

STK4011 and STK9011

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Multivariate distributions

Covers (most of) the following material from chapter 4:

- Section 4.6: pages 177-178 and 182-186

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We consider a continuous multivariate random vector $\mathbf{X} = (X_1, \dots, X_n)$

For any set $A \subset \mathcal{R}^n$ we have that

$$P(\mathbf{X} \in A) = \int \dots \int_A f(\mathbf{x}) d\mathbf{x} = \int \dots \int_A f(x_1, \dots, x_n) dx_1 \dots dx_n$$

where $\mathbf{x} = (x_1, \dots, x_n)$ and $f(\mathbf{x}) = f(x_1, \dots, x_n)$ is the **joint pdf**

The **marginal pdf** of (X_1, \dots, X_k) is given by

$$f(x_1, \dots, x_k) = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f(x_1, \dots, x_n) dx_{k+1} \dots dx_n$$

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Let $\mathbf{X} = (X_1, \dots, X_n)$ be a multivariate random vector with joint pdf $f(\mathbf{x}) = f(x_1, \dots, x_n)$ and let $f_{X_i}(x_i)$ denote the marginal pdf of X_i ; $i = 1, \dots, n$

Then X_1, \dots, X_n are **mutually independent** random variables if for $(x_1, \dots, x_n) \in \mathcal{R}^n$ we may write

$$f(x_1, \dots, x_n) = \prod_{i=1}^n f_{X_i}(x_i)$$

Theorem 4.6.11

X_1, \dots, X_n are **mutually independent** if and only if there exist functions $g_i(x_i)$, $i = 1, \dots, n$, such that for all $(x_1, \dots, x_n) \in \mathcal{R}^n$ we may write

$$f(x_1, \dots, x_n) = \prod_{i=1}^n g_i(x_i)$$

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Let $\mathbf{X} = (X_1, \dots, X_n)$ be a multivariate random vector with joint pdf $f(\mathbf{x}) = f(x_1, \dots, x_n)$ and let $g(\mathbf{x}) = g(x_1, \dots, x_n)$ be a real valued function

Then the **expected value** of $g(\mathbf{X})$ is given by

$$\begin{aligned} E g(\mathbf{X}) &= \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} g(\mathbf{x}) f(\mathbf{x}) d\mathbf{x} \\ &= \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} g(x_1, \dots, x_n) f(x_1, \dots, x_n) dx_1 \dots dx_n \end{aligned}$$

If X_1, \dots, X_n are mutually independent random variables and $g_i(x_i)$, $i = 1, \dots, n$, are real valued functions such that $g_i(x_i)$ is a function only of x_i , then

$$E \left(\prod_{i=1}^n g_i(X_i) \right) = \prod_{i=1}^n E g_i(X_i)$$

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Let X_1, \dots, X_n be mutually independent random variables with mgfs $M_{X_i}(t)$, $i = 1, \dots, n$, and let a_1, \dots, a_n and b_1, \dots, b_n be constants. Then the mgf of $Z = \sum_{i=1}^n (a_i X_i + b_i)$ is

$$M_Z(t) = \exp\left(t \sum b_i\right) \prod_{i=1}^n M_{X_i}(a_i t)$$

This result may be used to prove that if X_1, \dots, X_n are independent and $X_i \sim n(\mu_i, \sigma_i^2)$, then

$$Z = \sum_{i=1}^n (a_i X_i + b_i) \sim n\left(\sum a_i \mu_i + b_i, \sum a_i^2 \sigma_i^2\right)$$

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Multivariate transformations

Let $\mathbf{X} = (X_1, \dots, X_n)$ be a multivariate random vector with joint pdf $f_{\mathbf{X}}(\mathbf{x}) = f_{\mathbf{X}}(x_1, \dots, x_n)$ and support

$$\mathcal{A} = \{\mathbf{x} : f_{\mathbf{X}}(\mathbf{x}) > 0\}$$

Consider a new random vector $\mathbf{U} = (U_1, \dots, U_n)$ defined by $U_i = g_i(\mathbf{X}) = g_i(X_1, \dots, X_n)$, $i = 1, \dots, n$

The transformation $(u_1, \dots, u_n) = (g_1(\mathbf{x}), \dots, g_n(\mathbf{x}))$ is a one-to-one transformation from \mathcal{A} onto

$$\mathcal{B} = \{\mathbf{u} : \mathbf{u} = (u_1, \dots, u_n) = (g_1(\mathbf{x}), \dots, g_n(\mathbf{x})) \text{ for } \mathbf{x} \in \mathcal{A}\}$$

The inverse transformation is given by

$$x_i = h_i(\mathbf{u}) = h_i(u_1, \dots, u_n), \quad i = 1, \dots, n$$

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Then the joint pdf of $\mathbf{U} = (U_1, \dots, U_n)$ is given by (for $\mathbf{u} \in \mathcal{B}$):

$$f_{\mathbf{U}}(u_1, \dots, u_n) = f_{\mathbf{X}}(h_1(u_1, \dots, u_n), \dots, h_n(u_1, \dots, u_n)) |J(u_1, \dots, u_n)|$$

Here $J(u_1, \dots, u_n)$ is the Jacobian:

$$J(u_1, \dots, u_n) = \begin{vmatrix} \frac{\partial x_1}{\partial u_1} & \frac{\partial x_1}{\partial u_2} & \dots & \frac{\partial x_1}{\partial u_n} \\ \frac{\partial x_2}{\partial u_1} & \frac{\partial x_2}{\partial u_2} & \dots & \frac{\partial x_2}{\partial u_n} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial x_n}{\partial u_1} & \frac{\partial x_n}{\partial u_2} & \dots & \frac{\partial x_n}{\partial u_n} \end{vmatrix}$$

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Example 4.6.13

We have the joint pdf

$$f_{\mathbf{X}}(x_1, x_2, x_3, x_4) = \begin{cases} 24e^{-x_1 - x_2 - x_3 - x_4} & 0 < x_1 < x_2 < x_3 < x_4 < \infty \\ 0 & \text{otherwise} \end{cases}$$

Consider the transformation

$$\begin{aligned} U_1 &= X_1 & U_2 &= X_2 - X_1 \\ U_3 &= X_3 - X_2 & U_4 &= X_4 - X_3 \end{aligned}$$

The support of the distribution of the U_i s is

$$\mathcal{B} = \{(u_1, u_2, u_3, u_4) : u_i > 0, i = 1, 2, 3, 4\}$$

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The inverse transformation is

$$\begin{aligned} X_1 &= U_1 & X_2 &= U_1 + U_2 \\ X_3 &= U_1 + U_2 + U_3 & X_4 &= U_1 + U_2 + U_3 + U_4 \end{aligned}$$

The Jacobian becomes

$$J(u_1, u_2, u_3, u_4) = \begin{vmatrix} \frac{\partial x_1}{\partial u_1} & \frac{\partial x_1}{\partial u_2} & \frac{\partial x_1}{\partial u_3} & \frac{\partial x_1}{\partial u_4} \\ \frac{\partial x_2}{\partial u_1} & \frac{\partial x_2}{\partial u_2} & \frac{\partial x_2}{\partial u_3} & \frac{\partial x_2}{\partial u_4} \\ \frac{\partial x_3}{\partial u_1} & \frac{\partial x_3}{\partial u_2} & \frac{\partial x_3}{\partial u_3} & \frac{\partial x_3}{\partial u_4} \\ \frac{\partial x_4}{\partial u_1} & \frac{\partial x_4}{\partial u_2} & \frac{\partial x_4}{\partial u_3} & \frac{\partial x_4}{\partial u_4} \end{vmatrix} = \begin{vmatrix} 1 & 0 & 0 & 0 \\ 1 & 1 & 0 & 0 \\ 1 & 1 & 1 & 0 \\ 1 & 1 & 1 & 1 \end{vmatrix} = 1$$

The joint pdf of the U_i s becomes (when all $u_i > 0$)

$$\begin{aligned} f_{\mathbf{U}}(u_1, u_2, u_3, u_4) &= 24e^{-u_1 - (u_1+u_2) - (u_1+u_2+u_3) - (u_1+u_2+u_3+u_4)} \\ &= 24e^{-4u_1 - 3u_2 - 2u_3 - u_4} \\ &= 4e^{-4u_1} 3e^{-3u_2} 2e^{-2u_3} e^{-u_4} \end{aligned}$$

The U_i s are independent and

$$U_i \sim \text{exponential}(1/(5-i)), \quad i = 1, 2, 3, 4$$