y_N$  which is  $X_n$  run backward.

Is  $Y_n$  a Markov chain?

If yes, what is the transition matrix?

$$P(Y_n = j | Y_{n-1} = i) = P(X_{N-n} = j | X_{N-n-1} = i)$$

$$= \frac{P(X_{N-n} = i | X_{N-n-1} = j) P(X_{N-n-1} = j)}{P(X_{N-n} = i)}$$

$$= P_{ji} \frac{\pi_j}{\pi_i} = Q_{ij}$$

Now  $P(Y_0 = i_0, \dots, Y_N = i_N) = P(X_0 = i_N, \dots, X_N = i_0)$

$$= \pi_{i_0} P_{i_0 i_1} \dots P_{i_{N-1} i_N} = \pi_{i_0} Q_{i_1 i_0} \dots Q_{i_N i_{N-1}}$$

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so  $P(Y_n = i_n | Y_{n-1} = i_{n-1}, \dots, Y_0 = i_0) = P(Y_n = i_n | Y_{n-1} = i_{n-1})$

which implies that  $Y_n$  is a Markov chain with transition probability  $Q = \{Q_{ij}\}$

Defn: If  $P_{ij} = \frac{\pi_j}{\pi_i} P_{ji}$  the Markov chain is time reversible.

An alternative expression is

$\pi_i P_{ij} = \pi_j P_{ji}$  all  $i, j$

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