

# UNIVERSITY OF OSLO

## Faculty of Mathematics and Natural Sciences

Examination in: STK4050/9050 — Statistical simulations and computation.

Day of examination: Thursday December 8th 2011.

Examination hours: 09.00–13.00.

This problem set consists of 3 pages.

Appendices: None.

Permitted aids: Approved calculator.

Please make sure that your copy of the problem set is complete before you attempt to answer anything.

### Problem 1

- (a) For a Weibull distributed random variable  $T > 0$  we may write the survival probability

$$P(T > t) = \exp(-\theta t^\alpha)$$

where  $\theta > 0$  and  $\alpha > 0$  are parameters. Derive the inverse transform method for simulating Weibull distributed random variables.

- (b) Assume that  $Y$  is a gamma-distributed variable with shape parameter equal to 2 and a rate parameter  $\lambda$ , so that  $Y$  has density  $g(y) = \lambda^2 y \exp(-\lambda y)$ .

Explain why the inverse transform is more difficult in this case.

Present two approaches to simulating random variables from this distribution.

- (c) Assume that  $(T, \theta)$  has simultaneous density

$$g(t, \theta) = \theta \alpha t^{\alpha-1} \exp(-\theta t^\alpha) \lambda \exp(-\lambda \theta)$$

Explain why we then have that  $T$  conditional on  $\theta$  is Weibull with parameters  $\alpha$  and  $\theta$  as parameterized in question (a).

What is the marginal (unconditional) distribution of  $\theta$ ?

- (d) Derive the conditional distribution of  $\theta$  in question (c) given that the random variable  $T = t$ .

Present the Gibbs sampling algorithm for sampling from the bivariate distribution  $g(t, \theta)$  in question (c).

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## Problem 2

Assume that we want to evaluate an integral (expectation)  $l = \int H(x)f(x)dx = E_f[H(X)]$  where  $H(x)$  is some function, and  $X$  a random variable with density  $f(x)$ .

- (a) Argue that if  $g(x)$  is a density such that  $g(x) > 0$  if  $H(x)f(x) > 0$  then we may estimate  $l$  by sampling  $Y_i, i = 1, \dots, n$  independently from  $g(x)$  and calculating the average  $\hat{l} = \frac{1}{n} \sum_{i=1}^n H(Y_i)f(Y_i)/g(Y_i)$ .

State an expression for the variance of  $\hat{l}$ .

- (b) Assume that  $H(x) \geq 0$ . Argue that the optimal density  $g(x)$  to sample from satisfies  $g(x) \propto H(x)f(x)$ .

Discuss also why this result is not directly useful in practice, but may still give suggestions for sampling distributions  $g(x)$ .

Furthermore, extend the optimality result to functions  $H(x)$  not necessarily non-negative.

- (c) Assume that  $X$  has density  $f(x) = C \exp(-x^{1.5})$  for  $x > 0$  and  $f(x) = 0$  for  $x < 0$  (where approximately  $C = 1.1077325$ ). We want to estimate the probability  $X > 5$ . Why is it not a good idea to simply sample  $n$  random variables  $X_i$  from  $f(x)$  and calculating the proportion  $> 5$ ?

Importance sampling like in question (a) is a better idea. Suggest a sampling density  $g(x)$  that may give a quite precise estimate.

- (d) We may instead be interested in calculating  $E_f[X^2]$ . Suggest a sampling density  $g(x)$  suitable for this problem.

Why would a good solution to question (c) be a poor choice in this case?

## Problem 3

Assume that  $X_0, X_1, X_2, \dots$  is a time-homogeneous Markov chain on a finite state space  $\{1, 2, \dots, m\}$  with transition probabilities  $p_{ij} = P(X_n = j | X_{n-1} = i)$ . Assume also that the Markov chain is aperiodic and irreducible.

- (a) Assume that for some values  $\pi_1, \pi_2, \dots, \pi_m$  the Markov chain satisfies the detailed balance equations

$$\pi_i p_{ij} = \pi_j p_{ji}$$

Explain why the  $\pi_i$  defines the stationary distribution for  $X_n$  and why it will also be the limit distribution for  $X_n$  when  $n \rightarrow \infty$ .

The equation  $\pi_i p_{ij} = \pi_j p_{ji}$  is also referred to as reversibility. Give an explanation for this term.

- (b) The Metropolis-Hastings algorithm generates Markov chains given as follows: From current state of the chain  $X_n = i$

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- (i) Sample a proposal value  $Y = j$  from a conditional distribution  $q(y|X_n = i)$
- (ii) Accept  $Y$ , i.e. let  $X_{n+1} = j$  with probability

$$\alpha_{ij} = \min\left(1, \frac{\pi_j q(i|j)}{\pi_i q(j|i)}\right)$$

- (iii) Reject  $Y$ , i.e. let  $X_{n+1} = i$  with probability  $1 - \alpha_{ij}$ .

Show that the Metropolis-Hastings algorithm satisfies the detailed balance equation and conclude about how one may sample from the distribution given by the  $\pi_i$  from the algorithm.

Discuss properties that the proposal distribution should have for the sampling from the stationary distribution to be efficient.

END