

Chapter 8 - Hierarchical DSTMs: Implementation and Inference

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Hierarchical Dynamical Spatio-Temporal Models

- Data in Process models
- Observation types
- Linear observations
- Kalman filter
- Kalman smoother
- nonlinear/non Gaussian

Bayesian approach: Also include model for parameters

Methodology for inference in Hierarchical Dynamical Spatio-Temporal Models

- General Problem
- Sequential vs non sequential
- Kalman filter
- EM-algorithm
- MCMC
- Sequential Monte Carlo, particle filter
- INLA

Hierarchical model

- Model for $p[\mathbf{Z}|\mathbf{Y}, \boldsymbol{\theta}_D]$
- Model for $p[\mathbf{Y}|\boldsymbol{\theta}_P]$
- Bayesian approach: Model for $p[\boldsymbol{\theta}_D, \boldsymbol{\theta}_P]$

Inference: Extract information about $\boldsymbol{\theta}$ and \mathbf{Y} from \mathbf{Z}

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Inference: Extract information about $\boldsymbol{\theta}$ and \mathbf{Y} from \mathbf{Z}

Likelihood:

$$p[\mathbf{Z}|\boldsymbol{\theta}] = \int_{\mathbf{Y}} p[\mathbf{Z}|\mathbf{Y}, \boldsymbol{\theta}_D] p[\mathbf{Y}|\boldsymbol{\theta}_P] d\mathbf{Y}$$

Bayesian posterior

$$p[\boldsymbol{\theta}, \mathbf{Y}|\mathbf{Z}] = \frac{p[\boldsymbol{\theta}, \mathbf{Y}] p[\mathbf{Z}|\mathbf{Y}, \boldsymbol{\theta}]}{p[\mathbf{Z}]}$$
$$p[\mathbf{Z}] = \int_{\boldsymbol{\theta}} p[\mathbf{Z}|\boldsymbol{\theta}] p[\boldsymbol{\theta}] d\boldsymbol{\theta}$$

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How to obtain these quantities:

- Huge computational problem
- Very active research field
- Some general methods
- Software for specific (classes of) models

- Data $\mathbf{Z} = (\mathbf{Z}_1, \dots, \mathbf{Z}_T)$
- $p(\mathbf{Z}) = \prod_t p(\mathbf{Z}_t | \mathbf{Z}_1, \dots, \mathbf{Z}_{t-1})$
- Sequential updating:
$$p(\boldsymbol{\theta}, \mathbf{Y}_{1:t} | \mathbf{Z}_{1:t}), \quad t = 1, 2, 3, \dots$$
- Non-sequential updating
$$p(\boldsymbol{\theta}, \mathbf{Y}_{1:T} | \mathbf{Z}_{1:T})$$

- Kalman filter - sequential
- Markov chain Monte Carlo - nonsequential
- Sequential Monte Carlo - sequential
- INLA - nonsequential
- Ensemble Kalman Filter - sequential
- Ensemble (Kalman) Smoother - nonsequential (not in book)

Model

$$\mathbf{Y}_t = \mathbf{M}_t \mathbf{Y}_{t-1} + \boldsymbol{\eta}_t,$$

$$\boldsymbol{\eta}_t \stackrel{\text{ind}}{\sim} N(0, \mathbf{Q}_t)$$

$$\mathbf{Z}_t = \mathbf{H}_t \mathbf{Y}_t + \boldsymbol{\varepsilon}_t,$$

$$\boldsymbol{\varepsilon}_t \stackrel{\text{ind}}{\sim} N(0, \mathbf{R}_t)$$

Aim: Calculate $p(\mathbf{Y}_t | \mathbf{Z}_{1:t})$. Enough with

$$\hat{\mathbf{Y}}_{t|t} = E[\mathbf{Y}_t | \mathbf{Z}_{1:t}]$$

$$\mathbf{P}_{t|t} = \text{Var}[\mathbf{Y}_t | \mathbf{Z}_{1:t}]$$

Kalman filter

$$\mathbf{P}_{t|t-1} = \mathbf{M}_t \mathbf{P}_{t-1|t-1} \mathbf{M}_t^T + \mathbf{Q}_t$$

$$\hat{\mathbf{Y}}_{t|t-1} = \mathbf{M}_t \hat{\mathbf{Y}}_{t-1|t-1}$$

$$\mathbf{S}_t = \mathbf{H}_t^T \mathbf{P}_{t|t-1} \mathbf{H}_t + \mathbf{R}_t$$

$$\mathbf{K}_t = \mathbf{P}_{t|t-1} \mathbf{H}_t^T \mathbf{S}_t^{-1}$$

$$\mathbf{P}_{t|t} = [\mathbf{I} - \mathbf{K}_t \mathbf{H}_t] \mathbf{P}_{t|t-1}$$

$$\hat{\mathbf{Y}}_{t|t} = \hat{\mathbf{Y}}_{t|t-1} + \mathbf{K}_t [\mathbf{Z}_t - \mathbf{H}_t \hat{\mathbf{Y}}_{t|t-1}]$$

Likelihood:

$$L(\boldsymbol{\theta}) = p(\mathbf{Z}; \boldsymbol{\theta}) = \prod_{t=1}^T p(\mathbf{Z}_t | \mathbf{Z}_{1:t-1}; \boldsymbol{\theta})$$

Parameter estimation:

Kalman filter give $p(\mathbf{Y}_t|\mathbf{Z}_{1:t})$ given parameters.

- $L(\boldsymbol{\theta}) = p(\mathbf{Z}|\boldsymbol{\theta}) = \prod_{t=1}^T p(\mathbf{Z}_t|\mathbf{Z}_{1:t-1}; \boldsymbol{\theta})$
- This can be obtained directly from the Kalman filter (exercise)
- Can optimize wrt $\boldsymbol{\theta}$ to obtain ML estimates
- Can also do Bayesian versions.

Properties

- Computationally very efficient, sequential (online) inference
- Can use alternative filters calculating $\mathbf{P}_{t|t}^{-1}$ and utilizing that this often is sparse
- Can be extended to nonlinear models through linear approximations

Maximum likelihood estimates in case of missing data (or latent variables) , and unknown parameters.

General formulation:

Iterate Expectation and Maximization until convergence...

- E-step: Calculate expectation of log likelihood
 $E(\ln L(\theta|Z))|Z_{obs}, \hat{\theta}^{(i-1)}) = q(\theta|\hat{\theta}^{(i-1)})$
- M-step: Find θ that maximizes $q(\theta|\hat{\theta}^{(i-1)})$, and call this $\hat{\theta}^{(i)}$

Expectation (E-step) is to "get rid of" the latent variables (or missing data) to get back to the standard maximum likelihood estimator (M-Step).

$$\mathbf{Y}_t = \mathbf{M}\mathbf{Y}_{t-1} + \boldsymbol{\eta}_t,$$

$$\boldsymbol{\eta}_t \stackrel{\text{ind}}{\sim} N(0, \mathbf{Q})$$

$$\mathbf{Z}_t = \mathbf{H}_t\mathbf{Y}_t + \boldsymbol{\varepsilon}_t,$$

$$\boldsymbol{\varepsilon}_t \stackrel{\text{ind}}{\sim} N(0, \mathbf{R})$$

Unknown parameters: \mathbf{M} , \mathbf{R} , \mathbf{Q} , and $\boldsymbol{\mu}_Y$. Known parameter: \mathbf{H}_t

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Start with an initial guess on the parameters: $\mathbf{M}^{(0)}$, $\mathbf{R}^{(0)}$, $\mathbf{Q}^{(0)}$, $\boldsymbol{\mu}_Y^{(0)}$

Then iterate:

- E-step: Run a Kalman smoother (Filter-smoother/ Forward-backward) with current parameter estimates $\mathbf{M}^{(i)}$, $\mathbf{R}^{(i)}$, $\mathbf{Q}^{(i)}$, $\boldsymbol{\mu}_Y^{(i)}$ to obtain updated state estimates (estimate of latent process), i.e the estimates of $\mathbf{Y}_{1:T}$

$$\begin{aligned} \mathbf{Y}_t &= \mathbf{M}\mathbf{Y}_{t-1} + \boldsymbol{\eta}_t, & \boldsymbol{\eta}_t &\stackrel{\text{ind}}{\sim} N(0, \mathbf{Q}) \\ \mathbf{Z}_t &= \mathbf{H}_t\mathbf{Y}_t + \boldsymbol{\varepsilon}_t, & \boldsymbol{\varepsilon}_t &\stackrel{\text{ind}}{\sim} N(0, \mathbf{R}) \end{aligned}$$

Unknown parameters: \mathbf{M} , \mathbf{R} , \mathbf{Q} , and $\boldsymbol{\mu}_Y$. Known parameter: \mathbf{H}_t

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- M-step: Assume the state space (latent process) is observed, and use them in the maximum likelihood estimation together with $\mathbf{Z}_{1:T}$ to get updated parameter estimates of $\mathbf{M}^{(i+1)}$, $\mathbf{R}^{(i+1)}$, $\mathbf{Q}^{(i+1)}$, $\boldsymbol{\mu}_Y^{(i+1)}$.

Interest in $\hat{g} = E[g(\mathbf{Y}, \boldsymbol{\theta})|\mathbf{Z}]$

Monte Carlo

- Sample $\{(\mathbf{Y}^{(s)}, \boldsymbol{\theta}^{(s)})\}$ from $p[\mathbf{Y}, \boldsymbol{\theta}|\mathbf{Z}]$
- Approximate \hat{g} by

$$\frac{1}{S} \sum_{s=1}^S g(\mathbf{Y}^{(s)}, \boldsymbol{\theta}^{(s)})$$

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- Difficult to sample from $p[\mathbf{Y}, \boldsymbol{\theta}|\mathbf{Z}]$
- MCMC: Tool for complex simulation
- Very general (we have looked at Gibbs sampler)
- Non-sequential (offline) inference
- Some general software (Winbugs), SLOWWWW
- Usually need to implement from scratch to make it efficient
- Time-consuming both in implementation time and running time
- Separate courses for this

Sequential Monte Carlo (SMC)

- Kalman filter: Calculate $p[\mathbf{Y}_t | \mathbf{Z}_{1:t}, \boldsymbol{\theta}_D]$ analytically
- SMC: Approximate $p[\mathbf{Y}_t | \mathbf{Z}_{1:t}, \boldsymbol{\theta}_D]$ by Monte Carlo samples
- Utilize samples from time $t - 1$ when sampling at time t
- Differ from MCMC in performing simulations **sequentially**
- Very efficient in low dimensions of \mathbf{Y}_t , slow for high dimensions
- Active field, much progress made continuously!

"Sampling based Kalman filter".

Initial forecast:

Sample $\tilde{\mathbf{Y}}_0^{(l)} \sim p(\mathbf{y}_0)$, for $l = 1, \dots, L$

Iterate for all time steps:

- Forecast step is sampling:

Sample $\tilde{\mathbf{y}}_t^l \sim p(\mathbf{y}_t | \mathbf{Y}_{t-1}^{(l)})$, for $l = 1, \dots, L$

Set $\tilde{\mathbf{Y}}_{0:t}^{(l)} = [\tilde{\mathbf{Y}}_{0:t-1}^{(l)} \tilde{\mathbf{Y}}_t^{(l)}]$

- Filter step is importance re-sampling:

* Evaluate the importance weight (i.e. likelihood) $w_t^{(l)} = p(\mathbf{Z}_t | \tilde{\mathbf{Y}}_t^{(l)})$

* Resample with replacement L particles $[\mathbf{Y}_{0:t}^{(l)}, l=1, \dots, L]$ from $\tilde{\mathbf{Y}}_{0:t}^{(l)}$ using the importance weight for resampling i.e. the probability of sampling particle k (at time t) is $w_t^{(k)} / \sum_l w_t^{(l)}$

Integrated nested Laplace approximation (INLA)

Interest in calculation of

$$L(\theta) = \int_{\mathbf{Y}} p[\mathbf{Z}|\mathbf{Y}, \theta_D] p[\mathbf{Y}|\theta_P] d\mathbf{Y} = \int_{\mathbf{Y}} e^{f(\mathbf{Y})} d\mathbf{Y}$$

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Approximate (using $\hat{\mathbf{Y}} = \operatorname{argmax}_{\mathbf{Y}} f(\mathbf{Y})$, $\mathbf{f}'(\hat{\mathbf{Y}}) = \mathbf{0}$)

$$f(\mathbf{Y}) \approx \hat{f}(\mathbf{Y})$$

$$= f(\hat{\mathbf{Y}}) + \mathbf{f}'(\hat{\mathbf{Y}})(\mathbf{Y} - \hat{\mathbf{Y}}) + \frac{1}{2}(\mathbf{Y} - \hat{\mathbf{Y}})^T \mathbf{f}''(\hat{\mathbf{Y}})(\mathbf{Y} - \hat{\mathbf{Y}})$$

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which gives (with $d = \dim(\mathbf{Y})$)

$$L(\theta) \approx e^{f(\hat{\mathbf{Y}})} \frac{(2\pi)^{d/2}}{|\mathbf{f}''(\hat{\mathbf{Y}})|^{1/2}}$$

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INLA

- **Nested** Laplace approximations
- Utilize sparsity in precision matrices

- Require latent field to be Gaussian
- Great flexibility with respect to observation processes
- Extremely fast compared to MCMC
- The number of models it can cover is increasing frequently
- R overhead makes it relatively easy to use
- Only give **marginal** distributions, not simultaneous ones.