

SOLUTION TO EXERCISES

WEEK 45

Exercise 5.1

We have counting processes $N_1(t), \dots, N_n(t)$
 with intensity processes $\lambda_i(t) = \gamma_i(t) e^\beta$
 where $\gamma_i(t)$ is an at risk indicator ($i=1, \dots, n$)

We introduce $R_i(t) = \int_0^t \gamma_i(u) du$ and let
 $N_0(t) = \sum_{i=1}^n N_i(t)$ and $R_0(t) = \sum_{i=1}^n R_i(t)$.

a) The log-likelihood is given by [cf. (5.5)]

$$l(\beta) = \sum_{i=1}^n \int_0^T \log \lambda_i(t) dN_i(t) - \int_0^T \lambda_0(t) dt$$

By inserting $\lambda_i(t) = \gamma_i(t) e^\beta$ and $\lambda_0(t) = \gamma_0(t) e^\beta$

we find

$$l(\beta) = \sum_{i=1}^n \int_0^T \log(\gamma_i(t) e^\beta) dN_i(t) - e^\beta \int_0^T \gamma_0(t) dt$$

$$= \beta \sum_{i=1}^n \int_0^T dN_i(t) - e^\beta R_0(T) + \text{const}$$

$$= \beta N_0(T) - e^\beta R_0(T) + \text{const}$$

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The score function becomes

$$U(\beta) = \ell'(\beta) = N \cdot \lambda - e^{\beta} R \cdot \lambda$$

This gives for the MLE:

$$e^{\hat{\beta}} R \cdot \lambda = N \cdot \lambda$$

$$e^{\hat{\beta}} = \frac{N \cdot \lambda}{R \cdot \lambda} = \hat{\nu}$$

$$\text{So } \hat{\beta} = \log \hat{\nu}$$

b) By the general result for MLE for counting process models, we have that [cf section 5.1.4]

$$\hat{\beta} \stackrel{\text{approx}}{\sim} N(\beta, \sigma^2)$$

where $\hat{\sigma}^2 = 1/I(\hat{\beta})$, where $I(\beta)$

is the observed information

$$I(\beta) = -U'(\beta) = -\ell''(\beta)$$

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Thus $I(\beta) = e^{\beta} R_0(\beta)$, so

the variance may be estimated by

$$\hat{\sigma}^2 = \frac{1}{e^{\hat{\beta}} R_0(\hat{\beta})} = \frac{1}{\frac{N_0(\hat{\beta})}{R_0(\hat{\beta})} R_0(\hat{\beta})} = \frac{1}{N_0(\hat{\beta})}$$

Remark: We may alternatively derive the result in questions a and b by the delta-method, which states that [cf. exercise 3.3] for a smooth function $g(v)$, we have that

$$g(\hat{v}) \approx g(v) + g'(v)(\hat{v} - v)$$

from where we obtain that

$$g(\hat{v}) \overset{\text{approx}}{\sim} N(g(v), [g'(v)]^2 \sigma^2)$$

where σ^2 is the asymptotic variance of \hat{v} . By the result of example 5.1, we have that

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σ^2 may be estimated by

$$\hat{\sigma}^2 = \hat{v}^2 / N_0(r) = \hat{v} / R_0(r)$$

Thus we obtain that

$\hat{\beta} = \log \hat{v}$ is approximately normally distributed around $\beta = \log v$ with

a variance that may be estimated

$$\text{by } [(\log \hat{v})']^2 \hat{\sigma}^2 = \left(\frac{1}{\hat{v}}\right)^2 \frac{\hat{v}^2}{N_0(r)} = \frac{1}{N_0(r)}$$

c) To obtain a confidence interval for $v = e^\beta$, we first derive a confidence interval for β , namely

$$\hat{\beta} \pm 1.96 \frac{1}{\sqrt{N_0(r)}}$$

By exponentiating the limits of this confidence interval, we obtain the following 95% confidence interval for v :

$$e^{\hat{\beta} \pm 1.96 / \sqrt{N_0(r)}} \quad \text{i.e.} \quad \hat{v} e^{\pm 1.96 / \sqrt{N_0(r)}}$$

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Exercise 5.4

We have counting processes $N_1(t), \dots, N_n(t)$ with intensity processes $\lambda_i(t) = \nu Y_i(t)$

a) The log-likelihood is given by

$$\begin{aligned} l(\nu) &= \sum_{i=1}^n \int_0^T \log \lambda_i(t) dN_i(t) - \int_0^T \lambda_i(t) dt \\ &= \sum_{i=1}^n \int_0^T \log(\nu Y_i(t)) dN_i(t) - \nu \int_0^T Y_i(t) dt \end{aligned}$$

$$= \log \nu \cdot N_0(T) - \nu R_0(T) + \text{const}$$

Hence the score function becomes

$$U(\nu) = l'(\nu) = \frac{N_0(T)}{\nu} - R_0(T)$$

b) At the true parameter value we have that $N_i(t) = \int_0^t \nu Y_i(t) dt + M_i(t)$ and hence

$$N_0(T) = \nu R_0(T) + M_0(T)$$

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Thus the score evaluated at the true value of v is given by

$$\begin{aligned}U(v) &= \frac{N_0(p)}{v} - R_0(p) \\&= \frac{vR_0(p) + M_0(p)}{v} - R_0(p) \\&= R_0(p) + \frac{M_0(p)}{v} - R_0(p) = \frac{M_0(p)}{v}\end{aligned}$$

From this it follows that

$$E(U(v)) = E\left(\frac{M_0(p)}{v}\right) = 0$$

and that the variance of the score equals

$$\text{Var}(U(v)) = \frac{1}{v^2} \text{Var}(M_0(p))$$

$$= \frac{1}{v^2} E\langle M_0 \rangle(p)$$

$$= \frac{1}{v^2} E\left\{ \int_0^p \lambda_0(t) dt \right\}$$

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$$= \frac{1}{v^2} E \left\{ v \int_0^T Y_0(t) dt \right\}$$

$$= E(R_0(\pi)) / v$$

c) The observed information is given by

$$i(v) = -U'(v) = \frac{N_0(\pi)}{v^2}$$

When evaluated at the true parameter value we have

$$\begin{aligned} i(v) &= \frac{N_0(\pi)}{v^2} = \frac{v R_0(\pi) + M_0(\pi)}{v^2} \\ &= \frac{R_0(\pi)}{v} + \frac{M_0(\pi)}{v^2} \end{aligned}$$

and hence

$$E(i(v)) = E \left(\frac{R_0(\pi)}{v} \right)$$

i.e. the same as the variance of the score