

Solutions to exercises - Week 36

Beta function:

- Additional exercise

Exponential family of distributions:

- Exercises 3.28b-d, 3.30b and 3.33b-c

Location/scale distributions:

- Exercises 3.38, and 3.39

Bivariate distributions:

- Exercises 4.4a-b and 4.5

Additional exercise

We have the gamma function

$$\Gamma(\alpha) = \int_0^{\infty} x^{\alpha-1} e^{-x} dx$$

and the beta function

$$B(\alpha, \beta) = \int_0^1 x^{\alpha-1} (1-x)^{\beta-1} dx$$

We will prove the relation

$$B(\alpha, \beta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha + \beta)}$$

We have to prove that

$$\Gamma(\alpha)\Gamma(\beta) = \Gamma(\alpha + \beta)B(\alpha, \beta)$$

Now we may write

$$\begin{aligned}\Gamma(\alpha)\Gamma(\beta) &= \int_0^{\infty} x^{\alpha-1} e^{-x} dx \int_0^{\infty} y^{\beta-1} e^{-y} dy \\ &= \int_0^{\infty} \int_0^{\infty} x^{\alpha-1} e^{-x} y^{\beta-1} e^{-y} dy dx \\ &= \int_0^{\infty} \int_0^{\infty} x^{\alpha-1} y^{\beta-1} e^{-(x+y)} dy dx\end{aligned}$$

We then perform a change of variables:

$$u = x + y \quad \text{and} \quad v = \frac{x}{x + y}$$

Note that $u > 0$ and $0 < v < 1$

This gives $x = uv$ and $y = u(1 - v)$

The Jacobian becomes

$$J(u, v) = \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix} = \begin{vmatrix} v & u \\ 1 - v & -u \end{vmatrix} = -vu - (1 - v)u = -u$$

We then obtain

$$\begin{aligned} \Gamma(\alpha)\Gamma(\beta) &= \int_0^{\infty} \int_0^1 (uv)^{\alpha-1} [u(1-v)]^{\beta-1} e^{-u} | -u | dv du \\ &= \int_0^{\infty} \int_0^1 u^{\alpha+\beta-1} e^{-u} v^{\alpha-1} (1-v)^{\beta-1} dv du \\ &= \int_0^{\infty} u^{\alpha+\beta-1} e^{-u} du \int_0^1 v^{\alpha-1} (1-v)^{\beta-1} dv = \Gamma(\alpha + \beta) B(\alpha, \beta) \end{aligned}$$

Exercise 3.28b

$$f(x | \alpha, \beta) = \begin{cases} \frac{1}{\beta^\alpha \Gamma(\alpha)} x^{\alpha-1} e^{-x/\beta} & \text{if } x > 0 \\ 0 & \text{otherwise} \end{cases}$$

For **all** x we may write:

$$f(x | \alpha, \beta) = \underbrace{I_{\{x>0\}}(x)}_{h(x)} \underbrace{\frac{1}{\beta^\alpha \Gamma(\alpha)}}_{c(\alpha, \beta)} \exp \left\{ \underbrace{(\alpha - 1)}_{w_1(\alpha, \beta)} \underbrace{\log x}_{t_1(x)} + \underbrace{\left(-\frac{1}{\beta}\right)}_{w_2(\alpha, \beta)} \underbrace{x}_{t_2(x)} \right\}$$

Exercise 3.28c

$$f(x | \alpha, \beta) = \begin{cases} \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1-x)^{\beta-1} & \text{if } 0 < x < 1 \\ 0 & \text{otherwise} \end{cases}$$

For **all** x we may write:

$$f(x | \alpha, \beta) = \underbrace{I_{(0,1)}(x)}_{h(x)} \underbrace{\frac{1}{B(\alpha, \beta)}}_{c(\alpha, \beta)} \exp \left(\underbrace{(\alpha-1)}_{w_1(\alpha, \beta)} \underbrace{\log x}_{t_1(x)} + \underbrace{(\beta-1)}_{w_2(\alpha, \beta)} \underbrace{\log(1-x)}_{t_2(x)} \right)$$

Exercise 3.28d

$$f(x|\lambda) = \begin{cases} \frac{\lambda^x}{x!} e^{-\lambda} & \text{for } x = 0, 1, 2, \dots \\ 0 & \text{otherwise} \end{cases}$$

For **all** x we may write:

$$f(x|\lambda) = \underbrace{I_{\{0,1,2,\dots\}}(x)}_{h(x)} \underbrace{\frac{1}{x!}}_{c(\lambda)} \exp \underbrace{\log(\lambda)}_{w_1(\lambda)} \underbrace{x}_{t_1(x)}$$

Exercise 3.30b

From exercise 3.28.d we have:

$$f(x|\lambda) = \underbrace{I_{\{0,1,2,\dots\}}(x)}_{h(x)} \underbrace{\frac{1}{x!} e^{-\lambda}}_{c(\lambda)} \exp \underbrace{\log(\lambda)}_{w_1(\lambda)} \underbrace{x}_{t_1(x)}$$

The relations in Theorem 3.4.2 become

$$\mathbf{E} \left(\frac{1}{\lambda} X \right) = 1 \qquad \mathbf{Var} \left(\frac{1}{\lambda} X \right) = 0 - \mathbf{E} \left(-\frac{1}{\lambda^2} X \right)$$

From these relations we obtain

$$\mathbf{E} X = \lambda$$

$$\mathbf{Var} X = \lambda^2 \left\{ 0 + \frac{1}{\lambda^2} \lambda \right\} = \lambda$$

Exercise 3.33.b

Assume $X \sim n(\theta, a\theta^2)$

The pdf takes the form

$$f(x | \theta) = \frac{1}{\sqrt{2\pi} \sqrt{a\theta^2}} \exp\left(-\frac{(x - \theta)^2}{2a\theta^2}\right) \quad a > 0 \text{ known}$$

This may be written

$$f(x | \theta) = \underbrace{\frac{1}{\sqrt{2\pi} \sqrt{a\theta^2}} \exp\left(-\frac{1}{2a}\right)}_{c(\theta)} \exp\left(\underbrace{\frac{1}{2a\theta^2}}_{w_1(\theta)} \underbrace{-x^2}_{t_1(x)} + \underbrace{\frac{1}{a\theta}}_{w_2(\theta)} \underbrace{x}_{t_2(x)}\right)$$

For the normal distribution the full parameter space is

$$(\mu, \sigma^2) : -\infty < \mu < \infty, \sigma^2 > 0$$

Here we have a curved exponential distribution with parameter space

$$(\mu, \sigma^2) : \mu = \theta, \sigma^2 = a\theta^2, -\infty < \theta < \infty$$

Exercise 3.33.c

Assume $X \sim \text{gamma}(\alpha, 1/\alpha)$

The pdf takes the form

$$f(x|\alpha) = \begin{cases} \frac{\alpha^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\alpha x} & \text{if } x > 0 \\ 0 & \text{otherwise} \end{cases}$$

For all x we may write:

$$f(x|\alpha) = \underbrace{I_{\{x>0\}}(x)}_{h(x)} \underbrace{\frac{\alpha^\alpha}{\Gamma(\alpha)}}_{c(\alpha)} \exp \left(\underbrace{(\alpha-1)}_{w_1(\alpha)} \underbrace{\log x}_{t_1(x)} + \underbrace{\alpha}_{w_2(\alpha)} \underbrace{(-x)}_{t_2(x)} \right)$$

For the gamma distribution the full parameter space is

$$(\alpha, \beta) : \alpha > 0, \beta > 0$$

Here we have a curved exponential distribution with parameter space

$$(\alpha, \beta) : \alpha > 0, \beta = 1 / \alpha$$

Exercise 3.38

$$Z \sim f(z) \quad \alpha = P(Z > z_\alpha) = \int_{z_\alpha}^{\infty} f(z) dz$$

$$X \text{ has pdf } f_X(x) = \frac{1}{\sigma} f\left(\frac{x - \mu}{\sigma}\right)$$

Then

$$\alpha = P(X > x_\alpha) = \int_{x_\alpha}^{\infty} f_X(x) dx = \int_{x_\alpha}^{\infty} \frac{1}{\sigma} f\left(\frac{x - \mu}{\sigma}\right) dx = \int_{\frac{x_\alpha - \mu}{\sigma}}^{\infty} f(z) dz$$

Therefore

$$\frac{x_\alpha - \mu}{\sigma} = z_\alpha \quad \text{and} \quad x_\alpha = \sigma z_\alpha + \mu$$

Exercise 3.39

$$f_X(x) = \frac{1}{\sigma\pi} \frac{1}{1 + (x - \mu)/\sigma^2} \quad \text{for } -\infty < x < \infty$$

The pdf of $Z = (X - \mu)/\sigma$ is given by

$$f(z) = \frac{1}{\pi} \frac{1}{1 + z^2} \quad \text{for } -\infty < z < \infty$$

Recall that $\frac{d}{du} \arctan(u) = \frac{1}{1 + u^2}$

The cdf of Z becomes

$$F(z) = \int_{-\infty}^z \frac{1}{\pi} \frac{1}{1 + u^2} du = \frac{1}{\pi} \arctan(u) \Big|_{-\infty}^z = \frac{1}{\pi} \arctan(z) + \frac{1}{2}$$

We have

$$P(Z \leq 0) = F(0) = \frac{1}{\pi} \arctan(0) + \frac{1}{2} = 0 + \frac{1}{2} = \frac{1}{2}$$

Further we have

$$P(Z \leq -1) = F(-1) = \frac{1}{\pi} \arctan(-1) + \frac{1}{2} = \frac{1}{\pi} \left(-\frac{\pi}{4} \right) + \frac{1}{2} = \frac{1}{4}$$

$$P(Z \geq 1) = 1 - F(1) = \frac{1}{2} - \frac{1}{\pi} \arctan(1) = \frac{1}{2} - \frac{1}{\pi} \frac{\pi}{4} = \frac{1}{4}$$

Finally for $X = \sigma Z + \mu$ we have

$$P(X \leq x) = P(Z \leq (x - \mu) / \sigma) = F((x - \mu) / \sigma)$$

$$P(X \leq \mu) = F(0) = 1/2$$

$$P(X \leq \mu - \sigma) = F(-1) = 1/4$$

$$P(X \geq \mu + \sigma) = 1 - P(X \leq \mu + \sigma) = 1 - F(1) = 1/4$$

Exercise 4.4.a

$$f(x, y) = \begin{cases} C(x + 2y) & 0 < x < 2, 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$

We have:

$$\begin{aligned} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy &= \int_0^1 \int_0^2 C(x + 2y) dx dy = C \int_0^1 \left[\frac{x^2}{2} + 2yx \right]_0^2 dy \\ &= C \int_0^1 2 + 4y dy = C \left[2y + 2y^2 \right]_0^1 = 4C \end{aligned}$$

From this we obtain $C = \frac{1}{4}$

Exercise 4.4.b

$$f(x, y) = \begin{cases} \frac{1}{4}(x + 2y) & 0 < x < 2, 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$

Marginal distribution of X (for $0 < x < 2$)

$$\begin{aligned} f_X(x) &= \int_{-\infty}^{\infty} f(x, y) dy = \int_0^1 \frac{1}{4}(x + 2y) dy \\ &= \left[\frac{1}{4}(xy + y^2) \right]_0^1 = \frac{1}{4}(x + 1) \end{aligned}$$

Exercise 4.5.a

$$f(x, y) = \begin{cases} x + y & 0 < x < 1, 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$

We have

$$P(X > \sqrt{Y}) = P(Y < X^2) = \int_0^1 \int_0^{x^2} (x + y) dy dx$$

$$= \int_0^1 \left[xy + \frac{y^2}{2} \right]_0^{x^2} dx = \int_0^1 \left(x^3 + \frac{x^4}{2} \right) dx$$

$$= \left[\frac{x^4}{4} + \frac{x^5}{10} \right]_0^1 = \frac{7}{20}$$

Exercise 4.5.b

$$f(x, y) = \begin{cases} 2x & 0 < x < 1, 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$

We have

$$\begin{aligned} P(X^2 < Y < X) &= \int_0^1 \int_{x^2}^x 2x \, dy \, dx = \int_0^1 2xy \Big|_{x^2}^x \, dx \\ &= \int_0^1 2x^2 - 2x^3 \, dx = \left[\frac{2x^3}{3} - \frac{2x^4}{4} \right]_0^1 \\ &= \frac{1}{6} \end{aligned}$$