

# UNIVERSITETET I OSLO

## Matematisk Institutt

EXAM IN: **STK 4060/9060 – Time Series**  
WITH: **Nils Lid Hjort**  
AUXILIA: **Calculator, plus one single sheet of paper  
with the candidate's own personal notes**  
TIME FOR EXAM: **Thursday 2/vi/2022, 15:00 – 19:00**

This exam set contains four exercises and comprises four pages (including a simple appendix on the last page).

### Exercise 1: spectral densities

SUPPOSE THAT  $x_t$  is a stationary time series process, for  $t = 0, \pm 1, \pm 2, \dots$ , with finite covariances  $\gamma(h) = \text{cov}(x_t, x_{t+h})$  for  $h = 0, \pm 1, \pm 2, \dots$ . Assume furthermore that the associated spectral domain distribution has a density  $f(\omega)$  on  $[-\frac{1}{2}, \frac{1}{2}]$ . You may take for granted here that  $\gamma(h)$  can be represented as  $\int_{-1/2}^{1/2} \exp(2\pi i \omega h) f(\omega) d\omega$ , for each  $h$ , and that the spectral density can be expressed as

$$f(\omega) = \sum_{h=-\infty}^{\infty} \gamma(h) \exp(-2\pi i \omega h) \quad \text{for } -\frac{1}{2} \leq \omega \leq \frac{1}{2}.$$

For the following, you may find it useful to check with the few facts for complex numbers given in the Appendix.

- (a) Show that  $f(\omega) = \gamma(0) + 2 \sum_{h=1}^{\infty} \gamma(h) \cos(2\pi h \omega)$ . If the  $x_t$  are actually independent, what is its  $f(\omega)$ ?
- (b) Now consider a transformation of the  $x_t$  process to a new  $y_t$  process, defined by

$$y_t = \sum_j c_j x_{t-j},$$

for certain coefficients  $c_j$ . Show that

$$\gamma^*(h) = \text{cov}(y_t, y_{t+h}) = \sum_{r,s} c_r c_s \gamma(|h+r-s|).$$

- (c) Use this to show that the  $y_t$  process has spectral density

$$f^*(\omega) = |A(\omega)|^2 f(\omega),$$

in terms of the squared modulus of  $A(\omega) = \sum c_j \exp(2\pi i j \omega)$ .

- (d) Then study the case of  $y_t = x_t + b x_{t-1}$ . Show that its spectral density can be written

$$f^*(\omega) = \{1 + b^2 + 2b \cos(2\pi \omega)\} f(\omega).$$

## Exercise 2: moving average of moving average

WE START OUT CONSIDERING a simple moving average process, of the type  $x_t = w_t + aw_{t-1}$  for  $t = 1, 2, \dots$ , in terms of i.i.d. white noise zero-mean variables  $w_0, w_1, \dots$ , and where we for simplicity take these to have variance  $\sigma_w^2 = 1$ .

- (a) Give formulae for the variance  $\gamma(0)$  and for the covariances  $\gamma(h) = \text{cov}(x_t, x_{t+h})$  for  $h = 1, 2, \dots$ . Show also that the one-step correlation, between  $x_t$  and  $x_{t+1}$ , is  $\rho(1) = a/(1 + a^2)$ . What is the range of possible values, for this correlation?
- (b) For an observed time series  $x_1, \dots, x_n$ , give a formula for the empirical autocorrelation  $\hat{\rho}(1)$ . Explain how you may estimate the parameter  $a$  based on this.
- (c) Show that the spectral density for the  $x_t$  process becomes

$$f(\omega) = 1 + a^2 + 2a \cos(2\pi\omega).$$

- (d) Then define the moving average of the initial moving average process, by  $y_t = x_t + bx_{t-1}$ . Find the spectral density  $f^*(\omega)$  using results from Exercise 1.
- (e) Show that the  $y_t$  also can be expressed as a moving average process of order two. Find formulae for the covariances  $\gamma(0), \gamma(1), \gamma(2)$  using this representation, and use this to find the spectral density once more. Verify that the two formulae you now have derived for the spectral density for the  $y_t$  process are the same.

## Exercise 3: the periodogramme and the Whittle log-likelihood

THE DISCRETE FOURIER TRANSFORM, for an observed time series  $x_1, \dots, x_n$ , is defined as

$$d(\omega) = (1/\sqrt{n}) \sum_{t=1}^n x_t \exp(-2\pi i\omega t).$$

- (a) Give a formula for the inverse Fourier transform, where  $x_1, \dots, x_n$  can be retrieved from  $d(\omega_j)$  computed at  $\omega_j = j/n$  for  $j = 0, 1, \dots, n$ . (You are not asked here to prove such a formula.)
- (b) The periodogramme for the observed sequence is  $I(\omega_j) = |d(\omega_j)|^2$ , for  $\omega_j = j/n$ . Give a formula for this  $I(\omega_j)$ , in terms of  $\sum_{t=1}^n x_t \cos(2\pi\omega_j t)$  and  $\sum_{t=1}^n x_t \sin(2\pi\omega_j t)$ .
- (c) If the  $x_t$  series is stationary, describe briefly how the periodogramme relates to its spectral density  $f(\omega)$ .
- (d) Consider now the simple order-one zero-mean moving average process studied in Exercise 2, with  $x_t = w_t + aw_{t-1}$  and with  $\sigma_w = 1$ , and where we have seen that the spectral density takes the form  $f(\omega) = 1 + a^2 + 2a \cos(2\pi\omega)$ . For an observed time series  $x_1, \dots, x_n$  from this model, define the Whittle log-likelihood function  $\ell^w(a)$ .

- (e) I have data for two separate and independently observed time series, say  $x_{A,1}, \dots, x_{A,100}$  and  $x_{B,1}, \dots, x_{B,100}$ , both assumed to follow the simple MA(1) model above, but with parameters  $a_A$  and  $a_B$  that are not necessarily equal. In the figure below I've plotted the Whittle log-likelihood functions  $\ell_A^w(a_A)$  and  $\ell_B^w(a_B)$ . For the A data,  $\ell_A^w$  is maximised for 0.5622 with 2nd order derivative  $-95.1402$ ; similarly, for the B data,  $\ell_B^w$  is maximised for 0.3093 with 2nd order derivative  $-105.6288$ . (i) Give approximate 95 percent confidence intervals for  $a_A$  and for  $a_B$ . (ii) Test the null hypothesis that the two parameters are equal.

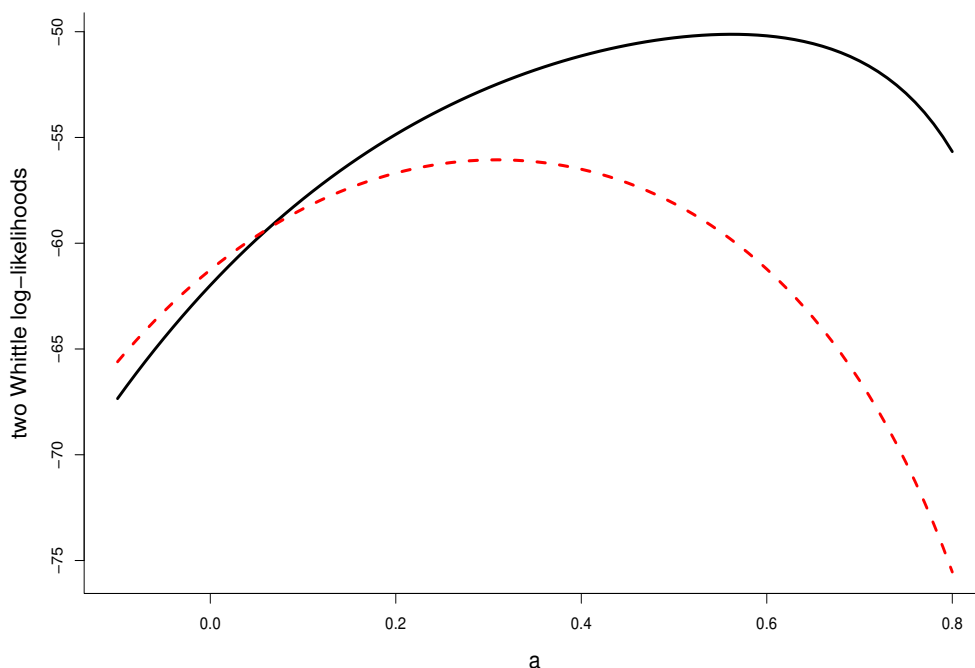


Figure: The two Whittle log-likelihood functions,  $\ell_A^w(a_A)$  (black, full) and  $\ell_B^w(a_B)$  (red, dashed), for the two time series datasets.

#### Exercise 4: equations for the AR(2) model

CONSIDER THE SECOND-ORDER AUTOREGRESSIVE MODEL, of the form

$$x_t = \phi_1 x_{t-1} + \phi_2 x_{t-2} + w_t,$$

taken here for simplicity to have zero mean, and with the  $w_t$  being i.i.d. with variance  $\sigma_w^2$ . Below, we let as usual  $\gamma(h) = \text{cov}(x_t, x_{t-h})$  for  $h = 0, \pm 1, \pm 2, \dots$

- (a) What is the requirement on  $(\phi_1, \phi_2)$ , to ensure that the  $x_t$  process is stationary and does not explode (i.e. is causal, to use the book's term)? Assume in the following that this requirement is met.
- (b) Multiply  $x_t - \phi_1 x_{t-1} - \phi_2 x_{t-2} - w_t = 0$  with  $x_t$ , and show that

$$\gamma(0) - \phi_1 \gamma(1) - \phi_2 \gamma(2) = \sigma_w^2.$$

- (c) Derive, in perhaps a similar manner, equations for  $\gamma(1)$  and  $\gamma(2)$ . Show that these can be organised as

$$\begin{pmatrix} \gamma(0) & \gamma(1) \\ \gamma(1) & \gamma(0) \end{pmatrix} \begin{pmatrix} \phi_1 \\ \phi_2 \end{pmatrix} = \begin{pmatrix} \gamma(1) \\ \gamma(2) \end{pmatrix}.$$

- (d) Having observed a time series  $x_1, \dots, x_n$  from this zero-mean AR(2) model, explain how you can use the equations of (d) to estimate  $\phi_1, \phi_2$ . How can you then estimate  $\sigma_w$ ?

### Appendix: just a few things for complex numbers

For  $z = a + ib$  a complex number, with  $i$  the famous square root of  $-1$ , its conjugate number is  $\bar{z} = a - ib$ , and its squared length is  $|z|^2 = z\bar{z} = a^2 + b^2$ . Also, for  $z = \sum_r c_r \exp(2\pi ir)$ , it follows that

$$|z|^2 = \sum_{r,s} c_r c_s \exp(2\pi i(r - s)).$$

Here, as usual,  $\exp(2\pi iu) = \cos(2\pi u) + i \sin(2\pi u)$ ; also, famously,  $\cos(-x) = \cos x$  and  $\sin(-x) = -\sin x$ .