



**UiO • Matematisk institutt**

Det matematisk-naturvitenskapelige fakultet

**STK-4051/9051 Computational Statistics Spring 2021**  
**Chapter 6**

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# Problem with presented methods in high dimensions, Rejection sampling

- Assume you want to sample iid  $X_i, i = 1, \dots, n$
- Two methods with rejection sampling  $\frac{f(x_i)}{g(x_i)} \leq \alpha^{-1}$ 
  1. Sample  $X_i, i = 1, \dots, n$  independently.
  2. Sample  $\{X_1, X_2, \dots, X_n\}$  simultaneously
- Expected time to sample complete
  1.  $n \cdot \alpha^{-1}$  (nice 😊)
  2.  $1 \cdot \alpha^{-n}$  (curse of dimensionality is back 😞)
- For complex distributions in high dimensions we need a joint proposal (within this framework)

# Factoring into 1D distributions

- Target

$$f(\mathbf{x}) = f(x_1)f(x_2|x_1) \cdots f(x_n|x_1, x_2, \dots, x_{n-1})$$

- Proposal

$$g(\mathbf{x}) = g(x_1)g(x_2|x_1) \cdots g(x_n|x_1, x_2, \dots, x_{n-1})$$

- Sample each component sequentially as 1D
- Markov property: (Simpler notation)
- $f(\mathbf{x}) = f(x_1)f(x_2|x_1)f(x_3|x_2) \cdots f(x_n|x_{n-1})$
- $g(\mathbf{x}) = g(x_1)g(x_2|x_1)g(x_3|x_2) \cdots g(x_n|x_{n-1})$

# Sequential Monte Carlo in a Markov structure

- Sequential Monte Carlo: First **high-dimensional setting**
- Assume now  $\mathbf{x} = \mathbf{x}_{1:t} = (x_1, \dots, x_t)$  have a **Markov** structure

$$f_t(\mathbf{x}_{1:t}) = f_1(x_1) \prod_{i=2}^t f_i(x_i | x_{i-1})$$

- Also assume a **proposal** distribution with **Markov property**:

$$g_t(\mathbf{x}_{1:t}) = g_1(x_1) \prod_{i=2}^t g_i(x_i | x_{i-1})$$

- Importance weights:

$$w(\mathbf{x}_{1:t}) = \frac{f_t(\mathbf{x}_{1:t})}{g_t(\mathbf{x}_{1:t})} = \frac{f_1(x_1)}{g_1(x_1)} \prod_{i=2}^t \frac{f_i(x_i | x_{i-1})}{g_i(x_i | x_{i-1})} = w(\mathbf{x}_{1:t-1}) \frac{f_t(x_t | x_{t-1})}{g_t(x_t | x_{t-1})}$$

- Opens up for **sequential** sampling/estimation
- Note: Easy to generalize to non-Markov settings as well
  - More computing at each step

# Sequential Monte Carlo

## Algorithm

- 1 Sample  $X_1 \sim g_1(\cdot)$ . Let  $w_1 = u_1 = f_1(x_1)/g_1(x_1)$ . Set  $t = 2$
- 2 Sample  $X_t|x_{t-1} \sim g_t(x_t|x_{t-1})$ .
- 3 Append  $x_t$  to  $\mathbf{x}_{1:t-1}$ , obtaining  $\mathbf{x}_t$
- 4 Let  $u_t = f_t(x_t|x_{t-1})/g_t(x_t|x_{t-1})$
- 5 Let  $w_t = w_{t-1} u_t$ , the importance weight for  $\mathbf{x}_{1:t}$
- 6 Increment  $t$  and return to step 2

Can simulate  $m$  sequences **in paralell!**

# Sequential Monte Carlo Example

- Assume Markov model

$$X_1 \sim N(0, \sqrt{2})$$

$$f_t(x_t | x_{t-1}) \propto |\cos(x_t - x_{t-1})| \exp \left\{ -\frac{1}{4}(x_t - x_{t-1})^2 \right\}$$

- Of interest:

$$\mu_t = E[X_t] \qquad = 0 \text{ due to symmetry}$$

$$\sigma_t^2 = \text{var}(X_t) = E[X_t^2] - (E[X_t])^2$$

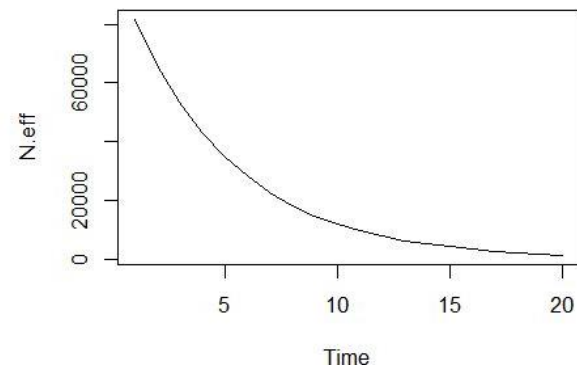
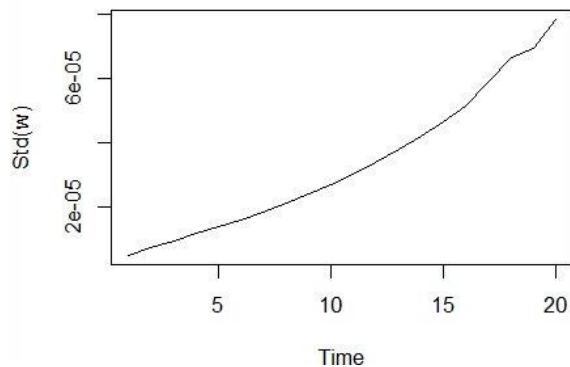
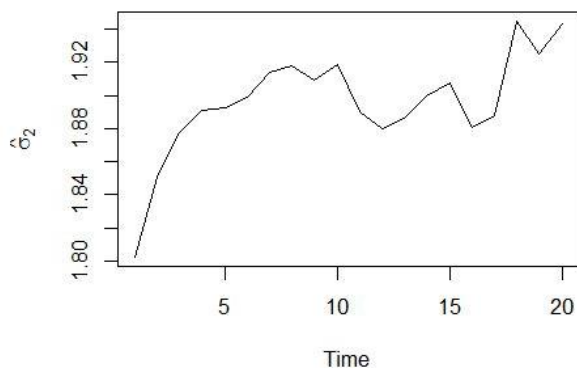
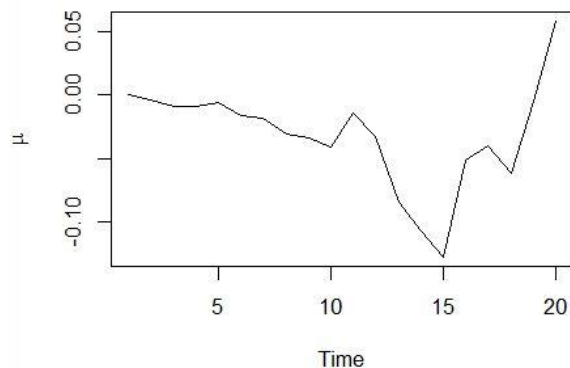
- Estimators by SMC simulation:

$$\hat{\mu}_t = \frac{\sum_{i=1}^n w_t^i x_t^i}{\sum_{i=1}^n w_t^i}$$

$$\hat{\sigma}_t^2 = \frac{\sum_{i=1}^n w_t^i (x_t^i)^2}{\sum_{i=1}^n w_t^i} - \left[ \frac{\sum_{i=1}^n w_t^i x_t^i}{\sum_{i=1}^n w_t^i} \right]^2 = \frac{\sum_{i=1}^n w_t^i (x_t^i - \hat{\mu}_t)^2}{\sum_{i=1}^n w_t^i}$$

- SMC\_cosnorm.R

# Dimension =20, m=100 000

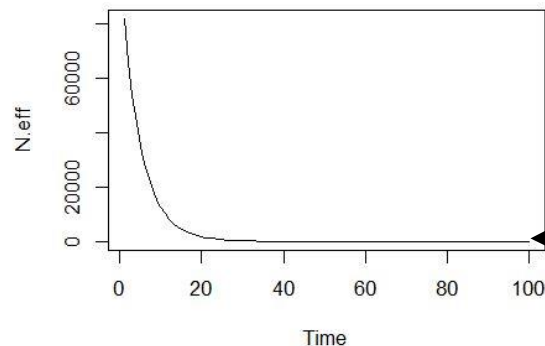
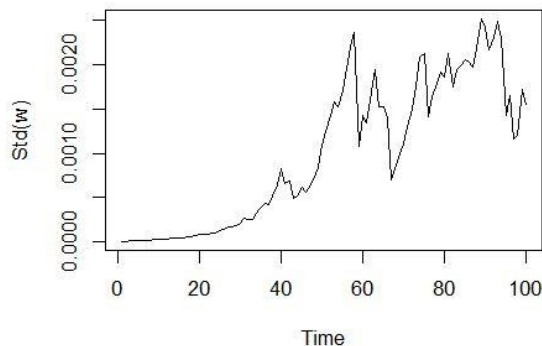
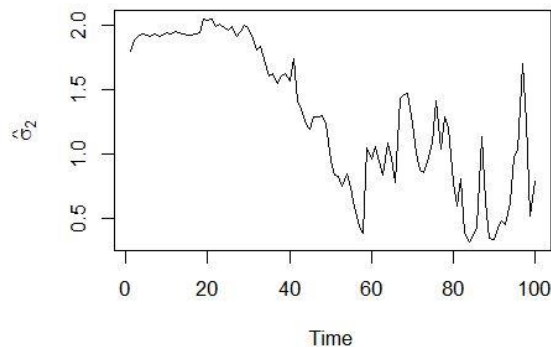
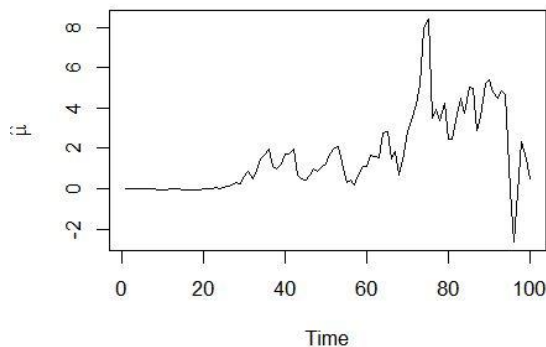


$$\hat{N}_{\text{eff}} = \frac{1}{\sum_{i=1}^n w_i^2}$$

Neff:  
1514

```
show(w.norm.sort[1:10])  
[1] 0.007377638 0.006993864 0.004586039 0.003991772 0.003901192 0.003686038 0.003401063 0.003339764 0.003094683 0.0029111
```

# Dimension =100, m=100 000



$$\hat{N}_{eff} = \frac{1}{\sum_{i=1}^n w_i^2}$$

N\_eff:  
3.78

```
show(w.norm.sort[1:10])  
[1] 0.47751939 0.06007770 0.04960048 0.03573471 0.03281825 0.02935132 0.02665531 0.02271621 0.02231759 0.02197812
```



# Weight degeneracy

- General rule:

$$\text{var}[Y] = E[\text{var}[Y|Z]] + \text{var}[E[Y|Z]] \geq \text{var}[E[Y|Z]]$$

- $Y = w_t, Z = \mathbf{X}_{1:t-1}$  ( $w_{t-1}$  given by  $\mathbf{x}_{1:t-1}$ ):

$$\begin{aligned} E[w_t | \mathbf{X}_{1:t-1}] &= w_{t-1} E\left[\frac{f_t(X_t | X_{t-1})}{g_t(X_t | X_{t-1})} \mid \mathbf{X}_{1:t-1}\right] \\ &= w_{t-1} \cdot 1 = w_{t-1} \end{aligned}$$

implying that

$$\text{var}[w_t] \geq \text{var}[w_{t-1}]$$

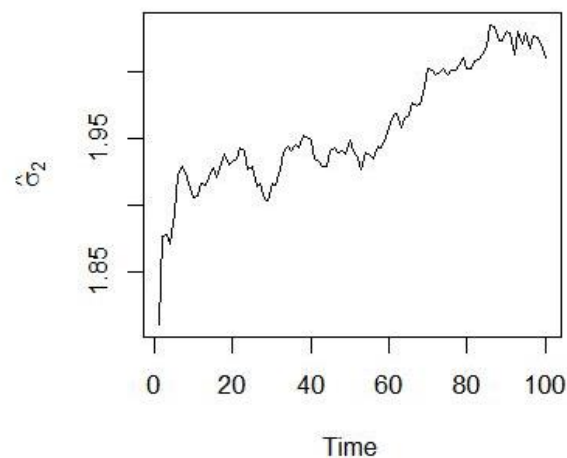
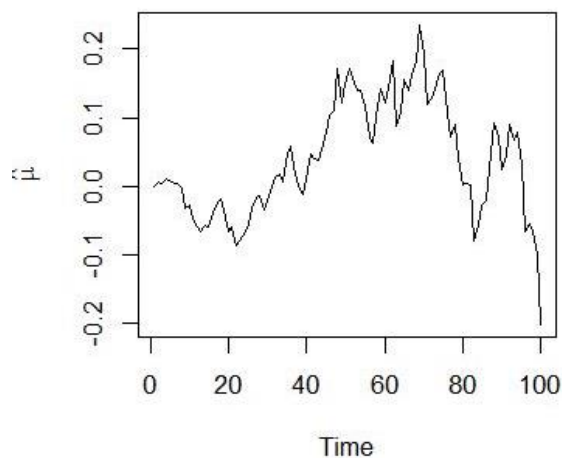
which indicates that the variance will **increase at each time-step**.

- **Practical consequence:**
  - Only a few samples will dominate the others
  - Variability of estimate will increase

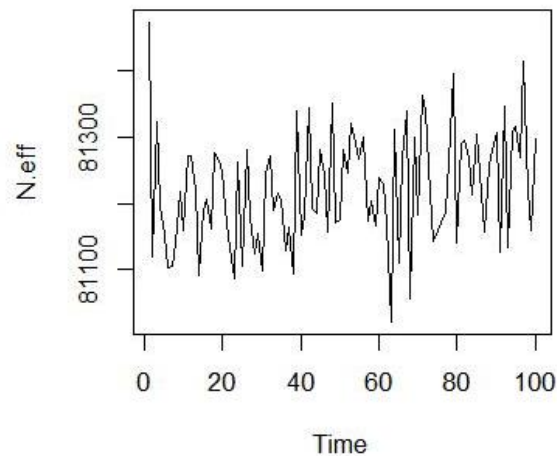
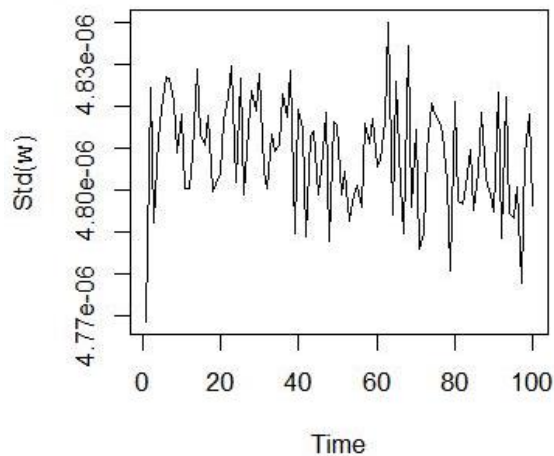
# Resampling

- **Degeneracy** of weights a serious problem.
- Solution: **Resampling** (SIR idea)
  - One possible choice for resampling: Apply SIR directly
  - Each iteration:
    - Sample  $\tilde{x}_t^i$ , independently from  $\{x_t^i\}$ , probability of each sample is  $w_t^i$
    - Set all new weights equal to  $1/N$
  - When  $N_{\text{eff}}$  is small
    - Sample  $\tilde{x}_t^i$ , independently from  $\{x_t^i\}$ , probability of each sample is  $w_t^i$
    - Set all new weights equal to  $1/N$
- `SMC_cosnorm.R`

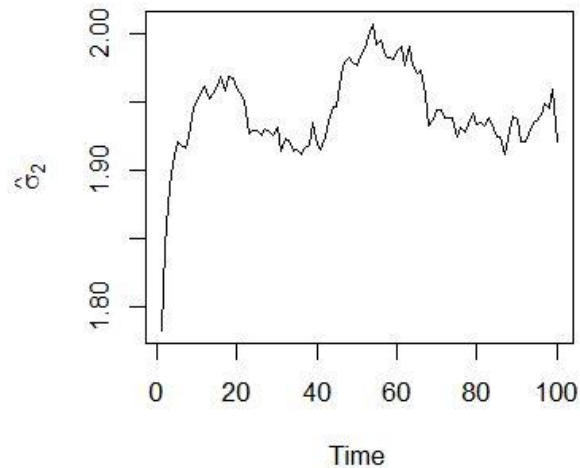
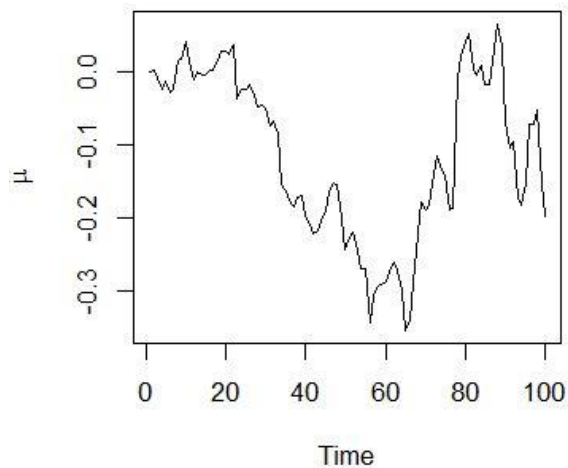
# Dimension =100, m=100 000



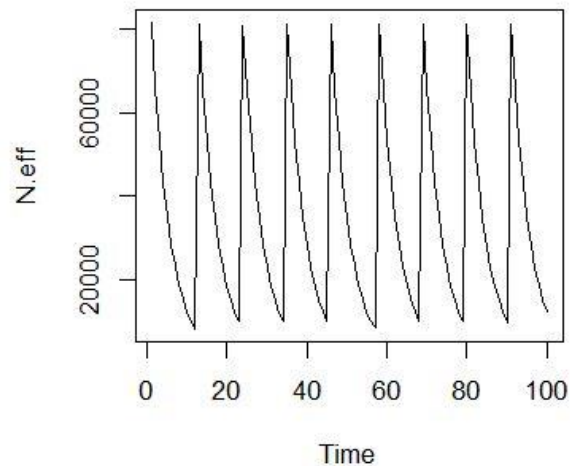
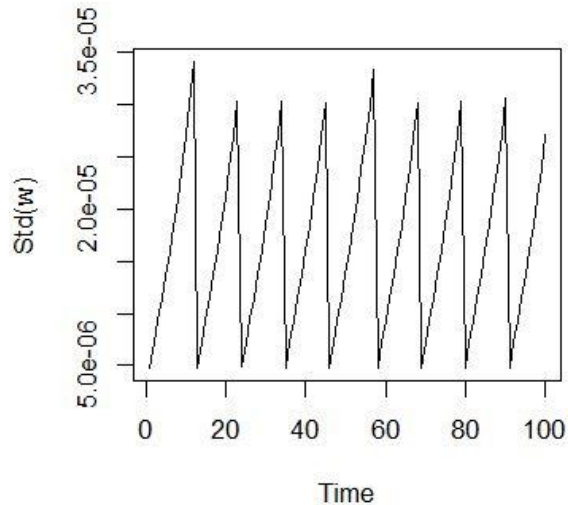
Resample  
each time



# Dimension =100, m=100 000



Resample  
 $n < 10\ 000$



# How to resample more generally?

- Resampling  $\tilde{x}_t^i$  from  $x_t^i$  :
  - Original (normalized) weights:  $w_t^i$
  - Resampling weights:  $\widetilde{w}_t^i$
  - The number of repeats for  $x_t^i$ :  $N_t^i$
- We need the expected number of resamples times the new weight to be the old weight
  - $E(N_t^i \cdot \widetilde{w}_t^i) = w_t^i$
- One choice (as above)
  - Sample  $\tilde{x}_t^i$ , independently from  $\{x_t^i\}$ , probability of each sample is  $w_t^i$
  - Set new weights  $1/N$

# Resampling

Original (normalized) weights:  $w_t^i$

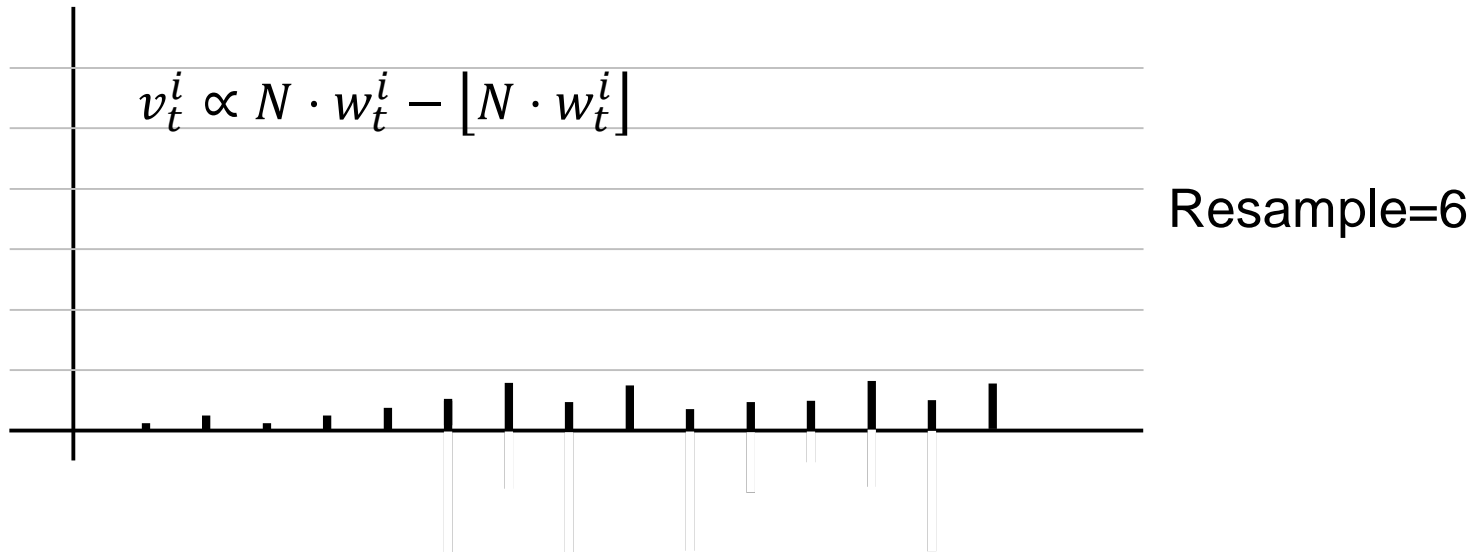
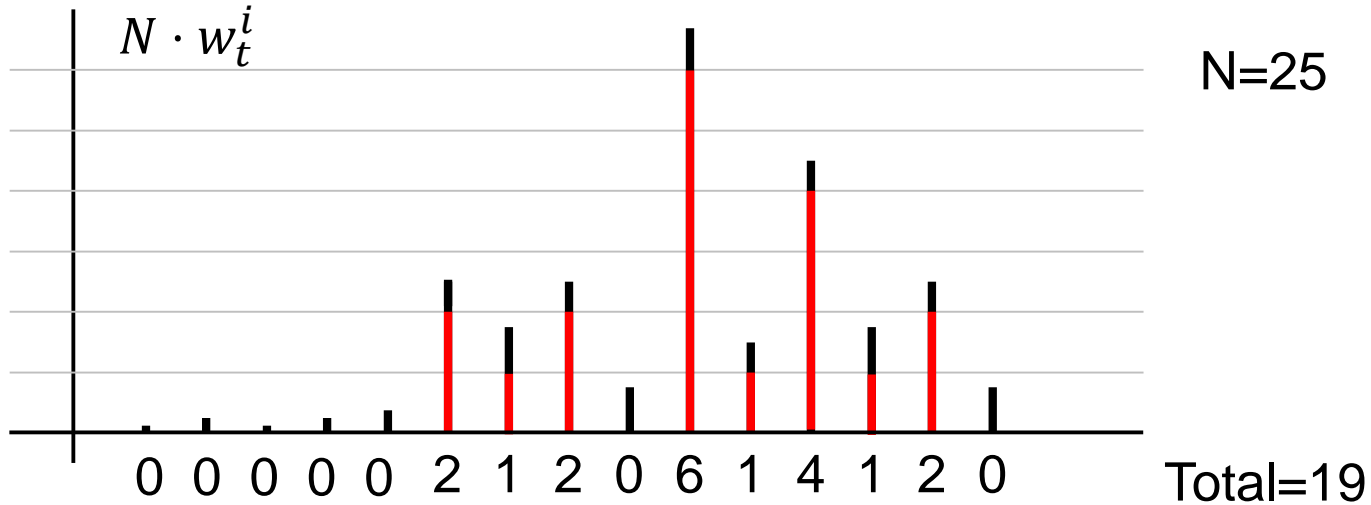
Resampling weights:  $\tilde{w}_t^i$

The number of repeats for  $x_t^i$ :  $N_t^i$

- Simplest option:
  - Resample with probabilities equal to  $w_t^i$ .
  - Put weights on resample to  $\tilde{w}_t^i = N^{-1}$
  - Number of repeats of  $x_t^i$ ,  $N_t^i$  is Binomial( $N$ ,  $w_t^i$ )
  - $E[N_t^i \tilde{w}_t^i] = Nw_t^i$
- More general resampling strategies are possible
- **Sufficient requirement:**  $E[N_t^i \tilde{w}_t^i] = Nw_t^i$
- Optimal strategy (for equally weighted samples)
  - For  $i = 1, \dots, N$ , put ( $\lfloor a \rfloor$  is the largest integer smaller than  $a$ )
 
$$\tilde{N}_t^i = \lfloor Nw_t^i \rfloor \quad (\text{Some will be zero})$$
  - Let  $\delta_t^i = w_t^i - \tilde{N}_t^i / N$
  - Define  $K = N - \sum_{i=1}^N \tilde{N}_t^i$  (remaining particles that have not been allocated)
  - Sample  $(D_t^1, \dots, D_t^K)$  from the multinomial distribution with probabilities proportional to  $(\delta_t^1, \dots, \delta_t^K)$ .
  - Put  $N_t^i = \tilde{N}_t^i + D_t^i$
  - Make  $N_t^i$  replicates of  $x_t^i$ , but all weights to  $1/N$

# Illustration of optimal resampling

(resample weight:  $1/N$ )



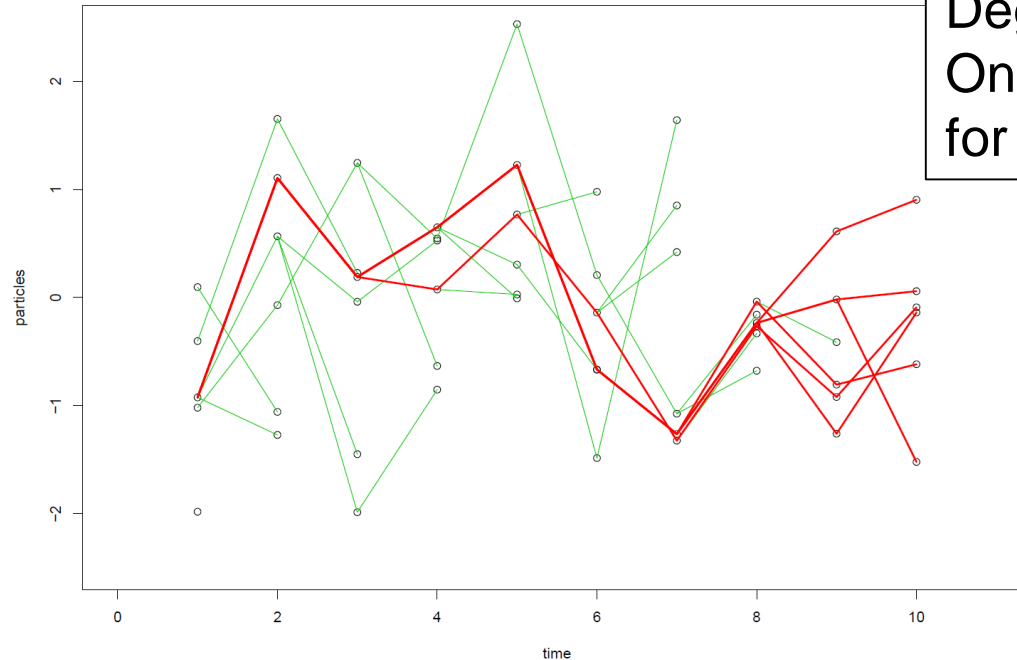
# Example optimal resampling

- $N = 5, \mathbf{w} = (0.3, 0.4, 0.05, 0.15, 0.2)$
- $N * \mathbf{w} = (1.5, 2.0, 0.25, 0.75, 1.0)$
- $[N \cdot w_t^i]$  •  $\tilde{\mathbf{N}} = (1, 2, 0, 0, 1)$
- $K = 5 - 4 = 1$
- $\tilde{\mathbf{N}}/N = (0.2, 0.4, 0.0, 0.0, 0.2)$
- $\delta = (0.1, 0.0, 0.05, 0.15, 0.0)$
- Sample  $\mathbf{D}$  from  $\text{Multinom}(1 : N, 1, (\frac{0.1}{0.3}, \frac{0.0}{0.3}, \frac{0.05}{0.3}, \frac{0.15}{0.3}, \frac{0.0}{0.3}))$   
e.g  $\mathbf{D} = (1, 0, 0, 0, 0)$
- Put  $\mathbf{N} = \tilde{\mathbf{N}} + \mathbf{D} = (2, 2, 0, 0, 1)$



# Resampling properties

- Resampling will introduce extra random noise at the **current** time-point
- Can reduce noise at **later** time points
- Gives a good approximation to  $f(x_n)$



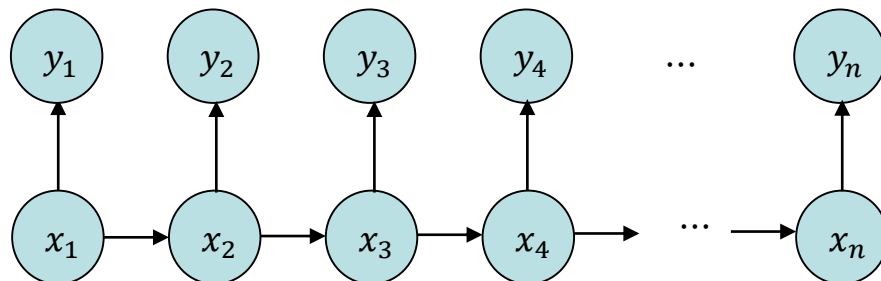
Degenerated sample path  
Only one particle left  
for  $t=1, 2$  and  $3$

- Does **not** give a good approximation to  $f(\mathbf{x})$  or  $f(x_1)$ !

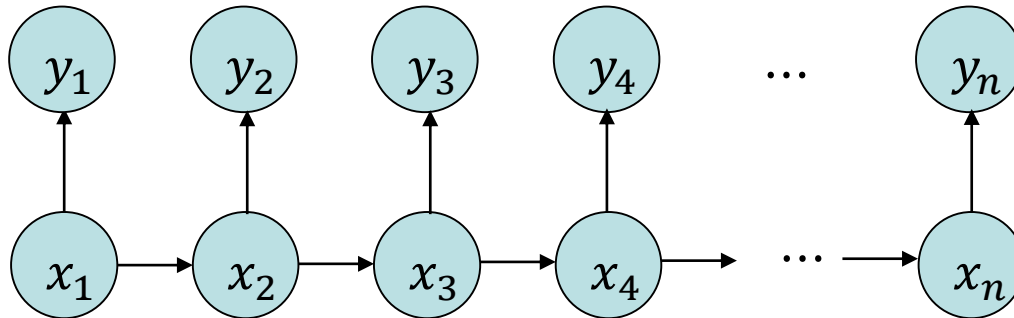
# Resampling properties

- Resampling will introduce extra random noise at the **current** time-point
- Can reduce noise at **later** time points
- Gives a good approximation to  $f(x_n)$
- Does **not** give a good approximation to  $f(\mathbf{x})$  or  $f(x_1)$ !

Makes it suited for filter problem  $p(\mathbf{x}_n | \mathbf{y}_{1:n})$



# Origin of method in state space models



- For the filtering problem the target is the distribution  $p(\mathbf{x}_n | \mathbf{y}_{1:n})$
- In chapter 4, we consider a hidden Markov model the distribution of  $x$  (state) is discrete here  $x$  (the state) it can be continuous

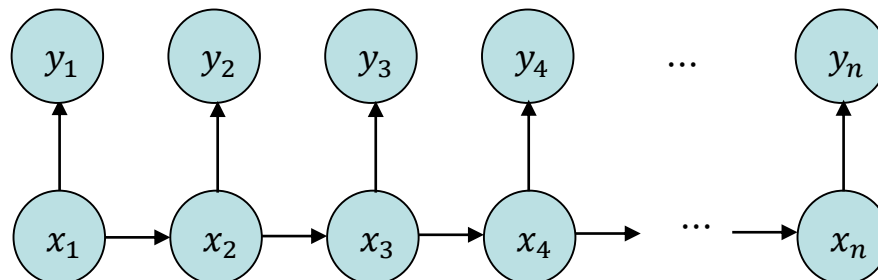
# Hidden Markov models - state space models

- Assume

$$X_1 \sim p_{x_1}(x_1)$$

$$X_t \sim p_x(x_t|x_{t-1})$$

$$Y_t \sim p_y(y_t|x_t)$$



- $\{y_t\}$  observed,  $\{x_t\}$  **hidden**
- Chapter 4:  $\{x_t\}$  discrete. Now possibly **continuous**
- Aim:**  $f(\mathbf{x}_{1:t}|\mathbf{y}_{1:t})$  or  $f(\mathbf{x}_t|\mathbf{y}_{1:t})$
- Recursive relationship (misprint in book):

$$\begin{aligned} f(\mathbf{x}_{1:t}|\mathbf{y}_{1:t}) &= \frac{f(\mathbf{x}_{1:t}, y_t|\mathbf{y}_{1:t-1})}{f(y_t|\mathbf{y}_{1:t-1})} \\ &= \frac{f(\mathbf{x}_{1:t-1}|\mathbf{y}_{1:t-1})p_x(x_t|x_{t-1})p_y(y_t|x_t)}{f(y_t|\mathbf{y}_{1:t-1})} \\ &\propto f(\mathbf{x}_{1:t-1}|\mathbf{y}_{1:t-1})p_x(x_t|x_{t-1})p_y(y_t|x_t) \end{aligned}$$

# Sequential Monte Carlo and HMM

- Assume  $g_t(x_t|x_{t-1}) = p_x(x_t|x_{t-1})$

$$\begin{aligned}
 w_t &= \frac{f(\mathbf{x}_{1:t}|\mathbf{y}_{1:t})}{g(\mathbf{x}_{1:t})} \\
 &\propto \frac{f(\mathbf{x}_{1:t-1}|\mathbf{y}_{1:t-1})p_x(x_t|x_{t-1})p_y(y_t|x_t)}{p_{x_1}(x_1)\prod_{s=2}^t p_x(x_s|x_{s-1})} \\
 &= \frac{f(\mathbf{x}_{1:t-1}|\mathbf{y}_{1:t-1})}{g(\mathbf{x}_{1:t-1})} \frac{p_x(x_t|x_{t-1})p_y(y_t|x_t)}{p_x(x_t|x_{t-1})} \\
 &= w_{t-1}p_y(y_t|x_t)
 \end{aligned}$$

- **Algorithm**

- 1 Sample  $X_1^i \sim p_{x_1}(\cdot), i = 1, \dots, n.$
- 2 Let  $w_1^{*i} = u_1^i = p_y(y_1|x_1^i)$ , normalize to  $w_1^i = w_1^{*i} / \sum_j w_1^{*j}$ . Set  $t = 2$
- 3 Sample  $X_t^i|x_{t-1}^i \sim p_x(x_t|x_{t-1}^i), i = 1, \dots, n.$
- 4 Append  $x_t^i$  to  $\mathbf{x}_{1:t-1}^i$ , obtaining  $\mathbf{x}_t^i$
- 5 Let  $u_t^i = p_y(y_t|x_t^i)$
- 6 Let  $w_t^{*i} = w_{t-1}^i u_t^i$ , **normalize** to  $w_t^i = w_t^{*i} / \sum_j w_t^{*j}$ .
- 7 If  $\hat{N}_{eff}$  is small, perform resampling
- 8 Increment  $t$  and return to step 3

# Terrain navigation (next time)

- Assume movement model for airplane

$$\mathbf{X}_t = \mathbf{x}_{t-1} + \mathbf{d}_t + \boldsymbol{\varepsilon}_t$$

$\mathbf{d}_t$  = Drift of plane measured by internal navigation system (assumed known)

$$\boldsymbol{\varepsilon}_t = \mathbf{R}_t^T \mathbf{Z}_t$$

$$\mathbf{R}_t = \frac{1}{x_{1,t-1}^2 + x_{2,t-1}^2} \begin{pmatrix} -x_{1,t-1} & x_{2,t-1} \\ -x_{2,t-1} & -x_{1,t-1} \end{pmatrix}$$

$$\mathbf{Z}_t \sim N_2 \left( \mathbf{0}, q^2 \begin{pmatrix} 1 & 0 \\ 0 & k^2 \end{pmatrix} \right)$$

$$q = 400, k = 0.5$$

$$Y_t = m(\mathbf{x}_t) + \delta_t$$

$m(\mathbf{x}_t)$  = Elevation at point  $\mathbf{x}_t$

- Example\_6\_7.R