

# Dynamic Regression Models (Lect 15)

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- ▶ HGL: Ch 9; BN: Kap 10
- ▶ The HGL Ch 9 is a long chapter, and the testing for autocorrelation part we have already covered.
- ▶ HGL starts the chapter with the Finite Distributed lag model (DL), for example

$$Y_t = \beta_0 + \beta_1 X_t + \beta_2 X_{t-1} + \varepsilon_t \quad (1)$$

and discuss estimation/testing with classical assumptions for  $\varepsilon_t$ , and without

- ▶ But (1) is “almost” a usual static model, and because economic relationships are often more genuinely dynamic, it has low practical relevance.
- ▶ Therefore we focus the “ARDL” part of Ch 9, and starts with the simplest version of that model class

## Autoregressive first order model AR(1) model I

- ▶ The simplest “dynamic regression model”. It has properties that carry over to more general models (ARDL below).
- ▶ Assume that we have  $t = 1, 2, \dots, T$  independent and identically distributed random variables  $\varepsilon_t$ :

$$\varepsilon_t \sim IID(0, \sigma_\varepsilon^2), t = 1, 2, \dots, T$$

Then, from

$$Y_t = \beta_0 + \beta_1 Y_{t-1} + \varepsilon_t, \quad |\beta_1| < 1, \quad \varepsilon_t \sim IID(0, \sigma_\varepsilon^2), \quad (2)$$

we know something precise about the *conditional* distribution of  $Y_t$  given  $Y_{t-1}$ , and more generally *the history of  $Y$  up to period  $t - 1$* .

## Autoregressive first order model AR(1) model II

- ▶  $|\beta_1| < 1$  secures stationarity (HGL 9.1.3) for this model.
- ▶ We will refer to  $Y_t$  as given by (2) as a *1st order autoregressive process*, usually denoted AR(1).
- ▶ In direct parallel to the previous models we can write

$$Y_t = E(Y_t | Y_{t-1}) + \varepsilon_t = \beta_0 + \beta_1 Y_{t-1} + \varepsilon_t \quad (3)$$

where

$$E(\varepsilon_t Y_{t-1}) = 0 \quad (4)$$

by construction (in fact by assumption of  $|\beta_1| < 1$ , but leave that for another course)

- ▶ (4) is necessary for pre-determinedness of  $Y_{t-1}$ .
  - ▶ But is  $E(\varepsilon_{t+j} Y_{t-1}) = 0$  for  $j = 1, 2, \dots$  as well?
  - ▶ And what about  $E(\varepsilon_{t-1-j} Y_{t-1})$  for  $j = 1, 2$ ?

## Autoregressive first order model AR(1) model III

- ▶ To answer these questions: need to consider the solution of (2), which is a stochastic difference equation.

## Solution I

- ▶  $|\beta_1| < 1$  defines  $Y_t$  as a *causal-process*: Stochastic shocks/impulses/news represented by  $\varepsilon$  come before (or in the same period) as the response in  $Y_t$ .
- ▶ The backward-recursive solution of a causal-process is dynamically stable. We show in class that it is:

$$Y_t = \beta_0 \sum_{i=0}^{t-1} \beta_1^i + \beta_1^t Y_0 + \sum_{i=0}^{t-1} \beta_1^i \varepsilon_{t-i} \quad (5)$$

where  $Y_0$  is the *initial condition*.

The conditional expectation is

$$E(Y_t | Y_0) = \beta_0 \sum_{i=0}^{t-1} \beta_1^i + \beta_1^t Y_0$$

## Solution II

while the **unconditional expectation of**  $Y_t$  is defined for the situation where  $t \rightarrow \infty$ :

$$E(Y_t) = \frac{\beta_0}{1 - \beta_1} \quad (6)$$

For simplicity, we regard  $Y_0$  as a deterministic parameter. Then the variance is found as:

$$\begin{aligned} \text{Var}(Y_t) &= \text{Var}\left(\sum_{i=0}^{t-1} \beta_1^i \varepsilon_{t-i}\right) = \sigma_\varepsilon^2 \sum_{i=0}^{t-1} (\beta_1^2)^i \\ &\stackrel{t \rightarrow \infty}{=} \frac{\sigma_\varepsilon^2}{1 - \beta_1^2} \end{aligned} \quad (7)$$

## Pre-determinedness of lagged $Y$ I

The solution for  $Y_{t-1}$  (make use of (5)!) shows that:

$$E(Y_{t-1}\varepsilon_t) = E\left(\sum_{i=0}^{t-2} \beta_1^i \varepsilon_{t-i-1}\right)\varepsilon_t = 0$$

and

$$E(Y_{t-1}\varepsilon_{t+j}) = 0 \text{ for } j = 1, 2, \dots$$

But also that:

$$E(Y_{t-1}\varepsilon_{t-i}) \neq 0 \text{ for } i = 1, 2,$$

$Y_{t-1}$  is a pre-determined explanatory variable.



## Bias and consistency I

- ▶ To save notation: Consider the case of  $E(Y_t) = 0 \implies \beta_0 = 0$ .
- ▶ The OLS estimator  $\hat{\beta}_1$  is

$$\hat{\beta}_1 = \frac{\sum_{t=2}^T Y_t Y_{t-1}}{\sum_{t=2}^T Y_{t-1}^2} = \sum_{t=2}^T \left( \frac{\beta_1 Y_{t-1}^2}{\sum_{t=2}^T Y_{t-1}^2} \right) + \sum_{t=2}^T \left( \frac{Y_{t-1} \varepsilon_t}{\sum_{t=2}^T Y_{t-1}^2} \right) \quad (8)$$

$\implies$

$$E(\hat{\beta}_1 - \beta_1) = E\left(\frac{\sum_{t=2}^T Y_{t-1} \varepsilon_t}{\sum_{t=2}^T Y_{t-1}^2}\right)$$

- ▶ Cannot show that  $E$  of the bias term is zero

## Bias and consistency II

- ▶ Both the denominator and numerator are random variables, and they are not independent: For example will  $\varepsilon_2$  “be in” the numerator and (because of  $Y_2 = \varepsilon_2$ ) also in  $Y_2 \times Y_2$  in the denominator.
- ▶ But, with reference to the Law of large numbers and Slutsky's theorem we have

$$\text{plim} (\hat{\phi}_1 - \phi_1) = \frac{\text{plim} \frac{1}{T} \sum_{t=2}^T Y_{t-1} \varepsilon_t}{\text{plim} \frac{1}{T} \sum_{t=2}^T Y_{t-1}^2} = \frac{0}{\frac{\sigma_\varepsilon^2}{1-\beta_1^2}} = 0.$$

since  $E(Y_{t-1}\varepsilon_t) = 0$  (numerator) and  $|\beta_1| < 1$  (implies the existence of the variance).

## Bias and consistency III

- ▶ The OLS estimator  $\hat{\beta}_1$  in the AR(1) is consistent, and it can be shown to be asymptotically normal:

$$\sqrt{T} \left( \hat{\beta}_1 - \beta_1 \right) \xrightarrow{d} N \left( 0, (1 - \beta_1^2) \right) \quad (9)$$

which entails that *t-ratios* can be compared with critical values from the normal distribution.

- ▶ Therefore: the large sample inference theory for the regression model extends to the AR(1) model.

## Analysis of finite sample bias in AR(1)

In (2), the finite sample bias can be shown to be approximately

$$E\left(\hat{\beta}_1 - \beta_1\right) \approx \frac{-2\beta_1}{T},$$

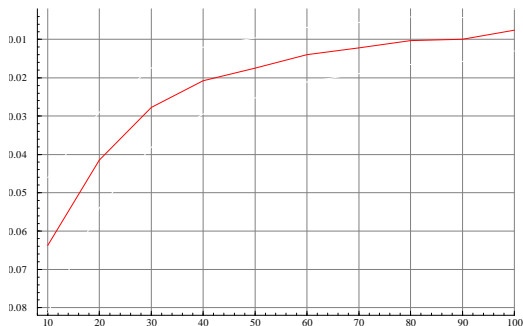
We can make this more concrete with a Monte-Carlo analysis. In the experiment, the DGP is

$$Y_t = 0.5Y_{t-1} + \varepsilon_{Yt}, \quad \varepsilon_{Yt} \sim NIID(0, 1),$$

and  $T = 10, 11, \dots, 99, 100$ . We use 1000 replications for each  $T$  and estimate the bias:

$$\hat{E}\left(\hat{\beta}_{1(T)} - \beta_1\right) = \frac{1}{1000} \sum_{i=1}^{1000} \left(\hat{\beta}_{1(T)i} - \beta_1\right).$$

## Bias in the AR(1) model



$$\begin{aligned} \hat{E} \left( \hat{\beta}_{1(10)} - 0.5 \right) &= \\ -0.058 &> \\ \dots &\approx \frac{-2 \times 0.5}{10} = -0.1 \end{aligned}$$

$$\begin{aligned} \hat{E} \left( \hat{\beta}_{1(100)} - 0.5 \right) &= \\ -0.008 &> \\ \dots &\approx \frac{-2 \times 0.5}{100} = \\ -0.01. & \end{aligned}$$

## Monte Carlo analysis of AR(1) with exogenous regressor

$$Y_t = \beta_0 + \beta_1 Y_{t-1} + \beta_2 X_t + \varepsilon_t, \quad |\beta_1| < 1, \quad \varepsilon_t \sim IID(0, \sigma_\varepsilon^2). \quad (10)$$

which we will also refer to as an AutoRegressive Distributed Lag model, ARDL.

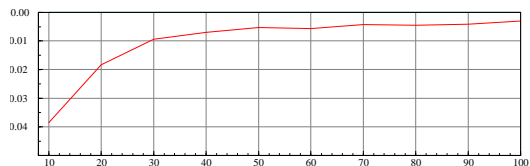
- ▶ We assume that  $X_t$  is strictly exogenous

Monte Carlo DGP:

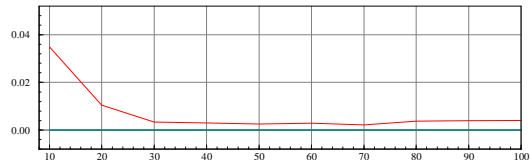
$$\begin{aligned} Y_t &= 0.5Y_{t-1} + 1 \cdot X_t + \varepsilon_{Yt}, & \varepsilon_{Yt} &\sim NIID(0, 1), \\ X_t &= 0.5X_{t-1} + \varepsilon_{Xt}, & \varepsilon_{Xt} &\sim NIID(0, 2), \end{aligned}$$

There are now two biases,  $\hat{E}(\hat{\beta}_{1(T)} - 0.5)$  and  $\hat{E}(\hat{\beta}_{2(T)} - 1)$

## Biases in the ADL model



$$\hat{E} \left( \hat{\beta}_1(\tau) - 0.5 \right)$$



$$\hat{E} \left( \hat{\beta}_2(\tau) - 1 \right)$$

## Conclusions

- ▶ The OLS biases are small, and the speeds of convergence to zero are high
- ▶ OLS estimation, and the use  $t$ -ratios and  $F$ -statistics for testing extend to dynamic models, *given that the model is correctly specified*, disturbances that have the usual classical assumptions conditional on  $Y_{t-1}$  and  $X_t$ .
- ▶ In particular: Avoid residual autocorrelation because it will destroy pre-determinedness of  $Y_{t-1}$ !
- ▶ The tests we have covered for Non-Normality, Heteroskedasticity and Autocorrelation in (Lect 13 and 14) are valid mis-specification tests also for ARDL models!



## Dynamic response to shocks

One purpose of estimating an ARDL model:

$$Y_t = \beta_0 + \beta_1 Y_{t-1} + \beta_2 X_t + \beta_3 X_{t-1} + \varepsilon_t \quad (11)$$

with classical assumptions for  $\varepsilon_t$  conditional on  $Y_{t-1}$ ,  $X_t$  and  $X_{t-1}$

is to estimate the dynamic response of  $Y$  to a permanent or temporary change in  $X$ .

- ▶ When we consider changes in the  $X$ , the key concept is **dynamic multiplier**.
- ▶ Can also study a temporary shock to  $\varepsilon$  (for example of magnitude one standard deviation  $\sigma$ ) These dynamic effects are often called *impulse-responses*.
- ▶ In class: Derive dynamic multipliers (short), and show examples of estimated dynamic multipliers.
- ▶ Use of model in forecasting: Lecture 16.