## FASIT til eksamen i STK4540, 03.01.2012

## Problem 1

a) The pure premium is the expected value of the portfolio pay-out  $\mathcal{X}$  during a given period; i.e.  $\pi^{pu} = E(\mathcal{X})$ . It is a break-even price that excatly balances cost when overhead and profit are not taken into account. The market premium is  $\pi = (1 + \gamma)\pi^{pu}$  where the loading  $\gamma$  fluctuates with the state of the market.

**b)** If  $X_1, + \ldots, X_J$  are independent claims from J policy holders and  $\mathcal{X} = X_1 + \ldots + X_J$ , then

$$\frac{\operatorname{sd}(\mathcal{X})}{E(\mathcal{X})} = \frac{\bar{\sigma}}{\sqrt{J}\bar{\xi}}$$

where  $\bar{\xi}$  and  $\bar{\sigma}^2$  are the average mean claim and average variance of claims for the entire population. The ratio goes to zero as  $J \to \infty$ .

c) The rule of double expectations yield for identical risks

$$E(\mathcal{X}) = J\xi_{\mu}\xi_{z}$$
 and  $\operatorname{var}(\mathcal{X}) = J^{2}\xi_{z}^{2}\sigma_{\mu}^{2} + J\xi_{\mu}(\xi_{z}^{2} + \sigma_{z}^{2})$ 

where  $\xi_{\mu}$  and  $\sigma_{\mu}$  are mean and standard deviation for  $\mu \xi_z$  and  $\sigma_z$  for claims. It follows that

$$\frac{\operatorname{sd}(\mathcal{X})}{E(\mathcal{X})} \to \frac{\sigma_{\mu}}{\xi_{\mu}} \qquad \text{as} \qquad J \to \infty,$$

and the standard deviation does not vanish in relative terms unless  $\sigma_{\mu} = 0$ .

d) The reserve  $q_{\epsilon}$  is the solution of the equation  $\Pr(\mathcal{X} > q_{\epsilon}) = \epsilon$  for a given small number  $\epsilon$ . Let  $\lambda = J\mu$  be the *fixed* Poisson parameter for the portfolio. Monte Carlo produces approximate solutions through the commands

$$\begin{array}{ll} 1 \ X^* \leftarrow 0 & 2 \ \mathrm{Draw} \ N^* \sim \mathrm{Poisson}(J\mu) \\ 3 \ \mathrm{Repeat} \ N^* \ \mathrm{times} & \\ 4 \ \mathrm{Draw} \ Z^* \ \mathrm{and} \ X^* \leftarrow X^* + Z^* \\ 5 \ \mathrm{Return} \ X^* \end{array}$$

Repeat this *m* times and sort the simulated portfolio claims in descening order as  $X_{(1)}^* \ge \ldots \ge X_{(m)}^*$  which yields  $X_{(m\epsilon)}^*$  as the approximate reserve. When  $\mu$  is random, add the initial command

Draw  $\mu^*$ 

before going through the others with  $\mu = \mu^*$ .

e) The first row where the variation is much larger corresponds to the random  $\mu$ . The 95% and 99% percent reserve under the two scenarios are 104.3 and 138.3 under the random  $\mu$  regime and 74.1 and 91.4 under the fixed  $\mu$  regime.

## Problem 2

a)

$$\pi^{\rm re} = J\mu \int_b^\infty (z-b)f(z)dz = J\mu \int_b^\infty \{1 - F(z)\}\,dz$$

after integration by parts.

**b**) $\pi^{\text{re}} = J\mu\xi^{\text{re}}$  where  $\xi^{\text{re}} = \beta(1+b/\beta)^{-\alpha-1}/(\alpha-1)$ . **c**)  $\mathcal{X}^{\text{re}} = \min(Z_1, b) + \ldots + \min(Z_N, b)$  and replace row 4 in 1d) by Draw  $Z^*$  and  $X^* \leftarrow X^* + \min(Z^*, b)$ 

d) The pure re-insurance premium balances re-insurance payments on average, and the exceeding amount is an average loss.

e) A b between 5 and 10 is optimal.

**f**) Replace  $\beta$  by  $\beta(1+I)$ .

**g)** The premium with inflation becomes  $\pi^{\text{re}} = J\mu\xi^{\text{re}}$  where now  $\xi^{\text{re}} = \beta(1+I)(1+b/(1+I)\beta)^{-\alpha-1}/(\alpha-1)$ , and the denominator is more sensitively dependent on I when  $\alpha$  increases.

## Problem 3

a) By Bayes formula

$$f(\mu|n_1,\dots,n_K) = cf(n_1,\dots,n_K|\mu)f(\mu) = c'\mu^{n_1+\dots+n_K+\alpha-1}e^{-\mu(K+\alpha/\xi_\mu)}$$

after inserting for the density functions. This is a Gamma density function as stated. b) Obvious.

c)  $E(\bar{n}|\mu) = \mu$  which implies that  $E(\hat{\pi} - \pi|\mu) = 0$  and the result follows by the rule of double expectation.

**d**) The first part follows from the Gamma distribution identified in a). Now by the rule of double variance

$$\operatorname{var}(\hat{\pi} - \pi) = E\{\operatorname{var}(\hat{\pi} - \pi) | n_1, \dots, n_K\} + \operatorname{var}\{E(\hat{\pi} - \pi) | n_1, \dots, n_K)\}$$

or

$$\operatorname{var}(\hat{\pi} - \pi) = E\{\xi_z^2 \hat{\mu}^2 / (\alpha + K\bar{n})\} + 0$$

which leads to the result given after inserting for  $\hat{\mu}$ .

e) When K is infinite, everything about the client is known and if  $\alpha$  is infinite, there is no variation among customers. In either case estimates of  $\pi$  haven't any uncertainty. If  $\alpha = 0$ , the prior has infinite variance and is valueless so that  $\hat{\pi} = \xi_z \bar{n}$  with variance as stated.