

## Exercise 4.1: Cox Score function and Information

Assume that  $N_i(t)$  are counting processes with intensity processes  $\lambda_i(t) = Y_i(t)\alpha_0(t) \exp(\beta' x_i)$ . The log-partial likelihood can then be written

$$\log(L(\beta)) = \int_0^\tau \sum_{i=1}^n [\beta' x_i - \log(S^{(0)}(\beta, t))] dN_i(t)$$

where

$$S^{(0)}(\beta, t) = \sum_{i=1}^n Y_i(t) \exp(\beta' x_i)$$

Let also

- $S^{(1)}(\beta, t) = \sum_{k=1}^n x_k Y_k(t) \exp(\beta' x_k) = \frac{\partial S^{(0)}(\beta, t)}{\partial \beta}$
- $S^{(2)}(\beta, t) = \sum_{k=1}^n x_k x_k^\top Y_k(t) \exp(\beta' x_k) = \frac{\partial^2 S^{(0)}(\beta, t)}{\partial \beta^2}$

## 4.1a: Cox score function

$$\begin{aligned} U(\beta) &= \frac{\partial \log(L(\beta))}{\partial \beta} = \sum_{i=1}^n \int_0^{\tau} \left[ x_i - \frac{\partial \log(S^{(0)}(\beta, t))}{\partial \beta} \right] dN_i(t) \\ &= \sum_{i=1}^n \int_0^{\tau} \left[ x_i - \frac{S^{(1)}(\beta, t)}{S^{(0)}(\beta, t)} \right] dN_i(t) \end{aligned}$$

since

$$\frac{\partial \log(S^{(0)}(\beta, t))}{\partial \beta} = \frac{1}{S^{(0)}(\beta, t)} \frac{\partial S^{(0)}(\beta, t)}{\partial \beta} = \frac{S^{(1)}(\beta, t)}{S^{(0)}(\beta, t)}$$

## 4.1b: Cox information matrix

$$\begin{aligned}
 I(\beta) &= -\frac{\partial U(\beta)}{\partial \beta^\top} = -\sum_{i=1}^n \int_0^\tau \left[ -\frac{\partial(S^{(1)}(\beta,t)/S^{(0)}(\beta,t))}{\partial \beta^\top} \right] dN_i(t) \\
 &= \sum_{i=1}^n \int_0^\tau \left[ \frac{S^{(2)}(\beta,t)}{S^{(0)}(\beta,t)} - \frac{S^{(1)}(\beta,t)S^{(1)}(\beta,t)^\top}{S^{(0)}(\beta,t)^2} \right] dN_i(t) \\
 &= \int_0^\tau \left[ \frac{S^{(2)}(\beta,t)}{S^{(0)}(\beta,t)} - \frac{S^{(1)}(\beta,t)S^{(1)}(\beta,t)^\top}{S^{(0)}(\beta,t)^2} \right] dN_\bullet(t)
 \end{aligned}$$

since  $\frac{\partial(S^{(1)}(\beta,t)/S^{(0)}(\beta,t))}{\partial \beta^\top} =$

$$\begin{aligned}
 &\frac{1}{S^{(0)}(\beta,t)} \frac{\partial S^{(1)}(\beta,t)}{\partial \beta^\top} - S^{(1)}(\beta,t) \frac{\partial(1/S^{(0)}(\beta,t))}{\partial \beta^\top} \\
 &= \frac{S^{(2)}(\beta,t)}{S^{(0)}(\beta,t)} - \frac{S^{(1)}(\beta,t)S^{(1)}(\beta,t)^\top}{S^{(0)}(\beta,t)^2}
 \end{aligned}$$

## Exercise 4.2: Cox Score test and Log-rank test

Assume that  $x_i$  is a binary variable. Then,

$$S^{(1)}(0, t) = \sum_{i=1}^n x_i Y_i(t) \exp(0x_i) = Y_{\bullet 1}(t),$$

i.e. the no. at risk with  $x_i = 1$ . Correspondingly

$$S^{(0)}(0, t) = \sum_{i=1}^n Y_i(t) \exp(0x_i) = Y_{\bullet 0}(t) + Y_{\bullet 1}(t) = Y_{\bullet}(t),$$

i.e. the total no. at risk. Thus

$$U(0) = \sum_{i=1}^n \int [x_i - \frac{S^{(1)}(0, t)}{S^{(0)}(0, t)}] dN_i(t) = \mathbf{O}_1 - \mathbf{E}_1$$

where  $\mathbf{O}_1$  is no. deaths with  $x_i = 1$  and  $\mathbf{E}_1 = \int Y_{\bullet 1}(t) \frac{dN_{\bullet}(t)}{Y_{\bullet}(t)}$ , thus

$U(0)$  is equal to the log-rank statistic!

## Comparison Cox-regression and Log-rank

```
> summary(coxph(Surv(lifetime,dead)~ulcer,data=mel))
```

	coef	exp(coef)	se(coef)	z	p
ulcer	-1.47	0.23	0.295	-4.98	6.3e-07

  

	exp(coef)	exp(-coef)	lower .95	upper .95
ulcer	0.23	4.36	0.129	0.41

```
Rsquare= 0.13 (max possible= 0.937 )
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```
Likelihood ratio test= 28.4 on 1 df, p=9.68e-08
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Wald test = 24.8 on 1 df, p=6.3e-07
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```
Score (logrank) test = 29.6 on 1 df, p=5.41e-08
```

```
> survdiff(Surv(lifetime,dead)~ulcer,data=mel)
```

	N	Observed	Expected	(O-E)^2/E	(O-E)^2/V
ulcer=1	90	41	21.2	18.5	29.6
ulcer=2	115	16	35.8	10.9	29.6

```
Chisq= 29.6 on 1 degrees of freedom, p= 5.41e-08
```

## Variance of $U(0) = \mathbf{O}_1 - \mathbf{E}_1$

is estimated by

$$I(\beta) = \int \left\{ \frac{S^{(2)}(0, t)}{S^{(0)}(0, t)} - \left[ \frac{S^{(1)}(0, t)}{S^{(0)}(0, t)} \right]^2 \right\} dN_{\bullet}(t)$$

But since  $x_i$  is binary we get

$$S^{(2)}(0, t) = \sum_{i=1}^n x_i^2 Y_i(t) = \sum_{i=1}^n x_i Y_i(t) = S^{(1)}(0, t) = Y_{\bullet 1}(t)$$

Thus

$$I(0) = \int \left\{ \frac{Y_{\bullet 1}(t)}{Y_{\bullet}(t)} - \left[ \frac{Y_{\bullet 1}(t)}{Y_{\bullet}(t)} \right]^2 \right\} dN_{\bullet}(t) = \int \frac{Y_{\bullet 0}(t) Y_{\bullet 1}(t)}{Y_{\bullet}(t)^2} dN_{\bullet}(t)$$

which equals the previously derived variance formula of  $\mathbf{O}_1 - \mathbf{E}_1$  (and ignoring ties).

## Generalized log-rank test

For general  $x_i$  (not necessarily binary) we have

$$U(0) = \sum_{i=1}^n \int \left[ x_i - \frac{\sum_{k=1}^n x_k Y_k(t)}{Y(t)} \right] dN_i(t)$$

which has variance under  $H_0 : \beta = 0$  given by

$$I(0) = \int \left\{ \frac{\sum_{i=1}^n x_i^2 Y_i(t)}{Y(t)} - \left[ \frac{\sum_{i=1}^n x_i Y_i(t)}{Y(t)} \right]^2 \right\} dN_{\bullet}(t)$$

The generalized log-rank test is the given by that under  $H_0$ :

$$\frac{U(0)^2}{I(0)} \sim \chi_1^2$$