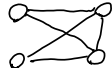


$n \rightarrow \infty$ X_n ergodic Markov chain
 $P(X_n=j) = \pi_j$, π_j stationary distribution
 Considered $Y_n = X_{N-n}$ $n = 0, \dots, N$
 Y_n is also a Markov chain and has transition probabilities $Q_{ij} = \frac{\pi_j}{\pi_i} P_{ji}$
 Definition 1 If $P_{ij} = \frac{\pi_j}{\pi_i} P_{ji}$ the Markov chain is reversible.
 Alternatively the condition is $\pi_i P_{ij} = \pi_j P_{ji}$
 Remark 4.8.1 $\pi_i P_{ij}$ is the rate where the backward chain moves from i to j
 $\pi_j P_{ji}$ is the rate for which the forward chain moves from j to i .
 That these two rates are equal is a strong assumption.
 Lemma. If $\exists \{x_j\}$ $x_j \geq 0$ $\sum x_i = 1$
 and $x_i P_{ij} = x_j P_{ji}$ all i, j
 then $x_i = \sum x_j P_{ji}$ so
 $x_i = \pi_i$ where π_i are the stationary or limiting probabilities.

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Proof. $\sum_j x_j P_{ji} = \sum_j x_i P_{ij} = x_i \sum_j P_{ij} = x_i$
 So since π is unique for ergodic chains
 $x_i = \pi_i$ all i
 Example Random walk on a graph.

 Random walk on a graph: Pick a node, then move to the next by picking a neighbor at random.
 Let $d_i = \#$ neighbor at node i
 $d = \sum d_i$, $P_{ij} = \frac{1}{d_i}$ j is i 's neighbor
 Let $\pi = (\frac{d_1}{d}, \dots, \frac{d_n}{d})$ $N = \#$ nodes.
 Is the chain reversible?
 $\pi_i P_{ij} = \begin{cases} \frac{d_i}{d} \cdot \frac{1}{d_i} = \frac{1}{d} & i, j \text{ neighbors} \\ 0 & \text{else} \end{cases}$
 $\pi_j P_{ji} = \begin{cases} \frac{d_j}{d} \cdot \frac{1}{d_j} = \frac{1}{d} & i, j \text{ neighbors} \\ 0 & \text{else} \end{cases}$
 Hence random walk on a graph is time reversible.

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4.9 Markov Chain Monte Carlo (MCMC)
 Consider X discrete take values at $\{1, 2, \dots\}$
 want $E h(X)$ same h .
 One option: Simulate independent X_1, X_2, \dots
 $\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n h(X_i) \rightarrow E h(X)$ under
 weak regularity by law large numbers.
 Although in principle possible this can be difficult for complicated distribution of X .
 Other option: Reall
 Proposition 4.6 p 417 If X_n is irreducible Markov chain with stationary probabilities π_j , $j = 0, 1, \dots$. Then for n -bounded $\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n h(X_k) \rightarrow \sum_{j=0}^{\infty} h(j) \pi_j$
 Thus if we can find a Markov chain X_n with stationary probabilities $\lim_{n \rightarrow \infty} P(X_n=j) = \pi_j$
 $E h(X) = \sum_{j=0}^{\infty} h(j) \pi_j = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n h(X_k)$

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The aim is therefore to construct Markov chain having the specified probabilities as stationary distribution.
 We will consider the case where $\pi_j = C \cdot b_j$ where $b_j, j = 1, 2, \dots$ are specified but C is unspecified.
 Let $Q = \{q(i, j)\}_{j=1}^{\infty}$ be any th specified Markov transition matrix on the integers $q(x, y)$ is called a proposal distribution, and should be easy to simulate.
 Use the algorithm 1. For $X_n = i$ generate Y
 $P(Y=j) = q(i, j)$
 2. If $Y=j$
 $X_{n+1} = \begin{cases} j & \text{with probability } \alpha(i, j) \\ i & \text{with probability } 1 - \alpha(i, j) \end{cases}$
 where $\alpha(i, j)$ is the acceptance probability (C to be specified)
 Remark 4.9.1 Possible for the chain X_n to remain in state i , i.e. $X_n = i$

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Now X_n has transition probabilities
 $P_{ij} = P(X_{n+1}=j | X_n=i) = \begin{cases} q(i, j) \alpha(i, j) & (i \neq j) \\ q(i, i) + \sum_{k \neq i} q(i, k) (1 - \alpha(i, k)) & (i = j) \end{cases}$
 Since $\sum_{j \neq i} P_{ij} + P_{ii} = \sum_{j \neq i} q(i, j) \alpha(i, j) + q(i, i) + \sum_{k \neq i} q(i, k) - \sum_{k \neq i} q(i, k) \alpha(i, k)$
 Is this chain X_n time reversible?
 i.e. $\pi(i) P_{ij} = \pi(j) P_{ji}$ $j \neq i$ or
 $\pi(i) q(i, j) \alpha(i, j) = \pi(j) q(j, i) \alpha(j, i)$ $j \neq i$

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Lemma: Let $\alpha(i, j) = \min \left\{ \frac{\pi(j) q(j, i)}{\pi(i) q(i, j)}, 1 \right\}$
 then $\textcircled{*}$ is satisfied.
 Proof. $\alpha(i, j) = \min \left\{ \frac{\pi(j) q(j, i)}{\pi(i) q(i, j)}, 1 \right\}$
 $\alpha(j, i) = \min \left\{ \frac{\pi(i) q(i, j)}{\pi(j) q(j, i)}, 1 \right\}$
 $\textcircled{*} \alpha(i, j) = \frac{\pi(j) q(j, i)}{\pi(i) q(i, j)} < 1$
 $\Rightarrow \frac{\pi(i) q(i, j)}{\pi(j) q(j, i)} > 1$ and $\alpha(j, i) = 1$
 So $\pi(i) q(i, j) \alpha(i, j) = \pi(i) q(i, j) \left[\frac{\pi(j) q(j, i)}{\pi(i) q(i, j)} \right] = \pi(j) q(j, i) = \pi(j) q(j, i) \alpha(j, i)$
 i.e. $\textcircled{*}$

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$\Rightarrow \frac{\pi(j)q(j,i)}{\pi(i)q(i,j)} > 1$
 $\Rightarrow \frac{\pi(i)q(i,j)}{\pi(j)q(j,i)} < 1$
 $\Rightarrow d(j,i) = \frac{\pi(i)q(i,j)}{\pi(j)q(j,i)}$
 $\Rightarrow \pi(i)q(i,j) \cdot d(j,i) = \pi(i)q(i,j) = \pi(j)q(j,i) \cdot d(j,i)$ i.e. π

Consider now $\pi(j) = \frac{b(j)}{B}$ $b(j)$ specified
 B unspecified.

Then $d(i,j) = \min \left\{ \frac{b(j)q(j,i)}{b(i)q(i,j)}, 1 \right\}$ since B

cancels in.

Remark 4.9.2 Usually $\pi(j)$ are not only stationary probabilities but will also be limiting probabilities.

The condition $P_{ij} > 0$ will suffice. The the chain cannot be aperiodic.

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Example Permutations of $\{1, \dots, n\}$ (x_1, \dots, x_n)
 $S = \{ (x_1, \dots, x_n) : \sum_{j=1}^n x_j > a \}$ $a > 0$ given.

How big is S ? Want to construct a Markov chain with $\pi_j = \frac{1}{\#S}$ $\#S =$ number of elements in S .

Each permutation in S is a node in a graph. Two nodes are connected if only two elements are changed ex. $(1, 2, \underline{3}, 4)$ and $(1, 2, \underline{4}, 3)$ are connected but not $(1, 2, \underline{3}, 4)$ and $(1, \underline{3}, 4, 2)$.

Proposed distribution: Let $N(s)$ be the neighbors of s and $\#N(s)$ the number of neighbors.

$q(s, \xi) = \frac{1}{\#N(s)}$ $\xi \in N(s)$
 Since $\pi_j = \frac{1}{\#S}$ $\pi(s) = \pi(\xi)$, then $d(s, \xi) = \min \left\{ \frac{q(s, \xi)}{q(\xi, s)}, 1 \right\} = \min \left\{ \frac{\#N(s)}{\#N(\xi)}, 1 \right\}$

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