# **Lecture note on the interpretation of regression coefficients**

# **1) The effect of** *X* **in the simple linear regression model**

To fix ideas, let *Y* = consumption of a certain class of goods and *X* = income, for a randomly chosen individual from the population.  $(X, Y)$  is jointly distributed with a joint pdf,  $f(x, y)$ ( the population distribution). The dependence of the response, *Y*, on the explanatory variable, *X*, is usually studied by means of the conditional distribution of *Y* for fixed values of *X* (i.e.,  $X = x$ ), with pdf,  $f(y|x) = f(x, y)/f_1(x)$ , where the marginal pdf for *X*, is

$$
f_1(x) = \int_{-\infty}^{\infty} f(x, y) dy
$$
. The regression function is simply the expected value of  $(Y | X = x)$  in

*f* (*y* | *x*), i.e.,  $\mu(x) = E(Y | x) = \int_0^{\infty} y f(y | x) dy$  $= E(Y | x) = \int_{-\infty}^{\infty} y f(y | x) dy$ , which expresses the expected response for a

given fixed value,  $X = x$ . In the simple linear regression model we postulate that  $\mu(x)$  is a linear function

$$
(1) \qquad \mu(x) = E(Y \mid x) = \alpha + \beta x
$$

In this model the regression coefficient,  $\beta$ , can be interpreted as the effect of a unit change of *X* (i.e.,  $(X = x) \rightarrow (X = x + 1)$ ) on the expected change of the response, *Y*.

**Elaboration.** Let  $Y_1$  be the consumption for a randomly chosen individual with income,  $X = x$ , and  $Y_2$  correspondingly for a randomly chosen individual with  $X = x + 1$ . Then the pdf's of  $Y_1$ , and  $Y_2$  are  $f(y|x)$  and  $f(y|x+1)$  with expected values,  $\mu(x)$  and  $\mu(x+1)$ , respectively. The expected difference becomes  $\beta$  since values,  $\mu(x)$  and  $\mu(x+1)$ , respectively. The expected difference becomes  $\beta$  sin<br> $E(Y_2 - Y_1) = E(Y_2) - E(Y_1) = \mu(x+1) - \mu(x) = \alpha + \beta(x+1) - \alpha - \beta x = \beta$ 

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$$

Note how the interpretation of  $\beta$  is derived from the meaning of the function,  $\mu(x) = E(Y | x)$ .

Note also that this interpretation *does not* apply to a single individual. It does not say anything about the expected response when a *single* individual increases the income from  $X = x$  to  $X = x + 1$ . For getting information on such effects we will need at least two observations of *X* and *Y* for each individual at two different points in time (i.e., panel data).

If we want the effect of 10 (say) units change in *X*, the same calculation gives

$$
\mu(x+10)-\mu(x)=10\beta.
$$

### **2) The effect of** *X***, controlling for** *Z* **(wealth)**

Now consider *Z* (e.g., wealth) as an additional explanatory variable that may influence *Y*. We want to find the effect on the expected response of a unit change of *X* - *controlling for Z*. Suppose the postulated regression function is

(2) 
$$
\mu(x, z) = E(Y | x, z) = \alpha + \beta x + \gamma z
$$

which is the expectation in the conditional distribution of *Y* for fixed values of  $X = x$  and  $Z = \overline{z}$ , with pdf,

$$
f(y \mid x, z) = \frac{f(x, y, z)}{f_1(x, z)}
$$
, where the marginal pdf of  $(X, Z)$  is  $f_1(x, z) = \int_{-\infty}^{\infty} f(x, y, z) dy$ ,

and where  $f(x, y, z)$  is the joint pdf of  $(X, Y, Z)$ .

We are now interested in the expected difference between two rv's,  $Y_1, Y_2$  (as in the elaboration under 1), where

- *Y*<sub>1</sub> is the consumption for a randomly chosen individual with income,  $X = x$  and Z=z (i.e.,  $Y_1 = (Y | X = x, Z = z)$ )
- $Y_2$  is the consumption for a randomly chosen individual with income,  $X = x + 1$  and  $Z=z$  (i.e.,  $Y_2 = (Y | X = x+1, Z = z)$ ).

Notice that  $Y_1$  and  $Y_2$  both have the same value, *z*, of *Z* (which is what we mean by "controlling for *Z*"). This is, of course, to make the comparison between  $Y_1$  and  $Y_2$  more fair.<br>
Then, the expected difference becomes<br>
(3)  $E(Y_2 - Y_1) = E(Y_2) - E(Y_1) = \mu(x+1, z) - \mu(x, z) = \alpha + \beta(x+1) + \gamma z - \alpha - \beta x - \gamma z = \beta$ Then, the expected difference becomes

(3) 
$$
E(Y_2 - Y_1) = E(Y_2) - E(Y_1) = \mu(x+1, z) - \mu(x, z) = \alpha + \beta(x+1) + \gamma z - \alpha - \beta x - \gamma z = \beta
$$

Thus,  $\beta$  can be interpreted as the expected change in the response (*Y*) between two subpopulations of individuals where all individuals in the first subpopulation have  $X = x$  and all individuals in the other have  $X = x + 1$ , and where all individuals in both groups *have the same value* of the wealth ( $Z = z$ ). This is often expressed by saying that  $\beta$  is "the effect of a unit change of *X* on (expected) *Y*, *ceteris paribus* – which translates to "everything else equal." Alternatively,  $\beta$  is sometimes called "*the partial effect* of a unit change in X (controlling for other explanatory variables)".

An advantage with this particular model is that the cet. par. effect of *X* reduces to a single parameter  $(\beta)$  no matter what the wealth  $(Z)$  is.

[Notice, in passing, that if the values of *Z* were different for  $Y_1$  and  $Y_2$ , e.g.,  $Z = z_1$  for  $Y_1$  and  $Z = z_2$  for  $Y_2$ , the calculation in (3) gives,

3<br>  $E(Y_2 - Y_1) = E(Y_2) - E(Y_1) = \mu(x+1, z_2) - \mu(x, z_1) = \beta + \gamma(z_2 - z_1)$ , which , which shows that the effect of a unit change in  $X$  - in that case - is partly due to differences in the wealth  $(if \gamma \neq 0, of course).]$ 

# **3) The effect of** *X***, controlling for** *Z* **(wealth) and** *V* **(age)**

Now we look at the conditional distribution of *Y* for fixed values,  $X = x, Z = z, V = v$ , with pdf 1  $(y | x, z, v) = \frac{f(x, y, z, v)}{f(x, v)}$  $\frac{\lambda}{(x, z, v)}$ *f*  $(y | x, z, v) = \frac{f(x, y, z, v)}{f(x, v)}$  $f(x, y, z, v)$ , where the marginal pdf of  $(X, Z, V)$  is  $f_1(x, z, v)$  $f_1(x, z, v) = \int_0^{\infty} f(x, y, z, v) dy$  $\mu(x, z, v) = \overline{E(Y | x, z, v)}$  $=\int f(x, y, z, v) dy$ . The expectation in this distribution is a function of *x*, *z*, and *v*,

The effect of a unit change in *X* (ceteris paribus), can be calculated as above

$$
\mu(x+1,z,v) - \mu(x,z,v)
$$

In the special case that we postulate a linear regression model,  $\mu(x, z, v) = \alpha + \beta x + \gamma z + \delta v$ , this calculation gives us bectal case that we postulate a linear regression model,  $\mu(x, z, v) = \alpha + \beta x + \gamma z + c$ <br>culation gives us<br> $\mu(x+1, z, v) - \mu(x, z, v) = \alpha + \beta(x+1) + \gamma z + \delta v - (\alpha + \beta x + \gamma z + \delta v) = \beta$ 

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$$

so the cet. par. effect of a unit change in *X* (sometimes also called "the income-effect on consumption") reduces to a single parameter,  $\beta$ , no matter what the wealth and age are.

#### **4) Modelling interaction between** *X* **(income) and** *Z* **(wealth)**

It is imaginable that the income-effect on consumption is different between rich and poor people. If this is the case, we say there is an *interaction* between income and wealth. An easy way to model this is to include the product term, *xz* (also called an interaction term), in the regression function

$$
E(Y \mid x, z) = \mu(x, z) = \alpha + \beta x + \gamma z + \partial x z
$$

1

Note that although this regression function is a non-linear function of *x* and *z*, we still call it *a linear regression model* since it is linear in the parameters,  $\alpha$ ,  $\beta$ ,  $\gamma$ , and  $\delta$ . Being linear in the parameters implies that it can be estimated by usual least squares techniques (e.g., OLS in the case of homoscedasticity, i.e., when we can postulate that  $var(Y | x, z) = constant)$ <sup>1</sup>.

The cet. par. effect of a unit change in *X* can be calculated as before

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\n
$$
\mu(x+1, z) - \mu(x, z) = \alpha + \beta(x+1) + \gamma z + \delta(x+1)z - (\alpha + \beta x + \gamma z + \delta x z) = \beta + \partial z
$$

<sup>&</sup>lt;sup>1</sup> If you wish to estimate this model by Stata (say), you need 4 variables, each with n observations, in the Stata data matrix: the response y, and 3 explanatory variables,  $x1 = x$ ,  $x2 = z$ , and  $x3 = xz$ . Then the following stata command for OLS does it: regr y x1 x2 x3

Hence (e.g.), if  $\delta$  < 0, the income-effect on consumption will be smaller for rich than for poor people.

If there are several explanatory variables, the price for including all sorts of interactions in the model is a large number of extra parameters in the regression function. For example, if we include *V*=age as an explanatory variable, a full interaction regression function could look like<br>  $\mu(x, z, v) = \alpha + \beta_1 x + \beta_2 z + \beta_3 v + \gamma_1 xz + \gamma_2 xv + \gamma_3 zv + \gamma_4 xzv$ 

$$
\mu(x, z, v) = \alpha + \beta_1 x + \beta_2 z + \beta_3 v + \gamma_1 xz + \gamma_2 xv + \gamma_3 zv + \gamma_4 xzv
$$

The interaction term,  $xzy$ , is called a  $2<sup>nd</sup>$  order interaction term. Check yourself that the cet. par. effect of a unit change of *X*, now becomes,

$$
\mu(x+1, z, v) - \mu(x, z, v) = \beta_1 + \gamma_1 z + \gamma_2 v + \gamma_4 z v
$$

### **5) The income effect on consumption may also depend on income**

Consider now, for simplicity, only *X* (income) as explanatory. Postulate a regression function

$$
E(Y \mid x) = \mu(x) = \alpha + \beta x + \gamma x^2
$$

(This is also a linear regression model that may be well estimated by OLS under the assumption of homoscedasticity since it is linear in the parameters,  $\alpha, \beta, \gamma$ .)<sup>2</sup>

The income-effect on consumption now becomes

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ome-effect on consumption now becomes  
\n
$$
\mu(x+1) - \mu(x) = \alpha + \beta(x+1) + \gamma(x+1)^2 - (\alpha + \beta x + \gamma x^2) = \beta + \gamma(2x+1)
$$

Hence (e.g.), if  $\gamma < 0$ , the income effect is smaller for high-income people than for lowincome people in this model.

<sup>&</sup>lt;sup>2</sup> To estimate this model in Stata (say), you need 3 variables (each with n observations) in the data matrix: the response y, and 2 explanatory variables,  $x1 = x$ ,  $x2 = x^2$ , and the OLS command becomes, regr y x1 x2