

and $f(x) = \begin{cases} \lambda \exp(-\lambda t) & t \geq 0 \\ 0 & \text{else} \end{cases}$

Proposition If X_1, \dots, X_n independent $X_i \sim \lambda_i \exp(-\lambda_i t)$
then
 (i) $\min(X_1, \dots, X_n) \sim \lambda \exp(-\lambda t)$ where $\lambda = \sum \lambda_i$
 (ii) $\min(X_1, \dots, X_n)$ and the rank order of X_1, \dots, X_n are independent.

Proof (i)
 seen already
 (ii) Consider a particular ordering and $P(X_{i_1} < \dots < X_{i_n} | \min X_i) <$
 But on event $\min X_i > t$ all variables are larger than t . By the memoryless property the remaining life times are exponentially distributed with original rates.
 Hence $P(X_{i_1} - t < \dots < X_{i_n} - t | \min X_i) = P(X_{i_1} < \dots < X_{i_n})$ which does not depend on $\min X_i$.

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Example
 ① 
 Both clerks bring new customer assigned to first free slot clerk.
 T : time spent in office by new customer.
 $E(T)$?
 Introduce R_1, R_2 remaining service times
 $E[T] = E[T | R_1 < R_2] \cdot P(R_1 < R_2)$
 $+ E[T | R_2 < R_1] \cdot P(R_2 < R_1)$
 $= E[T | R_1 < R_2] \cdot \frac{\lambda_1}{\lambda_1 + \lambda_2} + E[T | R_2 < R_1] \cdot \frac{\lambda_2}{\lambda_1 + \lambda_2}$
 Write $T = \min(R_1, R_2) + S$, S service time of new customer
 $E[T | R_1 < R_2] = E[R_1 | R_1 < R_2] + S$
 $E[S | R_1 < R_2] =$
 $E[R_1 | \min R_i = R_1] + \frac{1}{\lambda_1} = \frac{1}{\lambda_1 + \lambda_2} + \frac{1}{\lambda_1}$

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Similarly $E[T | R_2 < R_1] = \frac{1}{\lambda_1 + \lambda_2} + \frac{1}{\lambda_2}$

so $E[T] = (\frac{1}{\lambda_1 + \lambda_2} + \frac{1}{\lambda_1}) \cdot \frac{\lambda_1}{\lambda_1 + \lambda_2} + (\frac{1}{\lambda_1 + \lambda_2} + \frac{1}{\lambda_2}) \cdot \frac{\lambda_2}{\lambda_1 + \lambda_2}$
 $= \frac{3}{\lambda_1 + \lambda_2}$

② alternative solution
 $T = \min(R_1, R_2) + S$
 $E[T] = \frac{1}{\lambda_1 + \lambda_2} + E[S]$
 $E[S] = E[S | R_1 < R_2] \cdot \frac{\lambda_1}{\lambda_1 + \lambda_2} + E[S | R_2 < R_1] \cdot \frac{\lambda_2}{\lambda_1 + \lambda_2}$
 $= \frac{1}{\lambda_1} \cdot \frac{\lambda_1}{\lambda_1 + \lambda_2} + \frac{1}{\lambda_2} \cdot \frac{\lambda_2}{\lambda_1 + \lambda_2} = \frac{2}{\lambda_1 + \lambda_2}$

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Definition: Suppose X_1, \dots, X_n independent $X_i \sim \lambda_i \exp(-\lambda_i t)$
 Then the random variable $X_1 + \dots + X_n$ is hyperexponential.
 Distribution of $X_1 + \dots + X_n$?
 First $n=2$ m.g.f. $X_1 + X_2 = \frac{\lambda_1}{\lambda_1+t} \cdot \frac{\lambda_2}{\lambda_2+t} \cdot t \lambda_1 \lambda_2$
 $\frac{\lambda_1}{\lambda_1+t} \cdot \frac{\lambda_2}{\lambda_2+t} + (\frac{\lambda_1}{\lambda_1+t})(\frac{\lambda_2}{\lambda_2+t}) = \frac{\lambda_1 \lambda_2}{(\lambda_1+t)(\lambda_2+t)} \left[\frac{\lambda_1 + \lambda_2 + t}{\lambda_1 + \lambda_2} \right]$
 so the n.d.f. of $X_1 + X_2$ is
 $\frac{\lambda_1}{\lambda_1+t} \cdot \lambda_1 \exp(-\lambda_1 t) + \frac{\lambda_2}{\lambda_2+t} \cdot \lambda_2 \exp(-\lambda_2 t)$
 $= \sum_{i=1}^2 \prod_{j \neq i} \frac{\lambda_j}{\lambda_j+t} \cdot \lambda_i \exp(-\lambda_i t)$
 Let $C_{i,j,n} = \prod_{j \neq i} \frac{\lambda_j}{\lambda_j+t}$

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Lemma If X_1, \dots, X_n independent $X_i \sim \lambda_i \exp(-\lambda_i t)$
 Then $\{X_{i+1}, \dots, X_n\}(x) = \sum_{i=1}^n C_{i,n} \lambda_i \exp(-\lambda_i t)$

Proof O.K. for $n=2$.
 Suppose O.K. for n . Then
 $\prod_{i=1}^{n+1} \frac{\lambda_i}{\lambda_i+t} = \left(\prod_{i=1}^n \frac{\lambda_i}{\lambda_i+t} \right) \frac{\lambda_{n+1}}{\lambda_{n+1}+t}$
 $= \left(\sum_{i=1}^n C_{i,n} \frac{\lambda_i}{\lambda_i+t} \right) \frac{\lambda_{n+1}}{\lambda_{n+1}+t}$
 $= \sum_{i=1}^n C_{i,n} \left[\frac{\lambda_i}{\lambda_{n+1}-\lambda_i} \frac{\lambda_{n+1}}{\lambda_{n+1}+t} + \frac{\lambda_{n+1}}{\lambda_{n+1}-\lambda_i} \frac{\lambda_i}{\lambda_{n+1}+t} \right]$
 $= K_{n+1} \frac{\lambda_{n+1}}{\lambda_{n+1}+t} + \sum_{i=1}^n C_{i,n} \frac{\lambda_i}{\lambda_{n+1}+t}$
 But $\prod_{i=1}^{n+1} \frac{\lambda_i}{\lambda_i+t} = \frac{\lambda_1}{\lambda_1+t} \cdot \prod_{i=2}^{n+1} \frac{\lambda_i}{\lambda_i+t}$
 $= K_1 \frac{\lambda_1}{\lambda_1+t} + \sum_{i=2}^{n+1} C_{i,n+1} \frac{\lambda_i}{\lambda_i+t}$

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This implies by considering the densities and multiplying with $\partial P(\min X_i) \cdot t$ that
 $K_{n+1} = C_{n+1, n+1}$ which concludes the proof

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