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3.2 Force of mortality (dødsintensitet)

force of mortality of (x) at x+t :

$$\mu_{x+t} := \frac{g(t)}{1-G(t)} = -\frac{d}{dt} \log(1-G(t))$$

→ measure for the speed of mortality

Since

$$g(t) = \frac{1-G(t)}{ePx} \cdot \mu_{x+t}$$

(3.1.5) implies that

$$e_x = \int_0^{\infty} t \cdot ePx \cdot \mu_{x+t} dt \quad (3.2.1)$$

Further, since

$$\mu_{x+t} = -\frac{d}{dt} \log ePx$$

we get by integration that

$$ePx = e^{-\int_0^{\infty} \mu_{x+s} ds} \quad (3.2.2)$$

Rem. 3.2.1 : because of (3.1.4) we observe that

$$sq_{x+t} \approx \mu_{x+t} \cdot s \quad \text{for small } s$$

Now let us discuss some explicit distributions for T

3.3 Analytical distributions of T

→ explicit formula for G

→ favourable from the computational point of view

Examples :

1. De Moivre (1724) :

$$g(t) = \frac{1}{w-x} \quad 0 < t < w-x \quad (\text{uniform distr. on } [0, w-x])$$

⇒ w maximum age

$$\mu_{x+t} \stackrel{\text{def}}{=} \frac{g(t)}{1-G(t)} = \frac{1}{w-x-t} \quad 0 < t < w-x$$

2. Gompertz (1824) :

$$\mu_{x+t} := B c^{x+t} \quad t > 0$$

3. Makeham (1860) :

$$\mu_{x+t} := A + B c^{x+t} \quad t > 0$$

(3.2.2)

indep. of age

$$ePx = \exp(-At - m c^x (c^t - 1)) \quad m = B / \log c$$

4. Weibull (1939) :

$$\mu_{x+t} = K (x+t)^{n-1}$$

! Exerc. : Calc. the survival prob.

(14) For later use we have to introduce some r.v.'s related to T

3.4 Curtate future lifetime of (x)

1. K curtate future lifetime of (x)

$K := \lfloor T \rfloor$ ← Gauss bracket; rounds off to the next integer

e.g. $\lfloor 1.7 \rfloor = 1$, $\lfloor 2.9 \rfloor = 2$, $\lfloor 3.1 \rfloor = 3$

→ # of completed future years lived by (x)

def: $\Pr(K=j) = \Pr(j \leq T < j+1) = j p_x \cdot q_{x+j}$ (3.4.1)

surv. prob. ↓
↑
could prob. to die within 1 year after having reached $x+j$ years

expected value of K :

$e_x := E[K] = \sum_{j \geq 1} j \Pr(K=j)$ (3.4.1) = $\sum_{j \geq 1} j \cdot j p_x \cdot q_{x+j}$ (3.4.2)

or $e_x = \sum_{j \geq 1} \Pr(K \geq j) = \sum_{j \geq 1} j p_x$

2. S → $S := T - K \in [0, 1)$
 Rem. 3.4.1: (i) S has cont. distr. on $[0, 1]$
 (ii) $e_x \approx e_x + \frac{1}{2}$

3. $s^{(m)}$: $s^{(m)} := \frac{1}{m} \lfloor m \cdot S + 1 \rfloor$ ← Gauss bracket
 rounds to the next higher multiple of $\frac{1}{m}$
 e.g. $S = \frac{2}{m} + \frac{1}{2m} \rightarrow s^{(m)} = \frac{3}{m}$

Next we want to give an approximation of the distribution of T by means of one-year survival probabilities

3.5 Approximation of $G(t)$

Consider 1-year survival prob.

p_{x+j} $j=0,1,2,\dots$

given by a life table (see later)

→ ${}_k p_x$, $k \in \mathbb{N}$ known, since

${}_k p_x = p_x p_{x+1} p_{x+2} \dots p_{x+k-1}$, $k \in \mathbb{N}$

→ objective: approximation of $G(t) = \Pr(T \leq t)$

→ method: interpolation of q_x or p_{x+u} at intermediate ages $x+u$, $x \in \mathbb{N}_0$, $0 < u < 1$

(15) e.g. if M_{x+t} known by interpolation (B.Z.2) $\int_0^t M_{x+s} ds$ known for all t
 \Rightarrow ${}_t p_x = 1 - \frac{1}{e} \int_0^t M_{x+s} ds$ known for all t

1. method :

assumption : (i.e. $u q_{x+t} = u q_{x+j}$)
 $u q_x = u \cdot q_x, 0 \leq u \leq 1$ (linearity of $u q_x$)

$$\Rightarrow u p_x = 1 - u q_x$$

$$\stackrel{\text{def.}}{\Rightarrow} M_{x+t} = \frac{q_x}{1 - u q_x} \quad (3.5.1)$$

Rem. 3.5.1 : (3.5.1) holds, if e.g. S uniformly distributed and K, S are indep. (see Exc. 2)

2. method :

assumption : $M_{x+t} = M_{x+\frac{t}{2}}, 0 < t < 1$

\Rightarrow $\int_0^t M_{x+s} ds = \int_0^t M_{x+\frac{s}{2}} ds = \dots$

⑩ We now consider stochastic payment streams depending on the future lifetime T

4. Life insurance

An insurance policy is specified by

1. benefits payable by the insurer
2. premiums payable by the insured

Assumption: benefits and premiums may depend on the r.v. T

$Z :=$ PV of the benefits

$V :=$ PV of the premium payments

$\Rightarrow Z, V$ r.v.'s

$\rightarrow L := Z - V$ total loss of the insurer

\rightarrow reasonable to require that

$$E[L] = 0$$

$$\Leftrightarrow E[V] = E[Z]$$

$\rightarrow E[Z]$ net single premium

Remark 4.1: $E[Z]$ is not a measure for

\rightarrow the risk of the insurer
further characteristics of the distribution of Z (e.g. tails of distributions) are required for the assessment of risk

Examples of insurance policies based on T :

4.1 Elementary insurance types

4.1.1 Whole life and term insurance

1. whole life insurance (Enkel livsforsikring):

$$Z := v^{K+1} \quad \text{at } t = K+1$$

$$K = [T]$$

$$v = \frac{1}{1+i}$$

\rightarrow payment 1 at the end of the year of death

⑦ → net single premium :

$$A_x := E[Z] = E[v^{K+1}]$$

$$= \sum_{j \geq 0} v^{j+1} P_r(K=j)$$

$$\stackrel{(3.4.1)}{=} \sum_{j \geq 0} v^{j+1} {}_j p_x q_{x+j}$$

→ "risk measure" $\text{Var}[Z]$:

$$\text{Var}[Z] = E[Z^2] - A_x^2$$

$$\Rightarrow v = e^{-\delta}, \delta \text{ force of interest}$$

$$E[Z^2] = E[e^{-2\delta(K+1)}] =$$

$$\sum_{j \geq 0} e^{-2\delta(j+1)} {}_j p_x q_{x+j}$$

2. term insurance (risiko- eller korttidsforsikring)

$$Z = v^{K+1} \cdot 1_{\{K < n\}} \text{ at } t = K+1$$

→ payment of 1, if year of death $< n$ duration of the contract otherwise zero

→ net sing. prem :

$$A_{x:\overline{n}|} := E[Z] = \sum_{j=0}^{n-1} v^{j+1} {}_j p_x q_{x+j}$$

4.1.2 Pure endowments

pure endowment of duration n

$$Z = v^n \cdot 1_{\{K \geq n\}} \text{ at } t = n$$

→ paym. of 1, if the insured is alive at the end of n years else zero

→ net sing. prem. :

$$A_{x:\overline{n}|} := E[Z] = v^n {}_n p_x$$

4.1.3 Endowments (sammersatt (livsforsikring))

→ mixed case

→ Ex. 4.1.3.1 :

$$Z = \underbrace{v^{K+1} \cdot 1_{\{K < n\}}}_{\text{Term insur.}} + \underbrace{v^n \cdot 1_{\{K \geq n\}}}_{\text{Pure endowment}} \text{ at } t = K+1, n$$

→ net single prem. :

$$A_{x:\overline{n}|} := E[Z] = E[Z_1] + E[Z_2]$$

$$= A_{x:\overline{n}|} + A_{x:\overline{n}|}$$

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4.2 Payment at the moment of death

Consider e.g.

$$Z = v^T \quad \text{at } t=T$$

→ paym. of 1 at $t=T$, i.e. at the instant of death

$$\text{→ net sing. prem. : } \bar{A}_x = E[Z] = \int_0^\infty v^t \overbrace{g(t)}^{= eP_x \mu_{x+t}} dt$$

4.3 General types of life insurance

We consider benefits depending on time:

1. C_j payments at j -th year:

$$Z = C_{K+1} v^{K+1}$$

$$\Rightarrow E[Z] = \sum_{j \geq 0} C_{j+1} v^{j+1} {}_j p_x \cdot q_{x+j}$$

2. $c(t)$ cont. paym.:

$$Z = c(T) v^T$$

$$\Rightarrow E[Z] = \sum_{j \geq 0} E[Z | K=j] \cdot Pr(K=j)$$

$$\stackrel{T=K+S}{=} \sum_{j \geq 0} E[c(j+S) v^{j+S} | K=j] \cdot Pr(K=j) = \sum_{j \geq 0} {}_j p_x \cdot q_{x+j} \quad (\text{see Exerc. 2, Prob. 8})$$

4.4 Equations for A_x and \bar{A}_x

net sing. pr. of a whole life ins.

$$1. A_x : A_x = E[v^{K+1}] = \sum_{j \geq 0} v^{j+1} {}_j p_x q_{x+j}$$

$${}_j p_x = p_x \cdot {}_{j-1} p_{x+1}$$

$$\begin{aligned} \Rightarrow A_x &= v^{0+1} \underbrace{p_x}_{=1} \cdot q_{x+0} + \sum_{j \geq 1} \underbrace{v^{j+1} {}_j p_x}_{= v v^{j-1} p_x \cdot {}_{j-1} p_{x+1}} q_{x+j} \\ &= v \cdot q_x + v \cdot p_x \sum_{j \geq 0} \underbrace{v^j {}_j p_{x+1} q_{x+1+j}}_{= A_{x+1}} = v q_x + v p_x A_{x+1} \end{aligned}$$

$$\Rightarrow A_x = v q_x + v A_{x+1} p_x \quad (4.4.1)$$

where $Y = \begin{cases} v^{aT} & \text{if } T < 1 \\ v A_{x+1} & \text{if } T \geq 1 \end{cases}$

$$\xrightarrow{(4.4.1)} A_x = v A_{x+1} + v (1 - A_{x+1}) q_x = \underbrace{v A_{x+1}}_{\text{reserve}} + \underbrace{A_x \pi}_{\text{net sing. prem. of a 1-year term ins. with paym. } (1 - A_{x+1})} \cdot (1 - A_{x+1})$$

(19) 2. \bar{A}_x : $\bar{A}_x = E[CV^T] = \overbrace{A_{x+h}}^{\text{net sing. prem. of the same person aged } x+h}$

$$\Rightarrow \frac{\bar{A}_{x+h} - \bar{A}_x}{h} = \frac{(1 - v^h \cdot h p_x)}{h} \bar{A}_{x+h} - \frac{E[CV^T | T \leq h] \cdot h q_x}{h}$$

$h \rightarrow 0 \Rightarrow \delta \bar{A}_x = \frac{d}{dx} \bar{A}_x + \mu_x (1 - \bar{A}_x)$ (4.4.2) $\xrightarrow{h \rightarrow 0} g^{(0)} = \mu_x$

5. Life annuities

life annuities \rightarrow e.g. annuity-due, immediate annuity depending on the future life time T

Examples \rightarrow

5.1 Elementary life annuities

Ex 5.1.1 : whole life annuity-due (straks begynnende annuitet ar varighet $K+1$):

$$Y = \sum_{j=0}^K v^j = \ddot{a}_{\overline{K+1}|}$$

\rightarrow PV of annual paym. of 1 until $n=K$ (starting in zero)

\rightarrow net sing. prem.

$$\ddot{a}_x := E[Y] = \sum_{j \geq 0} \ddot{a}_{\overline{j+1}|} \overbrace{P_r(K=j)}^{= j p_x \cdot q_{x+j}}$$

Rewrite Y as

$$Y = \sum_{j \geq 0} v^j \cdot \underbrace{1_{\{K \geq j\}}}_{\text{PV of a pure endowment}}$$

$E[\cdot]$

$$\Rightarrow \ddot{a}_x = \sum_{j \geq 0} A_x \frac{1}{v^j}$$

Ex 5.1.2 : n-year temporary life annuity-due

$$Y = \ddot{a}_{\overline{K+1}|} \cdot 1_{\{K < n\}} + \ddot{a}_{\overline{n}|} \cdot 1_{\{K \geq n\}}$$

\rightarrow PV of an whole life annuity-due, if $K < n$ duration or of an annuity-due, if $K \geq n$

\rightarrow net sing. prem.

$$\ddot{a}_{x:\overline{n}|} := E[Y] = \sum_{j=0}^{n-1} \ddot{a}_{\overline{j+1}|} p_x q_{x+j} + \ddot{a}_{\overline{n}|} \cdot n p_x$$

Ex 5.1.3 : immediate life annuity :

$$Y = \sum_{j=1}^K v^j = a_{\overline{K}|}$$

\rightarrow PV of paym. of 1 from $t=1$ to $t=K$

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5.2 Life annuity with m payments per year

$$Y = \ddot{a}_{\overline{K+S^{(m)}}}^{(m)}$$

→ PV of paym. $\frac{1}{m}$ m times per year from $t=0$ to $t = K + S^{(m)} - \frac{1}{m}$

$S^{(m)} = \frac{1}{m} [mS+1]$
 m -th fraction of the year in which the insured dies

→ net sing. prem.

$$\ddot{a}_x^{(m)} = E[Y]$$

Rem. 5.2.1: If K, S indep., S unif. distr. on $[0,1]$ then

$$\ddot{a}_x^{(m)} = \alpha(m) \ddot{a}_x - \beta(m)$$

where $\alpha(m) \approx 1$ and $\beta(m) \approx \frac{m-1}{2m}$ (see Exerc. 2)

5.3 Variable life annuities

$$Y = \sum_{j=0}^K \tau_j v^j = \sum_{j \geq 0} \tau_j v^j \quad \{K \geq j\}$$

→ PV of paym. of τ_1, \dots, τ_K

→ net sing. prem.

$$E[Y] = \sum_{j \geq 0} \tau_j v^j j p_x$$

(Consider the continuous-time case:

$r(t) \quad t \geq 0$ paym. rate at time t

→ $Y = \int_0^T v^t r(t) dt$

→ net sing. prem.

$$E[Y] = \int_0^\infty v^t r(t) {}_t p_x dt \quad (\text{see Exerc. 3})$$

5.4 Equations for \ddot{a}_x, \bar{a}_x

1. \ddot{a}_x : From

$$\ddot{a}_x = \sum_{j \geq 0} v^j j p_x$$

and $j p_x = p_x \cdot j-1 p_{x+1}$

it follows

$$\ddot{a}_x = 1 + v \ddot{a}_{x+1} p_x \quad (5.4.1)$$

$$\Leftrightarrow \ddot{a}_x = \underbrace{1 + v \ddot{a}_{x+1}}_{\text{bond}} - \underbrace{v \ddot{a}_{x+1} q_x}_{\text{expected mortality gain}}$$