

Poisson processes

Def: N taking values in $\{0, 1, 2, \dots\}$ is a Poisson with rate λ if

$$P[N = n] = e^{-\lambda t} \frac{\lambda^n}{n!}$$

Mean: $E[N] = \lambda$

Variance: $\text{Var}(N) = \lambda$.

Thm: If N_1, N_2, \dots, N_n are Poisson r.v. with rates $\lambda_1, \lambda_2, \dots, \lambda_n$, then

$$N = N_1 + N_2 + \dots + N_n$$

is Poisson with rate $\lambda = \lambda_1 + \lambda_2 + \dots + \lambda_n$

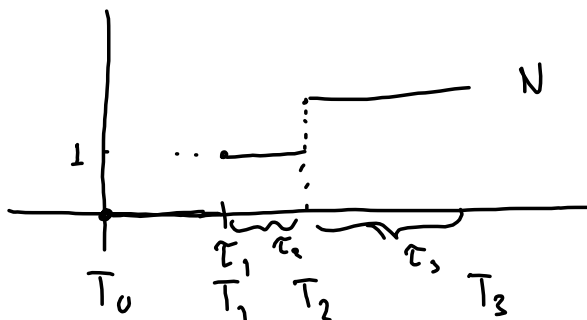
Def: A stochastic process $\{N_t\}_{t \in [0, \infty)}$ is a Poisson process with rate λ if

(i) $N(0) = 0$

(ii) N has independent increments

(iii) $N(t) - N(s)$ is Poisson disto. with rate $\lambda(t-s)$.

Dynamical picture: τ_1, τ_2, τ_3 are independent, exponential r.v. with rate λ :



$$T_n = \tau_1 + \tau_2 + \dots + \tau_n \quad (\text{n-th jump})$$

$$N(t, \omega) = \max \{n : T_n \leq t\}$$

Mean: $E[N(t) - N(s)] = \lambda(t-s)$

$$\text{Var}[N(t) - N(s)] = \lambda(t-s)$$

Nonhomogeneous Poisson processes

Rate function: $\lambda(t)$

Mean value function: $m(t) = \int_0^t \lambda(s) ds$

Def: N is a nonhom. Poisson process with rate function λ if

(i) $N(0) = 0$

(ii) N has independent increments

(iii) $N(t) - N(s)$ is Poisson distributed with ^{rate} $m(t) - m(s) = \int_s^t \lambda(u) du$

Infinitesimal description:

(i) $N(0) = 0$

(ii) N has ind. increments

(iii) $P[N(t+h) - N(t) = 1] = \lambda(t)h + o(h)$

(iv) $P[N(t+h) - N(t) > 1] = o(h)$

Example



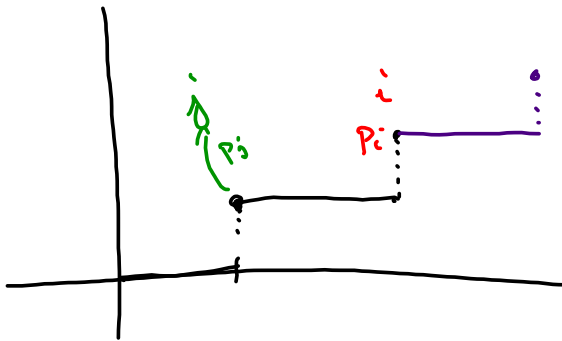
Operations on Poisson process

Adding: N_1, N_2, \dots, N_n independent Poisson process with rates $\lambda_1, \lambda_2, \dots, \lambda_n$, then

$$N(t) = N_1(t) + N_2(t) + \dots + N_n(t)$$

is a Poisson proc with rate $\lambda_1 + \lambda_2 + \dots + \lambda_n$.

Thinning: Categories: $1, 2, \dots, k$. With probability p_i the jump belongs to category i .



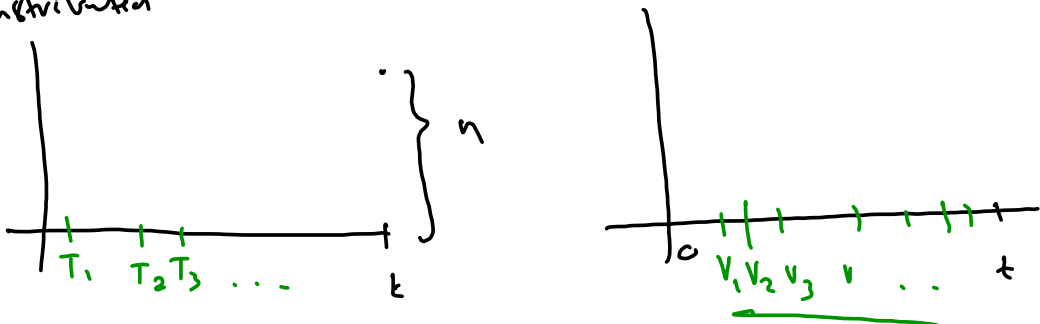
$N_j(t)$ = number of jumps before of type j

Poisson process with rate $p_i \lambda$.

Example: Cinema
 children $\rightarrow 0.4$
 adults $\rightarrow 0.6$

$N(t)$ with rate $\lambda \rightarrow N_c$ poisson with rate 0.4λ
 $\rightarrow N_a$ " " " 0.6λ .

Conditioning: If we know that $N(t) = n$, how are T_1, T_2, \dots, T_n distributed



Choose n random points from the interval and order them $V_1 < V_2 < \dots < V_n$. Then (V_1, V_2, \dots, V_n) has the same distribution as (T_1, T_2, \dots, T_n) .

Compound Poisson Processes

Poisson process N

$\gamma_1, \gamma_2, \gamma_3, \dots$ are iid and independent of N .

$$S(t) = \gamma_1(t) + \gamma_2(t) + \dots + \gamma_{N(t)}(t)$$

Example: Car insurance company:

Claims come as Poisson process N .

γ_i is the random claim for accident no i .

$$S(t) = \gamma_1(t) + \dots + \gamma_{N(t)}(t) \leftarrow \text{total claims by time } t. \text{ with rate } \lambda$$

Theorem: If N is a Poisson random variable and $\gamma_1, \gamma_2, \dots$ are iid and independent of N . If

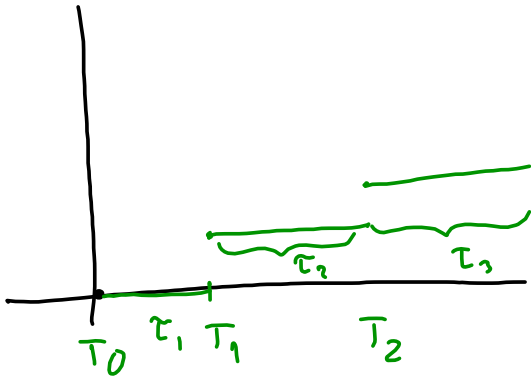
$$S = \gamma_1 + \gamma_2 + \dots + \gamma_{N_1}$$

then

$$E[S] = E[N] E[\gamma_i]$$

$$\text{var}(S) = \lambda E[\gamma_i^2]$$

Renewal process



Fix i.i.d random variable $E[\tau_i] < \infty$, and make the process jump as pictured below.

$$N(t) = \max\{n : T_n \leq t\}$$

Thm: If $P[\tau_i > 0] > 0$ and $\mu = E[\tau_i] < \infty$, then

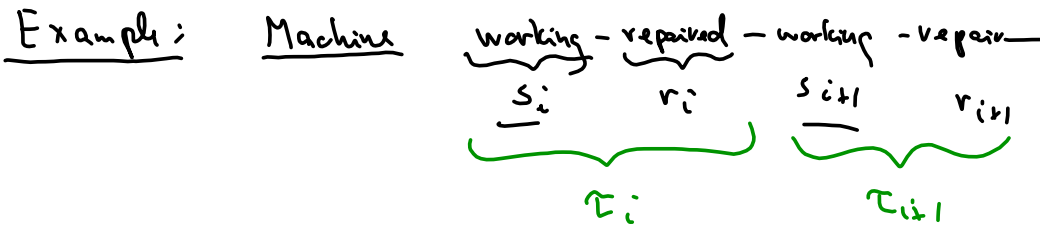
$$\frac{N(t)}{t} \rightarrow \frac{1}{\mu} \quad \text{with prob. 1}$$

Why $\mu N(t) \sim t$

Reward process: For each jump there is a reward r_i (i.i.d).

$$R(t) = \sum_{i=1}^{N(t)} r_i$$

$$\frac{R(t)}{t} \rightarrow \frac{E[r_i]}{E[\tau_i]}$$



Reward: s_i

$S(t) = \sum s_i$ = the period the machine working

$$\frac{S(t)}{t} \rightarrow \frac{E[s_i]}{E[\tau_i]} = \frac{E[s_i]}{E[s_i] + E[r_i]}$$

Problem 3, 2017

a) People coming in Poisson 4 per minute \leftarrow *could time minutes.*

Arriving exponentially with rate 4: Density $4e^{-4t}$

b) Time between lifts 5 seconds = $\frac{1}{12}$ minute.

Arrivals are Poisson with rate $\lambda \cdot t = 4 \cdot \frac{1}{12} = \frac{1}{3}$

$$P[\text{customers} = n] = e^{-\frac{1}{3}} \frac{\left(\frac{1}{3}\right)^n}{n!} = e^{-\frac{1}{3}} \frac{1}{3^n n!}$$

c) Prob one person leaving + one left waiting

$$= P[\text{customers} = 2] = e^{-\frac{1}{3}} \frac{1}{3^2 \cdot 2} \hat{=}$$

d) Prob of second lift leaving empty?

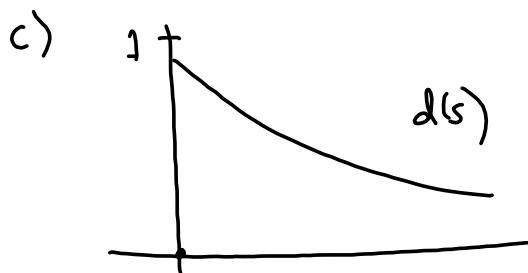
two ways

(1) no customers have arrived before second lift leaving

(2) one customer during first period, none during second.

$$\begin{aligned} & P[\text{customer} = 0]^2 + P[\text{customer} = 1] P[\text{customer} = 0] \\ &= \left(e^{-\frac{1}{3}}\right)^2 + \frac{e^{-\frac{1}{3}}}{3} \cdot e^{-\frac{1}{3}} = e^{-\frac{2}{3}} \left(1 + \frac{1}{3}\right) = \frac{4}{3} e^{-\frac{2}{3}} \end{aligned}$$

Prct: 2009



$$\begin{aligned}
 P[\Sigma_1 > t] &= \sum_{n=0}^{\infty} P[\Sigma_1 > t | N(t) = n] P[N(t) = n] && \sum_{n=0}^{\infty} e^{-\lambda} \frac{\lambda^n}{n!} = 1 \\
 &= \sum_{n=0}^{\infty} \left(\frac{t - D(t)}{t} \right)^n e^{-\lambda t} \frac{(\lambda t)^n}{n!} && \sum_{n=0}^{\infty} \frac{x^n}{n!} = e^x \\
 &= e^{-\lambda t} \sum_{n=0}^{\infty} \frac{(\lambda(t - D(t)))^n}{n!} = e^{-\lambda t} e^{\lambda(t - D(t))} && \sum_{n=0}^{\infty} \frac{\lambda^n}{n!} = e^{\lambda} \\
 &= e^{-\lambda D(t)} &&
 \end{aligned}$$

$$P[\Sigma_1 \leq t] = 1 - e^{-\lambda D(t)}$$

d) Explain that

$$P[\Sigma_1 > t | N(t) = n] = \left(\frac{t - D(t)}{t} \right)^n$$

has made n jumps - uniformly distributed.

Look at the $n=1$ case:

$$\begin{aligned}
 P[\Sigma_1 > t | N(t) = 1] &= \frac{1}{t} \int_0^t (1 - d(s)) ds = \frac{1}{t} [t - D(t)] \\
 &= \frac{t - D(t)}{t}
 \end{aligned}$$

What about n :

$$P[\Sigma_1 > t | N(t) = n] = \left(\frac{t - D(t)}{t} \right)^n \text{ by independence.}$$