UNIVERSITY OF OSLO

Faculty of Mathematics and Natural Sciences

Examination in	STK4400 — Risk and reliability analysis
Day of examination:	Wednesday June 15, 2016
Examination hours:	14.30-18.30
This problem set consists of 5 pages.	
Appendices:	None
Permitted aids:	Calculator

Please make sure that your copy of the problem set is complete before you attempt to answer anything.

Problem 1

(a) Give the definitions of a multistate strongly coherent system, a multistate coherent system and a multistate weakly coherent system. Explain the mutual relations between these three.

(b) Give the definition of a binary type multistate monotone system. What is the idea behind this measure?

Problem 2

(a) Assume that X_1, \ldots, X_n are associated random variables. Show that for $j \in \{1, \ldots, M\}$ we have for the multistate series system.

$$P[\min_{1 \le i \le n} X_i \ge j] \ge \prod_{i=1}^n p_i^j \tag{1}$$

(b) Let ϕ be a multistate structure function that is nondecreasing in each argument and we assume that

$$\min_{1 \le i \le n} x_i \le \phi(\boldsymbol{x}) \tag{2}$$

Assume that X_1, \ldots, X_n are associated random variables. Show that we have

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$$\prod_{i=1}^{n} p_i^j \le p_{\phi}^j \tag{3}$$

Comment on the result.

Problem 3

Let X be a vector of environmental variables with sample space $\mathcal{X} \subseteq \mathbb{R}^n$. Furthermore let $P_e \in (0, 0.5)$ be a given exceedence probability. We shall in this exercise look into how one can identify a convex set $\mathcal{B} \subset \mathcal{X}$ such that for all hyperplanes Π that are tangents to \mathcal{B} , we have $P[\mathbf{X} \in \Pi^+] = P_e$, where Π^+ denotes the halfspace that is bounded by the hyperplane Π and not containing \mathcal{B} . We also introduce Π^- denoting the halfspace that is complementary to Π^+ , such that $\mathcal{B} \subseteq \Pi^-$. The boundary of \mathcal{B} is denoted by $\partial \mathcal{B}$ and is what is called an *environmental contour*.

(a) Explain how such an environmental contour can be used to check whether a certain mechanical structure is safe.

We assume in the rest of the exercise that $\mathbf{X} = (T, H)$. For $\theta \in [0, 2\pi)$ we define the random variable $Y(\theta) = T \cos(\theta) + H \sin(\theta)$. We also introduce

the function $C(\theta)$ defined for $\theta \in [0, 2\pi)$:

$$C(\theta) = \inf\{C : P[T\cos(\theta) + H\sin(\theta) > C] = P_e\}.$$

(b) What property does $C(\theta)$ have related to the probability distribution of $Y(\theta)$?

We also introduce:

$$\begin{aligned} \Pi(\theta) &= \{(t,h) : t\cos(\theta) + h\sin(\theta) = C(\theta)\},\\ \Pi^+(\theta) &= \{(t,h) : t\cos(\theta) + h\sin(\theta) > C(\theta)\},\\ \Pi^-(\theta) &= \{(t,h) : t\cos(\theta) + h\sin(\theta) \le C(\theta)\}. \end{aligned}$$

Then it can be shown that the set \mathcal{B} can can be written as:

$$\mathcal{B} = \bigcap_{\theta \in [0, 2\pi)} \Pi^{-}(\theta).$$

We consider the intersection point $(t(\theta_1, \theta_2), h(\theta_1, \theta_2))$ between the hyperplanes $\Pi(\theta_1)$ $\Pi(\theta_2)$. (c) Show the first of the two following relations:

$$t(\theta_1, \theta_2) = \frac{\sin(\theta_2)C(\theta_1) - \sin(\theta_1)C(\theta_2)}{\sin(\theta_2 - \theta_1)}$$
$$h(\theta_1, \theta_2) = \frac{-\cos(\theta_2)C(\theta_1) + \cos(\theta_1)C(\theta_2)}{\sin(\theta_2 - \theta_1)}$$

Then let $(t(\theta), h(\theta)) = \lim_{\delta \to 0^+} (t(\theta, \theta + \delta), h(\theta, \theta + \delta)).$

(d) Show that:

$$\begin{pmatrix} t(\theta) \\ h(\theta) \end{pmatrix} = \begin{bmatrix} C(\theta) & -C'(\theta) \\ C'(\theta) & C(\theta) \end{bmatrix} \cdot \begin{pmatrix} \cos(\theta) \\ \sin(\theta) \end{pmatrix},$$

where $C'(\theta)$ denotes the derivative of $C(\theta)$, and explain in a short way how this expression can be used to identify $\partial \mathcal{B}$.

(e) What sort of contour is $\partial \mathcal{B}$ if $C(\theta) = R$ for all $\theta \in [0, 2\pi)$? Give a reason for your answer.

Assume we have carried through a Monte Carlo simulation based on the joint probability distribution of (T, H), and that we have generated n vectors:

$$(T_1, H_1), \dots, (T_n, H_n) \tag{4}$$

For a given $\theta \in [0, 2\pi)$ we then calculate:

$$Y_i(\theta) = T_i \cos(\theta) + H_i \sin(\theta), \quad i = 1, \dots, n$$
(5)

(f) Explain how we can use $Y_1(\theta), \ldots, Y_n(\theta)$ to estimate $C(\theta)$.

Problem 4

Consider events occuring on a time axis starting from time t = 0. Let $0 < S_1 < S_2 < \ldots$ be the event times and let $T_i = S_i - S_{i-1}$ be the times between events, where $S_0 = 0$. Let further N(t) for t > 0 denote the number of events in the time interval (0, t].

(a) Define what it means that the process S_1, S_2, \ldots is a non-homogeneous Poisson process (NHPP) with intensity function w(t). Define also the cumulative intensity function W(t) and write down the probability distribution of N(t) for a given t > 0.

A software system is subject to failures at random times caused by errors present in the code. Let N(t) be the cumulative number of failures experienced by time t. We

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assume that each failure is caused by exactly one error in the code, and that this error is successfully removed from the software after each failure. Hence N(t) also represents the cumulative number of errors detected and removed by time t.

A classical model for software reliability is the *time dependent error detection model of* Goel and Okumoto (the GO-model), where one assumes that N(t) is an NHPP with intensity function of the form

$$w(t) = \alpha \beta e^{-\beta t}$$

for parameters $\alpha > 0, \beta > 0$.

(b) Calculate the cumulative intensity function W(t). What is the interpretation of W(t)?

What is the limit of W(t) as $t \to \infty$? Explain why this limit suggests an interpretation of the parameter α as the "initial number of errors" in the code.

(c) Suppose that the software system has been run until time s > 0. The *conditional* reliability function R(t|s) of the software at time s is defined to be the probability that the software operates without failures for at least time t beyond time s. Show that

$$R(t|s) = \exp\left(-e^{-\beta s}W(t)\right).$$
(6)

The GO-model is in particular used in software testing. Suppose that values of α and β are estimated from previous data and hence can be assumed known. Then (6) can be used to calculate an *optimal testing time* s_0 by the requirement, on the testing time s, that the conditional reliability $R(t_0|s)$ should be at least equal to a given value r (0 < r < 1), for a given time t_0 .

(d) Derive the corresponding expression for s_0 in terms of α, β, r and t_0 . Discuss in particular the choice $t_0 = \infty$.

Recall the interpretation of the parameter α as the "initial number of errors" in the code. In the *Jelinski-Moranda model for software reliability* (the JM-model) the *true* initial number of errors, a, is an unknown parameter, which is hence a non-negative integer which we shall assume is at least one.

The basic assumption of the JM-model is that the time T_i between failure number i-1and failure number i is exponentially distributed with hazard rate

$$\lambda_i = (a - i + 1)b$$

for i = 1, 2, ..., a, where b > 0 is the second parameter of the model. It is further assumed that $T_1, T_2, ..., T_a$ are stochastically independent.

(e) Explain why the time S_a of the *a*th failure can be interpreted as the time when all errors in the code are found.

Derive an exact expression for $E(S_a)$, i.e., the expected time when the software is errorfree. Then verify the approximation

$$E(S_a) \approx \frac{\ln a}{b}$$

when a is large.