

# STK2130 – Lecture 3

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# Chapter 4 – Markov Chains

## Section 4.3 – Classification of States

## Section 4.3. Classification of States

Let  $\{X_n\}$  be a Markov chain with state space  $S$  and transition probability matrix  $\mathbf{P}$ .

State  $j$  is said to be **accessible** from state  $i$ , denoted as  $i \rightarrow j$ , if  $P_{ij}^n > 0$  for some  $n \geq 0$ .

Note that we have:

$$P_{ij}^n \leq P\left(\bigcup_{n=0}^{\infty} \{X_n = j\} \mid X_0 = i\right), \quad n = 0, 1, 2, \dots$$

Hence, we also have:

$$\begin{aligned} \sup_{n \geq 0} P_{ij}^n &\leq P\left(\bigcup_{n=0}^{\infty} \{X_n = j\} \mid X_0 = i\right) \\ &\leq \sum_{n=0}^{\infty} P\{X_n = j \mid X_0 = i\} = \sum_{n=0}^{\infty} P_{ij}^n. \end{aligned}$$

## Chapter 4 – Markov Chains

If  $i \rightarrow j$ , then by definition  $P_{ij}^n > 0$  for some  $n \geq 0$ , and hence we obviously also have that  $\sup_n P_{ij}^n > 0$ . This implies that:

$$P\left(\bigcup_{n=0}^{\infty} \{X_n = j\} \mid X_0 = i\right) \geq \sup_{n \geq 0} P_{ij}^n > 0.$$

Conversely, if  $P\left(\bigcup_{n=0}^{\infty} \{X_n = j\} \mid X_0 = i\right) > 0$ , then also:

$$\sum_{n=0}^{\infty} P_{ij}^n \geq P\left(\bigcup_{n=0}^{\infty} \{X_n = j\} \mid X_0 = i\right) > 0$$

which implies that  $P_{ij}^n > 0$  for some  $n \geq 0$ , i.e., that  $i \rightarrow j$ .

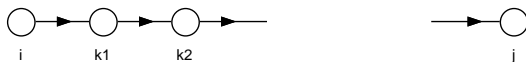
Hence, we conclude that  $i \rightarrow j$  if and only if:

$$P\left(\bigcup_{n=0}^{\infty} \{X_n = j\} \mid X_0 = i\right) > 0.$$

## Section 4.3. Classification of States (cont.)

A state diagram for a Markov chain is a directed graph where the nodes represent the states and the edges represent possible one-step transitions. More precisely, the state diagram contains an edge from node  $i$  to node  $j$  if and only if  $P_{ij} > 0$ .

If  $i, j \in \mathcal{S}$ , then  $i \rightarrow j$  if and only if the state diagram contains at least one directed path from  $i$  to  $j$ .



If such a path exists, we have:

$$P_{ij}^n \geq P_{i,k_1} \cdot P_{k_1,k_2} \cdots P_{k_{n-2},k_{n-1}} \cdot P_{k_{n-1},j} > 0.$$

# Communicating states

States  $i$  and  $j$  **communicate**, denoted as  $i \leftrightarrow j$ , if  $i \rightarrow j$  and  $j \rightarrow i$ .

The relation  $\leftrightarrow$  is an **equivalence relation**. That is  $\leftrightarrow$  satisfies the following properties:

- **Reflexivity**:  $i \leftrightarrow i$ .
- **Symmetry**:  $i \leftrightarrow j$  if and only if  $j \leftrightarrow i$ .
- **Transitivity**:  $i \leftrightarrow j$  and  $j \leftrightarrow k$  implies  $i \leftrightarrow k$ .

Reflexivity follows since we always have  $P_{ii}^0 = 1 > 0$ . Symmetry follows directly from the definition.

## Communicating states (cont.)

To prove transitivity we assume that  $i \leftrightarrow j$  and  $j \leftrightarrow k$ .

Hence, in particular  $i \rightarrow j$  and  $j \rightarrow k$ , implying that there exists  $m, n \geq 0$  such that  $P_{ij}^m > 0$  and  $P_{jk}^n > 0$ .

By the Chapman-Kolmogorov equations, we have:

$$P_{ik}^{m+n} = \sum_{r \in S} P_{ir}^m P_{rk}^n \geq P_{ij}^m \cdot P_{jk}^n > 0.$$

Hence, by definition  $i \rightarrow k$ .

By a similar argument we can show that  $k \rightarrow i$  as well.

Hence, we conclude that  $i \leftrightarrow k$ .

## Communicating states (cont.)

Two states that communicate are said to be **in the same (equivalence) class**.

Two classes of states are either **identical** or **disjoint**.

**PROOF:** Assume that  $\mathcal{A}, \mathcal{B} \subseteq \mathcal{S}$  represent two equivalence classes, and assume that  $\mathcal{A} \cap \mathcal{B} \neq \emptyset$ . That is, there exists a state  $i$  such that  $i \in \mathcal{A} \cap \mathcal{B}$ .

Then choose  $j \in \mathcal{A}$  and  $k \in \mathcal{B}$  arbitrarily.

Now,  $i, j \in \mathcal{A}$  implies that  $i \leftrightarrow j$  and  $i, k \in \mathcal{B}$  implies that  $i \leftrightarrow k$ .

Hence, by transitivity we also have  $j \leftrightarrow k$ . That is,  $j$  and  $k$  belong to the same equivalence class.

Since this holds for any  $j \in \mathcal{A}$  and  $k \in \mathcal{B}$ , this implies that  $\mathcal{A} = \mathcal{B}$  ■

The equivalence classes partition the state space  $\mathcal{S}$  into a number of disjoint sets. A Markov chain is called **irreducible** if the number of equivalence classes is **one**.



## Example 4.15

Consider a Markov chain with state space  $\mathcal{S} = \{0, 1, 2\}$  and transition probability matrix:

$$P = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} & 0 \\ \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \\ 0 & \frac{1}{3} & \frac{2}{3} \end{bmatrix}$$

We then observe:

Since  $P_{01} = \frac{1}{2} > 0$ , it follows that  $0 \rightarrow 1$

Since  $P_{10} = \frac{1}{2} > 0$ , it follows that  $1 \rightarrow 0$

Since  $P_{12} = \frac{1}{4} > 0$ , it follows that  $1 \rightarrow 2$

Since  $P_{21} = \frac{1}{3} > 0$ , it follows that  $2 \rightarrow 1$

Hence,  $0 \leftrightarrow 1$  and  $1 \leftrightarrow 2$ , and by transitivity  $0 \leftrightarrow 2$  as well. Thus, the Markov chain is irreducible.

## Example 4.15 (cont.)

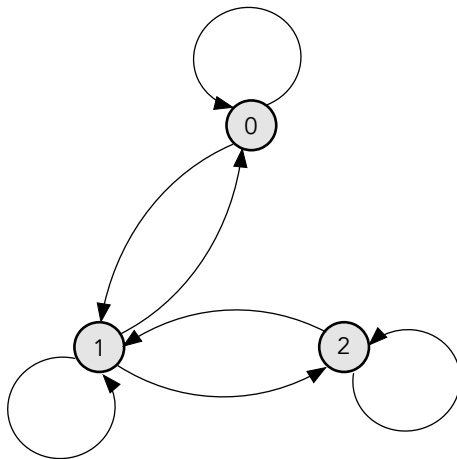


Figure: State diagram of an irreducible Markov chain with one class  $\{0, 1, 2\}$

## Example 4.16

A Markov chain with state space  $\mathcal{S} = \{0, 1, 2, 3\}$  and matrix:

$$P = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} & 0 & 0 \\ \frac{1}{2} & \frac{1}{2} & 0 & 0 \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

$$P_{01} = P_{10} = \frac{1}{2}, \quad \Rightarrow \quad 0 \leftrightarrow 1$$

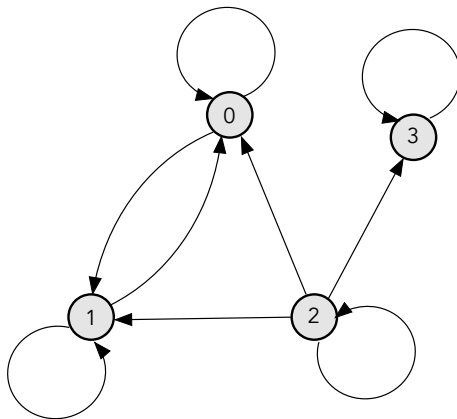
$$P_{0j} = P_{1j} = 0, \quad \Rightarrow \quad 0, 1 \not\leftrightarrow j, \quad j = 2, 3$$

$$P_{2i} = \frac{1}{4}, \quad \Rightarrow \quad 2 \rightarrow i, \quad i = 0, 1, 2, 3$$

$$P_{3i} = 0, \quad \Rightarrow \quad 3 \not\leftrightarrow i, \quad i = 0, 1, 2$$

The Markov chain has classes  $\{0, 1\}$ ,  $\{2\}$  and  $\{3\}$ , and is **not** irreducible.

## Example 4.16 (cont.)



**Figure:** State diagram of a Markov chain with three classes  $\{0, 1\}$ ,  $\{2\}$  and  $\{3\}$ .

# Recurrent and transient states

We consider the probabilities:

$$f_i = P\left(\bigcup_{r=1}^{\infty} \{X_r = i\} \mid X_0 = i\right), \quad i \in S.$$

- State  $i$  is **recurrent** if  $f_i = 1$ .
- State  $i$  is **transient** if  $f_i < 1$ .

Assume that  $X_0 = i$ , and let  $N_i$  denote the number of times  $X_n = i$ .

- If  $i$  is **recurrent**, then  $P(N_i = \infty \mid X_0 = i) = 1$ .
- If  $i$  is **transient**, then  $P(N_i = n \mid X_0 = i) = f_i^{n-1}(1 - f_i)$ ,  $n = 1, 2, \dots$

If  $i$  is transient and  $X_0 = i$ , then  $N_i \mid X_0 = i$  has a **geometric distribution** with  $E[N_i \mid X_0 = i] = 1/(1 - f_i) < \infty$ .

## Proposition 4.1

Let  $I_i^{(n)} = I(X_n = i)$ ,  $n = 0, 1, \dots$ . We can then write:

$$N_i = \sum_{n=0}^{\infty} I_i^{(n)}$$

Hence, we have:

$$\begin{aligned} E[N_i | X_0 = i] &= \sum_{n=0}^{\infty} E[I_i^{(n)} | X_0 = i] \\ &= \sum_{n=0}^{\infty} P[X_n = i | X_0 = i] = \sum_{n=0}^{\infty} P_{ii}^n \end{aligned}$$

- State  $i$  is **recurrent**, if  $\sum_{n=1}^{\infty} P_{ii}^n = \infty$ .
- State  $i$  is **transient**, if  $\sum_{n=1}^{\infty} P_{ii}^n < \infty$ .

## Corollary 4.2

If state  $i$  is recurrent, and  $i \leftrightarrow j$ , then state  $j$  is recurrent as well. Thus, recurrence is a **class property**.

PROOF: Since  $i \leftrightarrow j$ , there exists  $k$  and  $m$  such that  $P_{ij}^k > 0$  and  $P_{ji}^m > 0$ .

Hence, for any  $n = 1, 2, \dots$  we have:

$$P_{jj}^{m+n+k} \geq P_{ji}^m \cdot P_{ii}^n \cdot P_{ij}^k.$$

Summing over all  $n$ , and using that  $i$  is recurrent,  $P_{ij}^k > 0$  and  $P_{ji}^m > 0$  we get:

$$\sum_{n=1}^{\infty} P_{jj}^{m+n+k} \geq P_{ji}^m \cdot P_{ij}^k \cdot \sum_{n=1}^{\infty} P_{ii}^n = \infty$$

Hence, we conclude that  $j$  is recurrent as well ■

## Corollary 4.2 (cont.)

- If state  $i$  is transient and  $i \leftrightarrow j$ , then state  $j$  must also be transient. For if  $j$  were recurrent then, by Corollary 4.2,  $i$  would also be recurrent contradicting that  $i$  is transient. Thus, transience is a **class property** as well.
- If  $\{X_n\}$  is a Markov chain with a **finite** state space, then at least one of the states must be **recurrent**. If  $\{X_n\}$  is **irreducible** as well, then all states are **recurrent**.



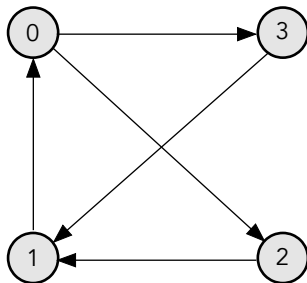
## Example 4.17

Consider a Markov chain with state space  $\mathcal{S} = \{0, 1, 2, 3\}$  and transition probability matrix:

$$P = \begin{bmatrix} 0 & 0 & \frac{1}{2} & \frac{1}{2} \\ 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 1 & 0 & 0 \end{bmatrix}$$

It is easy to verify that  $i \leftrightarrow j$  for all  $i, j \in \mathcal{S}$ . Hence, the Markov chain is **irreducible** and thus all states must be **recurrent** ■

## Example 4.17 (cont.)



**Figure:** State diagram of an irreducible Markov chain with one class  $\{0, 1, 2, 3\}$

## Example 4.18

Consider a Markov chain with state space  $\mathcal{S} = \{0, 1, 2, 3, 4\}$  and transition probability matrix:

$$P = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} & 0 & 0 & 0 \\ \frac{1}{2} & \frac{1}{2} & 0 & 0 & 0 \\ 0 & 0 & \frac{1}{2} & \frac{1}{2} & 0 \\ 0 & 0 & \frac{1}{2} & \frac{1}{2} & 0 \\ \frac{1}{4} & \frac{1}{4} & 0 & 0 & \frac{1}{2} \end{bmatrix}$$

This chain has classes  $\{0, 1\}$ ,  $\{2, 3\}$  and  $\{4\}$ .

The first two classes are **recurrent** and the third **transient** ■

## Example 4.18 (cont.)

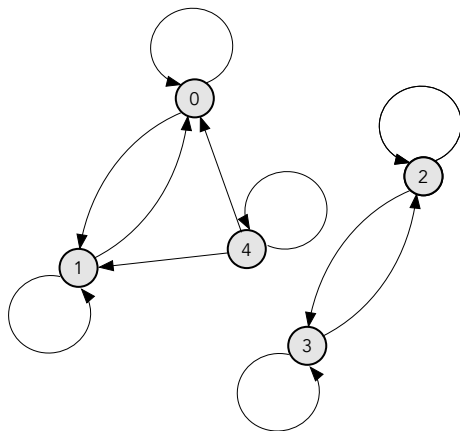


Figure: State diagram of a Markov chain with classes  $\{0, 1\}$ ,  $\{2, 3\}$  and  $\{4\}$

## Example 4.19 - Random walk

Consider a Markov chain with state space  $\mathcal{S} = \{\dots, -2, -1, 0, 1, 2, \dots\}$  and where  $0 < p < 1$  and:

$$P_{i,i+1} = p, \quad P_{i,i-1} = (1-p), \quad i \in \mathcal{S}.$$

It is obvious that  $i \leftrightarrow j$  for all  $i, j \in \mathcal{S}$ . Hence, according to Corollary 4.2 all states are either recurrent or transient.

In order to check for recurrence, it is sufficient to check if  $\sum_{n=1}^{\infty} P_{00}^n = \infty$ .

Thus, we assume that  $X_0 = 0$ , and observe that in this case  $X_n$  is **odd** if  $n$  is **odd**, and  $X_n$  is **even** if  $n$  is **even**. Hence, since 0 is **even**, we have:

$$P_{00}^{2n-1} = 0, \quad n = 1, 2, \dots$$

$$P_{00}^{2n} = \binom{2n}{n} p^n (1-p)^n = \frac{(2n)!}{n!n!} [p(1-p)]^n, \quad n = 1, 2, \dots$$

## Example 4.19 - Random walk (cont.)

We then use Stirling's formula for  $n!$ :

$$n! \approx n^{n+1/2} e^{-n} \sqrt{2\pi}$$

From this we get:

$$\frac{(2n)!}{n!n!} \approx \frac{(2n)^{2n+1/2} e^{-2n} \sqrt{2\pi}}{(n^{n+1/2} e^{-n} \sqrt{2\pi})^2} = \frac{(2n)^{2n+1/2} e^{-2n} \sqrt{2\pi}}{n^{2n+1} e^{-2n} (2\pi)} = \frac{2^{2n}}{\sqrt{n\pi}} = \frac{4^n}{\sqrt{n\pi}}$$

Hence:

$$P_{00}^{2n} = \frac{(2n)!}{n!n!} [p(1-p)]^n \approx \frac{(4p(1-p))^n}{\sqrt{n\pi}}$$

## Example 4.19 - Random walk (cont.)

This implies that:

$$\sum_{n=1}^{\infty} P_{00}^{2n} \approx \sum_{n=1}^{\infty} \frac{(4p(1-p))^n}{\sqrt{n\pi}}$$

This series is divergent if and only if  $p = \frac{1}{2}$ .

Hence, the states are **recurrent** if and only if  $p = \frac{1}{2}$ .