

STK 2130 - 3. april 2024

2.27 Rock concert with arrivals @ rate 50
→ 30 are female
20 are male (*)

→ THINNING

"Main" Poisson process with rate 50
splits into two independent
PP with rates 30 and 20

a) What is the probability for all 3 of the
first 3 arrivals to be females?

→ recommend: turn off STK 2130 brain,
it is a simple question.

From rates (*) → 3:2 ratio of females to
males.

→ chance of first guest being female = $\frac{3}{5}$.

— " — second — " — = $\frac{3}{5}$

EVEN GIVEN
THAT 1st guest is
female

(Markov property)

— " — third — " — = $\frac{3}{5}$

$$P(\text{first 3 guests female}) = \left(\frac{3}{5}\right)^3$$

b) Same situation: arrivals to concert @ $\begin{cases} 20 \text{ males} \\ 30 \text{ females} \end{cases}$

$$P(\text{both guests arriving in } [0, 3] \mid \text{both arrived in } [0, 5])$$

ONE WAY:

$$= \frac{P(\{\text{both guests in } [0, 3]\} \cap \{\text{both guests in } [0, 5]\})}{P(\text{both in } [0, 5])}$$

=... the STK1100 way.

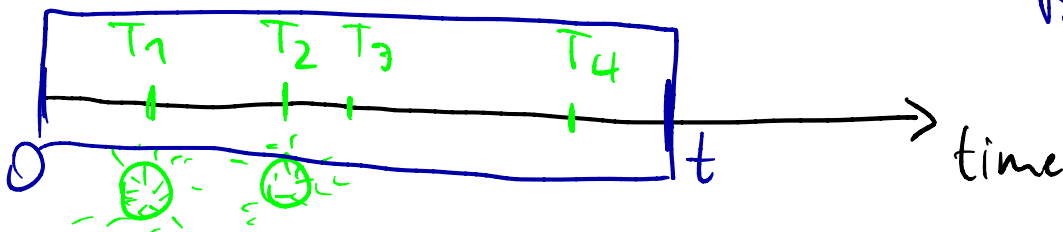
BUT:

THEOREM 2.15 (in the book)

Let T_1, T_2, T_3, \dots be arrival times of a PP
and $U_1, U_2, \dots, U_n \sim \text{Unit } [0, t]$, U_1, U_2, \dots, U_n indep.

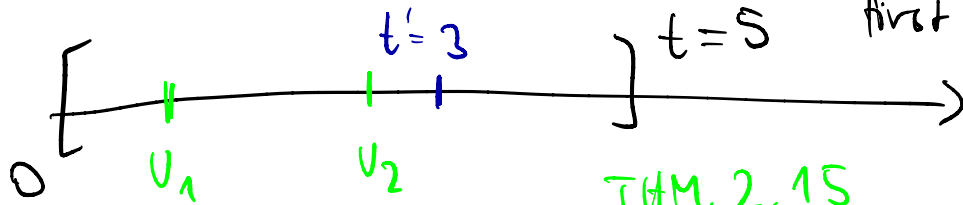
Conditioning on $N(t) = n$, (T_1, T_2, \dots, T_n)
has the same distribution as (U_1, U_2, \dots, U_n) .

$[0, t] \rightarrow$ impose that there are 4 arrival times



b) (cont'd)

$U_1, U_2 \sim \text{Unit}$



because it is given that both guests arrived in the first 5 minutes

THM. 2.15

$$\underbrace{P(U_1 \leq 3, U_2 \leq 3)}_{\text{THM. 2.15}} = P(\text{both guests arriving in } [0, 3] \mid \text{both arrived in } [0, 5])$$

U_1, U_2 independent,

$U_1, U_2 \sim \text{Unit}([0, 5])$

$$\stackrel{\text{indep.}}{=} P(U_1 \leq 3) \cdot P(U_2 \leq 3)$$

$$= \left(\frac{3}{5}\right) \cdot \left(\frac{3}{5}\right) = \underline{\underline{\left(\frac{3}{5}\right)^2}}$$

c) Still in the same setting:

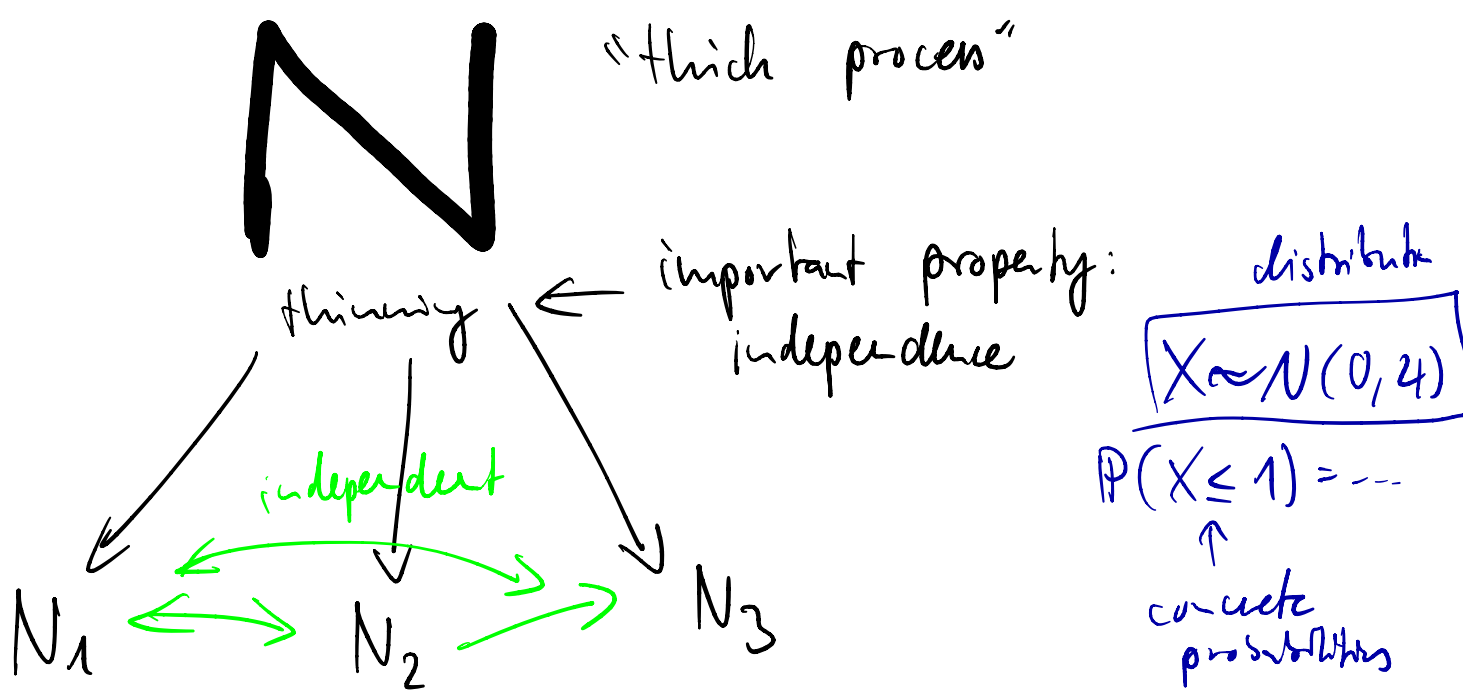
rock concert with PP arrival times @ rate 50.

Now, instead of $50 = \begin{cases} 30 \text{ female} \\ 20 \text{ male} \end{cases}$, split as

$50 = \begin{cases} 25 \text{ buy 1 ticket} \\ 20 \text{ buy 2 tickets} \\ 5 \text{ buy 3 tickets} \end{cases}$ (a different way of thinking the PPP)

$N_i :=$ PPP of buying i tickets

What is distribution of (N_1, N_2, N_3) ?



Independence of (N_1, N_2, N_3) implies that the joint distribution is simply three indep.

Poisson point processes with expected values

25, 20 and 5. $(N_1, N_2, N_3) \sim (PPP_1(25), PPP_2(20), PPP_3(5))$

$P(N_2(30 \text{ min}) = 4) = \dots$

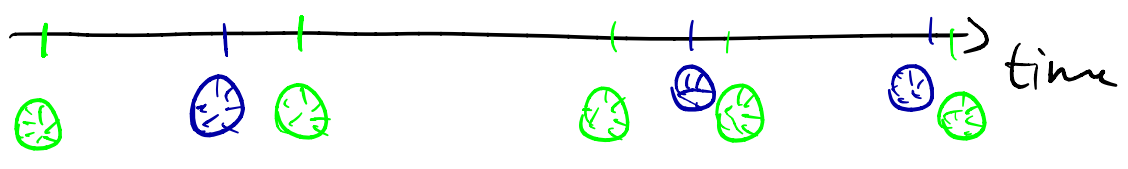
2.28

2 ways in which lightbulbs can be replaced:

failure $\sim \exp(200)$

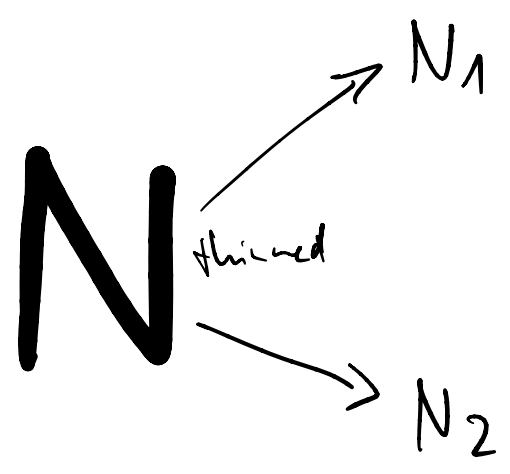
preventive replacement $\sim \exp(100)$

} arrival times, independently of each other



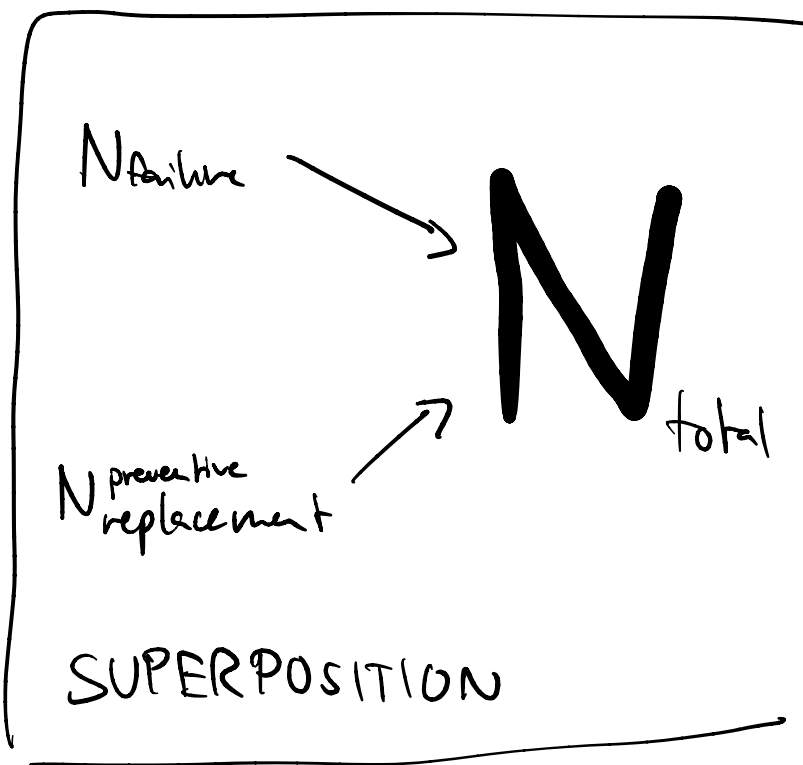
N_{failure}
 $N_{\text{preventive replacement}}$

} are both PPP, independent ones, with rates 200 and 100.



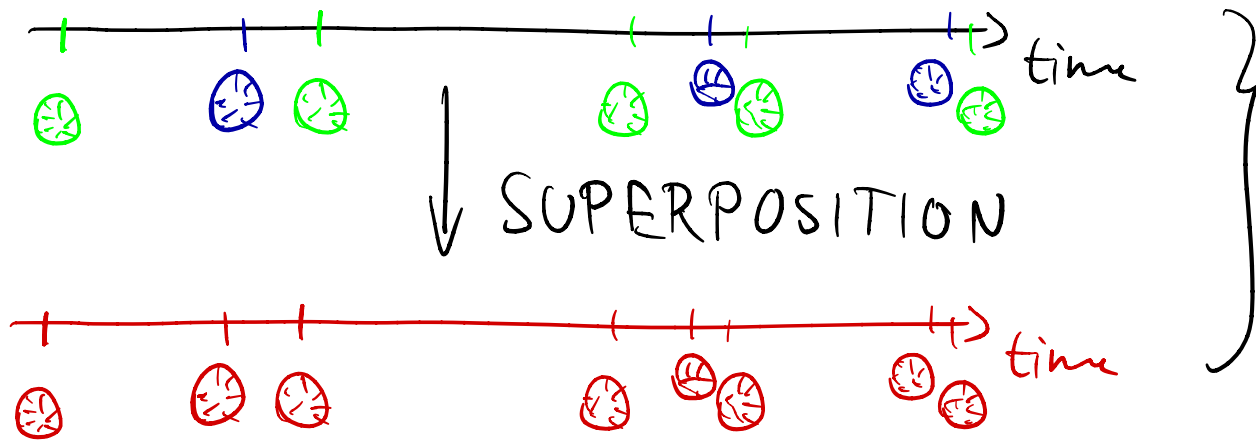
THINNING

→ prev. exercise



SUPERPOSITION

→ current exercise



"CLOCKS ASSUME THE SAME COLOUR"

(Failure + preventive replacement) \implies replacement

rates are added

⚠ careful with terminology \rightarrow expectations of the PPP get added

sanity check: does the expected number of events go up due to superposition (i.e. greater number of red clocks than green clocks?)

a) N_{failure} has expectation $\frac{1}{200}$ (expected arrivals per day)

$N_{\text{preventive replacement}}$ has expectation $\frac{1}{100}$ (expected arrivals per day)

$$\Rightarrow N_{\text{total}} = \frac{1}{200} + \frac{1}{100} = \frac{3}{200} \approx \frac{1}{67}$$

expected arrivals per day

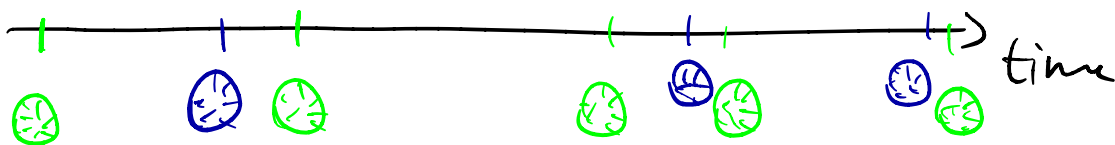
(depending on terminology \rightarrow new "rate" is 67 days)

b) in the long run, what fraction of replacements are due to failure?

\rightarrow compare exercise 2.27 a) $50 = \begin{cases} 30 \text{ female} \\ 20 \text{ male} \end{cases}$

$$\text{rate } \frac{30}{50} = \frac{3}{5}$$

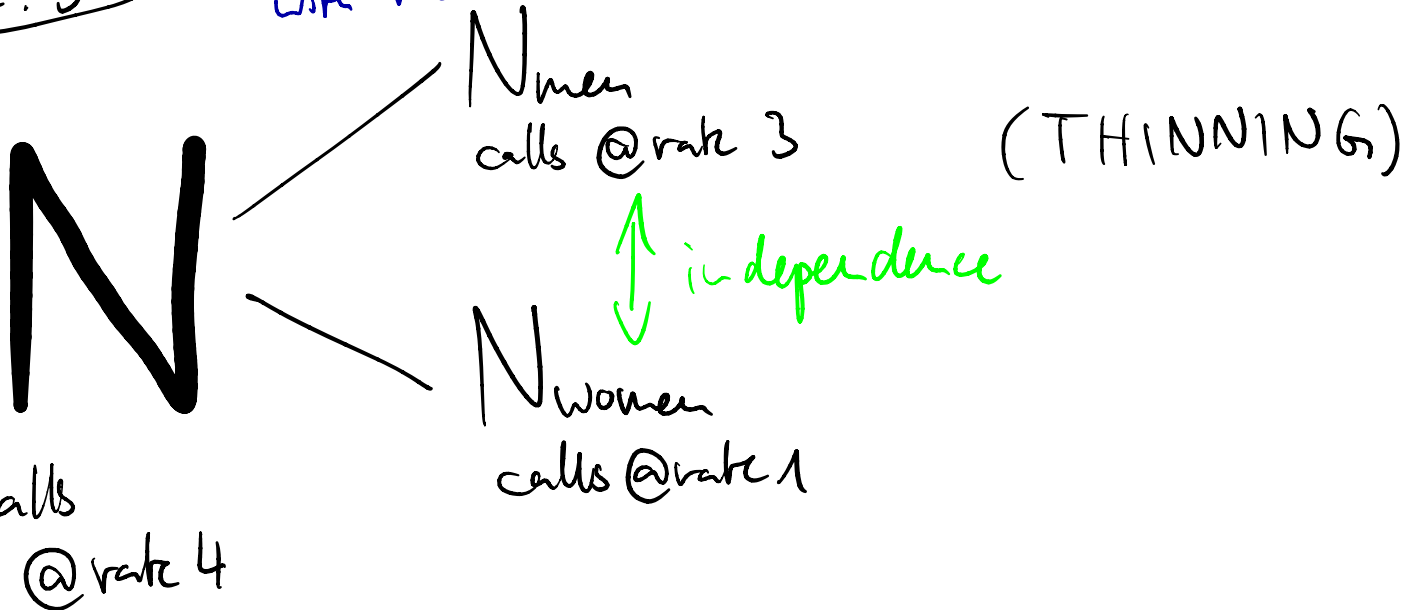
$$\frac{1}{67} = \begin{cases} \frac{1}{200} \text{ failure} \\ \frac{1}{100} \end{cases} \Rightarrow \text{share } \frac{\frac{1}{200}}{\frac{1}{67}} \approx \frac{1}{3}$$



Another way to understand the question:
What is the long term share of blue events out of all events.

2.30

N is a PPP with rate $\lambda \Rightarrow N(t) \sim \text{Poi}(\lambda t)$ (*)



a)

$$\mathbb{P}(N_{\text{men}}(t=1h) = 2, N_{\text{women}}(t=1h) = 3)$$

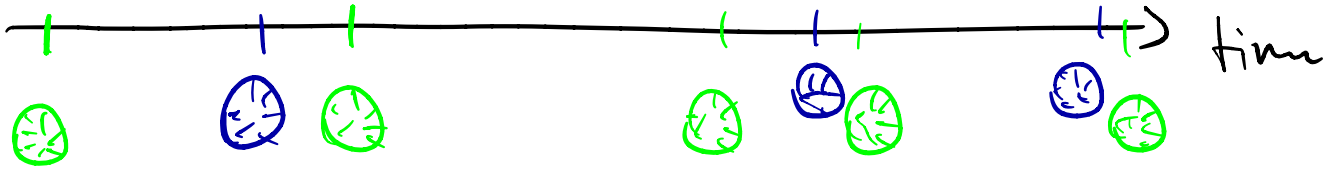
indep.

$$= \mathbb{P}(N_{\text{men}}(t=1h) = 2) \cdot \mathbb{P}(N_{\text{women}}(t=1h) = 3)$$

$$= e^{-3 \cdot 1} \cdot \frac{3 \cdot 1}{2!} \cdot e^{-1 \cdot 1} \cdot \frac{(1 \cdot 1)^3}{3!}$$

(rate 3) → 3
 (rate 1) → 1, (one hour) → 1
 (rate 1) → 1, (one hour) → 1

b) What is the probability that 3 men will have made calls before 3 women have?



P (at least three out of the first five calls are green)

$= P(\text{five calls out of five are green})$

$+ P(\text{four} \dots)$

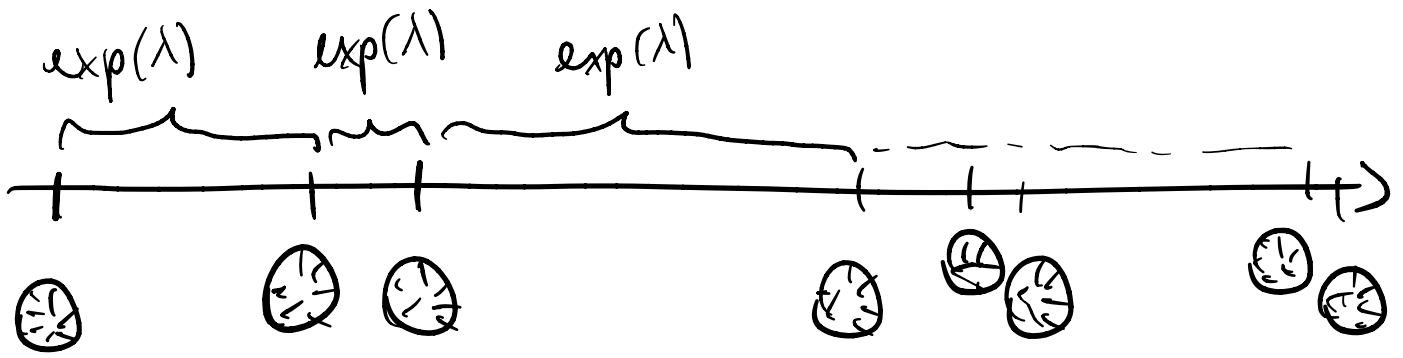
$+ P(\text{three} \dots)$

$$= \text{Bin}_{5, 3/4}(5) + \text{Bin}_{5, 3/4}(4) + \text{Bin}_{5, 3/4}(3)$$

$$= \binom{5}{3} \cdot \left(\frac{3}{4}\right)^3 \left(\frac{3}{4}\right)^2 + \binom{5}{4} \left(\frac{3}{4}\right)^4 \left(\frac{3}{4}\right) + \binom{5}{5} \left(\frac{3}{5}\right)^5$$

$= \dots$

RENEWAL PROCESS



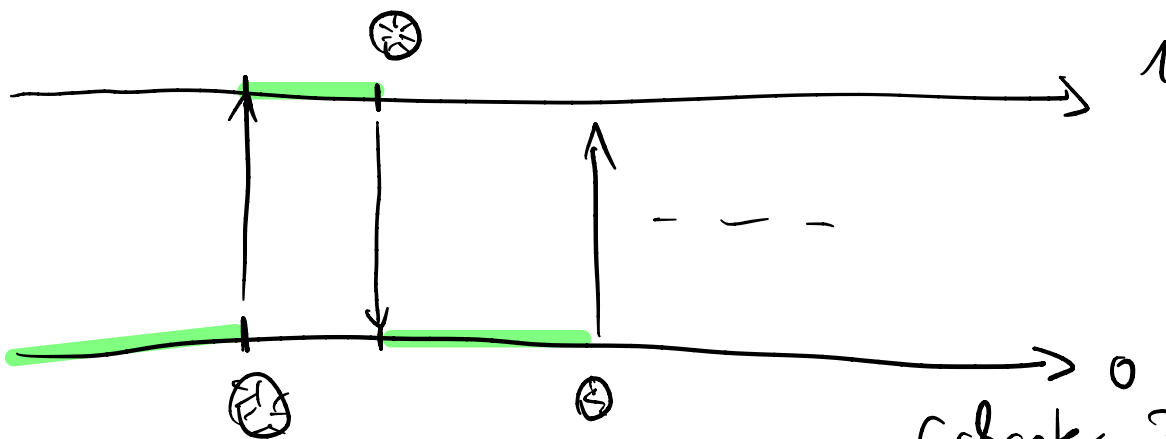
PPP is characterised by exponential waiting times.

Drop (only) this assumption

→ get some "generalised" form of PPP

RENEWAL PROCESS

MOST IMPORTANT APPLICATION:
JUMPING BETWEEN STATES



(chapter 3 of book)

Q: HOW MUCH TIME ON AVERAGE SPENT IN STATE 1 OR 0? → ANSWERS IN THE Book