

# ECON3150/4150 Spring 2016

## Lecture 7

### Hypothesis testing

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# Overview

- Hypothesis testing in SRLM
- Hypothesis testing in MLRM

# Hypothesis testing

- In the revision part we used the sample average as an estimator for the unknown population mean
- Now we are dealing with unknown parameters ( $\beta_0$  and  $\beta_1$ ) and we have derived estimators for them.
- As with sample average we formulate a null hypothesis that  $\beta_1$  is equal to some specific value  $\beta_{1,0}$  and an alternative hypothesis.

# Hypothesis testing

Given the null and alternative hypothesis:

- Compute the test statistic
- Choose your desired significance level
  - Compare test statistic to critical test statistic
  - Compare significance level to the p-value of the test-statistic
- Two outcomes:
  - Reject the null in favor of an alternatives.
  - Fail to reject the null.

Finding the critical value and the p-value require knowing the distribution of the statistic. The distribution of the OLS estimators depend on the distribution of the error term.

## Compute test statistic

The test statistic for the regression coefficient is the t statistic

$$t = \frac{\text{estimator} - \text{hypothesised value}}{\text{standard error of the estimator}}$$

- Since the standard error is always positive the t-statistic has the same sign as the difference between the estimator and the hypothesized value.
- For a given standard error the larger value of the estimator the larger value of the t-statistic.
- If the null hypothesis is that the true parameter is zero, a large estimator provides evidence against the null.
- t-values sufficiently far from the hypothesized value result in rejection of the null.

## Compute t-statistic

In the single regression model:  $Y = \beta_0 + \beta_1 X + u$

- Compute the t-statistics:

$$t = \frac{\hat{\beta}_1 - \beta_{1,0}}{SE(\hat{\beta}_1)}$$

- If the null hypothesis is that  $\beta_1 = 1$

$$t = \frac{\hat{\beta}_1 - 1}{SE(\hat{\beta}_1)}$$

The standard error is given by:

$$SE(\hat{\beta}_1) = \sqrt{\hat{\sigma}_{\beta_1}^2}$$

# Hypothesis testing

Example where the standard error is reported below in parenthesis:

$$wage = \beta_0 + \beta_1 educ + u$$

$$wage = \underset{(0.104)}{0.284} + \underset{(0.007)}{0.092}educ$$

If  $H_0 : \beta_1 = 0$ :

$$t = 0.092/0.007 = 13.14$$

## Distribution of estimators

We can use the t-statistic only if the regression coefficient is normally distributed. Under the OLS assumptions including the normality assumption, sampling distribution of the OLS estimators is normal.

Because:

- A random variable which is a linear function of a normally distributed variable is itself normally distributed.
- If we assume that  $u \sim N(0, \sigma^2)$  then  $Y_i$  is normally distributed.
- Since the estimators  $\hat{\beta}_0$  and  $\hat{\beta}_1$  is linear functions of the  $Y_i$ 's then the estimators are normally distributed.
- Thus  $\hat{\beta}_1 \sim N[\beta_1, \text{Var}(\hat{\beta}_1)]$



## Normality assumption

- In large samples we can invoke the CLT to conclude that the OLS satisfy asymptotic normality.
- Whenever  $Y$  takes on just a few values and we have few observations it cannot have anything close to a normal distribution.
- If the  $\hat{\beta}_1$  is not normally distributed the t-statistic does not have t distribution.
- The normal distribution of  $u$  is the same as the distribution of  $Y$  given  $X$ .

## Distribution of estimators

- If  $\hat{\beta}_1$  is either normally distributed or approximately normally distributed it can be standardized.

$$t = \frac{\hat{\beta}_j - \beta_j}{\text{se}(\hat{\beta}_j)}$$

- t statistic for  $\hat{\beta}_1$  is t-distributed
- Thus:  $(\hat{\beta}_1 - \beta_1)/\text{se}(\hat{\beta}_1) \sim t_{n-k-1}$
- $t_{n-k-1}$  represents the "t-distribution" with  $n - k - 1$  degrees of freedom.
- As the degrees of freedom in the t-distribution gets large, the t distribution approaches the standard normal distribution.

# Hypothesis testing in MLRM

Rather than testing  $\beta_1$  you can test any  $\beta_j$

- If  $H_0 : \beta_j = \beta_{j,0}$

$$t = \frac{\hat{\beta}_j - \beta_{j,0}}{\text{se}(\hat{\beta}_j)}$$

## Hypothesized value

- In practice it is unusual to have a feeling for the actual value of the coefficients.
- Very often the objective of the analysis is to demonstrate that  $Y$  is influenced by  $X$ , without having any specific prior notion of the actual coefficients of the relationships.
- In this case it is usual to define  $\beta_j = 0$  as the null, which means that the null hypothesis is that  $X$  has no partial effect on  $Y$ .
- If  $H_0 : \beta_j = 0$  then the t-statistic is just the coefficient divided by the standard error.

$$t = \frac{\hat{\beta}_j}{se(\hat{\beta}_j)}$$

- This statistic is automatically printed as part of the regression results.

# Example

```
1 . reg wage educ, robust
```

Linear regression

```
Number of obs =      935
F( 1, 933) =      95.65
Prob > F      =      0.0000
R-squared     =      0.1070
Root MSE     =      382.32
```

wage	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
educ	60.21428	6.156956	9.78	0.000	48.1312	72.29737
_cons	146.9524	80.26953	1.83	0.067	-10.57731	304.4822

Is the t-statistic for the coefficient of education large? What about the statistic of the intercept?

# The t-statistic

- The further the estimated coefficient is from the hypothesized value the more likely the hypothesis is incorrect.
- If  $H_0 : \beta_j = 0$  then if  $H_0$  is true  $\beta_j$  should be close to 0

$$\text{small} \left| \frac{\hat{\beta}_j}{\text{se}(\hat{\beta}_j)} \right|$$

- If  $H_0$  is false  $\hat{\beta}_j$  should be far away from 0.

$$\text{large} \left| \frac{\hat{\beta}_j}{\text{se}(\hat{\beta}_j)} \right|$$

But what is small and what is large?

# Critical value

The critical value of the t-distribution determines what is small and large.

- The critical value depends on your desired level of accuracy.
- The accuracy is determined by the significance level - the probability of rejecting  $H_0$  when it is in fact true. (type I error)
- A 5% significance level means that we mistakenly reject  $H_0$  5% of the time.
- The critical value increases as significance level falls, thus a null hypothesis that is rejected at a 5% level is automatically rejected at the 10% level as well.
- The lower the possible significance level the lower is the risk that we mistakenly reject the null.

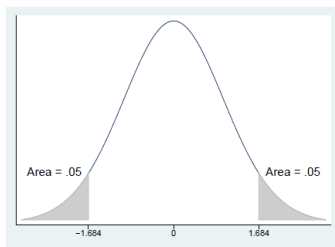
## Make conclusion

- When  $H_0 : \beta_j = 0$  is rejected at the 5% level we usually say
  - $X_j$  is statistically significant at the 5% level
  - $X_j$  is statistically different from zero at the 5% level.
- If  $H_0$  is not rejected we say that  $X$  is statistically insignificant at the 5% level.
- If we fail to reject  $H_0$  we never say that we accept  $H_0$  because there are many other values for  $\beta_1$  which cannot be rejected and they cannot all be true.



## Make conclusions

We could be wrong as we do not measure  $\beta_1$ . For a 10% significance level we falsely reject the null 10% of the time.



The graphs illustrates  $f(\hat{\beta}_1/(se(\hat{\beta}_1)))$ . 10% of the distribution is in the grey area.

# Finding critical values

**TABLE 2** Critical Values for Two-Sided and One-Sided Tests Using the Student t Distribution

Degrees of Freedom	Significance Level				
	20% (2-Sided)	10% (2-Sided)	5% (2-Sided)	2% (2-Sided)	1% (2-Sided)
	10% (1-Sided)	5% (1-Sided)	2.5% (1-Sided)	1% (1-Sided)	0.5% (1-Sided)
1	3.08	6.31	12.71	31.82	63.66
2	1.89	2.92	4.30	6.96	9.92
3	1.64	2.35	3.18	4.54	5.84
4	1.53	2.13	2.78	3.75	4.60
5	1.48	2.02	2.57	3.36	4.03
6	1.44	1.94	2.45	3.14	3.71
7	1.41	1.89	2.36	3.00	3.50
8	1.40	1.86	2.31	2.90	3.36
9	1.38	1.83	2.26	2.82	3.25
10	1.37	1.81	2.23	2.76	3.17
11	1.36	1.80	2.20	2.72	3.11
12	1.36	1.78	2.18	2.68	3.05
13	1.35	1.77	2.16	2.65	3.01
14	1.35	1.76	2.14	2.62	2.98
15	1.34	1.75	2.13	2.60	2.95
16	1.34	1.75	2.12	2.58	2.92
17	1.33	1.74	2.11	2.57	2.90
18	1.33	1.73	2.10	2.55	2.88
19	1.33	1.73	2.09	2.54	2.86
20	1.33	1.72	2.09	2.53	2.85
21	1.32	1.72	2.08	2.52	2.83
22	1.32	1.72	2.07	2.51	2.82
23	1.32	1.71	2.07	2.50	2.81
24	1.32	1.71	2.06	2.49	2.80
25	1.32	1.71	2.06	2.49	2.79
26	1.32	1.71	2.06	2.48	2.78
27	1.31	1.70	2.05	2.47	2.77
28	1.31	1.70	2.05	2.47	2.76
29	1.31	1.70	2.05	2.46	2.76
30	1.31	1.70	2.04	2.46	2.75
60	1.30	1.67	2.00	2.39	2.66
90	1.29	1.66	1.99	2.37	2.63
120	1.29	1.66	1.98	2.36	2.62
$\infty$	1.28	1.64	1.96	2.33	2.58

Values are shown for the critical values for two-sided ( $\neq$ ) and one-sided ( $>$ ) alternative hypotheses. The critical value for the

# P-value

- An alternative to comparing the t-statistic to the critical t-statistic is to compute the p-value for the t-statistic.
- The p-value is more informative as it gives you the smallest significance level at which the null hypothesis would have been rejected.
- A null that is rejected at a 5% level must have a p-value smaller than 5%.

# P-value

- We remember that the p-value is the probability of obtaining the observed t statistic (or one more extreme) as a matter of chance if the null hypothesis  $H_0 : \beta_j = 0$  is true.
- It therefore gives the lowest significance level at which the null hypothesis could be rejected.
- It is not a direct probability of rejecting  $H_0$  when it is actually true.

# P-value stata

```
1 . reg wage educ, robust
```

Linear regression

```
Number of obs =      935
F( 1, 933) =      95.65
Prob > F      =      0.0000
R-squared     =      0.1070
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## Computing p-values for t-tests

- The p-value (in SLRM) is calculated by computing the probability that a t random variable with  $(n-2)$  degrees of freedom is larger than  $t^{act}$  in absolute value.
- Thus the p-value is the significance level of the test when we use the value of the test statistic as the critical value for the test.

For the two sided test:

$$\begin{aligned} \text{p-value} &= Pr_{H_0}(|t| > |t^{act}|) = 2P(t > t^{act}) \\ &= Pr(|Z| > |t^{act}|) = 2\phi(-|t^{act}|) \text{ in large samples} \end{aligned}$$

# Finding p-value

Standard Normal Probabilities



Table entry for  $z$  is the area under the standard normal curve to the left of  $z$ .

$z$	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
-3.4	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0002
-3.3	.0005	.0005	.0005	.0004	.0004	.0004	.0004	.0004	.0004	.0003
-3.2	.0007	.0007	.0006	.0006	.0006	.0006	.0006	.0005	.0005	.0005
-3.1	.0010	.0009	.0009	.0009	.0008	.0008	.0008	.0008	.0007	.0007
-3.0	.0013	.0013	.0013	.0012	.0012	.0011	.0011	.0011	.0010	.0010
-2.9	.0019	.0018	.0018	.0017	.0016	.0016	.0015	.0015	.0014	.0014
-2.8	.0026	.0025	.0024	.0023	.0023	.0022	.0021	.0021	.0020	.0019
-2.7	.0035	.0034	.0033	.0032	.0031	.0030	.0029	.0028	.0027	.0026
-2.6	.0047	.0045	.0044	.0043	.0041	.0040	.0039	.0038	.0037	.0036
-2.5	.0062	.0060	.0059	.0057	.0055	.0054	.0052	.0051	.0049	.0048
-2.4	.0082	.0080	.0078	.0075	.0073	.0071	.0069	.0068	.0066	.0064
-2.3	.0107	.0104	.0102	.0099	.0096	.0094	.0091	.0089	.0087	.0084
-2.2	.0139	.0136	.0132	.0129	.0125	.0122	.0119	.0116	.0113	.0110
-2.1	.0179	.0174	.0170	.0166	.0162	.0158	.0154	.0150	.0146	.0143
-2.0	.0228	.0222	.0217	.0212	.0207	.0202	.0197	.0192	.0188	.0183
-1.9	.0287	.0281	.0274	.0268	.0262	.0256	.0250	.0244	.0239	.0233
-1.8	.0359	.0351	.0344	.0336	.0329	.0322	.0314	.0307	.0301	.0294
-1.7	.0446	.0436	.0427	.0418	.0409	.0401	.0392	.0384	.0375	.0367
-1.6	.0548	.0537	.0526	.0516	.0505	.0495	.0485	.0475	.0465	.0455
-1.5	.0668	.0655	.0643	.0630	.0618	.0606	.0594	.0582	.0571	.0559
-1.4	.0808	.0793	.0778	.0764	.0749	.0735	.0721	.0708	.0694	.0681
-1.3	.0968	.0951	.0934	.0918	.0901	.0885	.0869	.0853	.0838	.0823
-1.2	.1151	.1131	.1112	.1093	.1075	.1056	.1038	.1020	.1003	.0985
-1.1	.1357	.1335	.1314	.1292	.1271	.1251	.1230	.1210	.1190	.1170
-1.0	.1587	.1562	.1539	.1515	.1492	.1469	.1446	.1423	.1401	.1379
-0.9	.1841	.1814	.1788	.1762	.1736	.1711	.1685	.1660	.1635	.1611
-0.8	.2119	.2090	.2061	.2033	.2005	.1977	.1949	.1922	.1894	.1867
-0.7	.2420	.2389	.2358	.2327	.2296	.2266	.2236	.2206	.2177	.2148
-0.6	.2743	.2709	.2676	.2643	.2611	.2578	.2546	.2514	.2483	.2451
-0.5	.3085	.3050	.3015	.2981	.2946	.2912	.2877	.2843	.2810	.2776
-0.4	.3446	.3409	.3372	.3336	.3300	.3264	.3228	.3192	.3156	.3121
-0.3	.3821	.3783	.3745	.3707	.3669	.3632	.3594	.3557	.3520	.3483
-0.2	.4207	.4168	.4129	.4090	.4052	.4013	.3974	.3936	.3897	.3859
-0.1	.4602	.4562	.4522	.4483	.4443	.4404	.4364	.4325	.4286	.4247
-0.0	.5000	.4960	.4920	.4880	.4840	.4801	.4761	.4721	.4681	.4641

# Finding p-value

From the example last week:

```
1 . reg wage educ, robust
```

Linear regression

```
Number of obs =      935
F( 1, 933) =      95.65
Prob > F      =      0.0000
R-squared     =      0.1070
Root MSE     =      382.32
```

wage	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
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_cons	146.9524	80.26953	1.83	0.067	-10.57731	304.4822

The constant has a computed t value of 1.83. Since n is large we can use the z-table. The p-value is  $2\phi(-1.83)$ .



# Finding p-value

Standard Normal Probabilities



Table entry for  $z$  is the area under the standard normal curve to the left of  $z$ .

$z$	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
-3.4	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0002
-3.3	.0005	.0005	.0005	.0004	.0004	.0004	.0004	.0004	.0004	.0003
-3.2	.0007	.0007	.0006	.0006	.0006	.0006	.0006	.0006	.0005	.0005
-3.1	.0010	.0009	.0009	.0009	.0008	.0008	.0008	.0008	.0007	.0007
-3.0	.0013	.0013	.0013	.0012	.0012	.0011	.0011	.0011	.0010	.0010
-2.9	.0019	.0018	.0018	.0017	.0016	.0016	.0015	.0015	.0014	.0014
-2.8	.0026	.0025	.0024	.0023	.0023	.0022	.0021	.0021	.0020	.0019
-2.7	.0035	.0034	.0033	.0032	.0031	.0030	.0029	.0028	.0027	.0026
-2.6	.0047	.0045	.0044	.0043	.0041	.0040	.0039	.0038	.0037	.0036
-2.5	.0062	.0060	.0059	.0057	.0055	.0054	.0052	.0051	.0049	.0048
-2.4	.0082	.0080	.0078	.0075	.0073	.0071	.0069	.0068	.0066	.0064
-2.3	.0107	.0104	.0102	.0099	.0096	.0094	.0091	.0089	.0087	.0084
-2.2	.0139	.0136	.0132	.0129	.0125	.0122	.0119	.0116	.0113	.0110
-2.1	.0179	.0174	.0170	.0166	.0162	.0158	.0154	.0150	.0146	.0143
-2.0	.0228	.0222	.0217	.0212	.0207	.0202	.0197	.0192	.0188	.0183
-1.9	.0287	.0281	.0274	.0268	.0262	.0256	.0250	.0244	.0239	.0233
-1.8	.0359	.0351	.0344	.0336	.0329	.0322	.0314	.0307	.0301	.0294
-1.7	.0446	.0436	.0427	.0418	.0409	.0401	.0392	.0384	.0375	.0367
-1.6	.0548	.0537	.0526	.0516	.0505	.0495	.0485	.0475	.0465	.0455
-1.5	.0668	.0655	.0643	.0630	.0618	.0606	.0594	.0582	.0571	.0559
-1.4	.0808	.0793	.0778	.0764	.0749	.0735	.0721	.0708	.0694	.0681
-1.3	.0968	.0951	.0934	.0918	.0901	.0885	.0869	.0853	.0838	.0823
-1.2	.1151	.1131	.1112	.1093	.1075	.1056	.1038	.1020	.1003	.0985
-1.1	.1357	.1335	.1314	.1292	.1271	.1251	.1230	.1210	.1190	.1170
-1.0	.1587	.1562	.1539	.1515	.1492	.1469	.1446	.1423	.1401	.1379
-0.9	.1841	.1814	.1788	.1762	.1736	.1711	.1685	.1660	.1635	.1611
-0.8	.2119	.2090	.2061	.2033	.2005	.1977	.1949	.1922	.1894	.1867
-0.7	.2420	.2389	.2358	.2327	.2296	.2266	.2236	.2206	.2177	.2148
-0.6	.2743	.2709	.2676	.2643	.2611	.2578	.2546	.2514	.2483	.2451
-0.5	.3085	.3050	.3015	.2981	.2946	.2912	.2877	.2843	.2810	.2776
-0.4	.3446	.3409	.3372	.3336	.3300	.3264	.3228	.3192	.3156	.3121
-0.3	.3821	.3783	.3745	.3707	.3669	.3632	.3594	.3557	.3520	.3483
-0.2	.4207	.4168	.4129	.4090	.4052	.4013	.3974	.3936	.3897	.3859
-0.1	.4602	.4562	.4522	.4483	.4443	.4404	.4364	.4325	.4286	.4247
-0.0	.5000	.4960	.4920	.4880	.4840	.4801	.4761	.4721	.4681	.4641

$$p = 2 * 0.036$$

$$= 0.0672$$

# The Z-table

Remember that:

$$\phi(-Z) = 1 - \phi(Z)$$

Thus the table does not have to show the probability distribution for both positive and negative  $Z$ .

# Interpreting p-values

- Small p-values are evidence against the null, large p-values provide little evidence against the null.
- If for example the p-value = 0.5 then we would observe a value of the t statistic as extreme as we did in 50% of all random samples if the null hypothesis is true.
- If  $\alpha$  denotes the significance level then  $H_0$  is rejected if p-value  $< \alpha$ .

# Interpreting p-values

## p-value

**Correct interpretation:** Assuming that the null is true you would obtain the observed difference or more in  $p\%$  of studies due to random sampling error.

**Wrong interpretation:** P-value is the probability of making a mistake by rejecting a true null hypothesis.

- The p-value is calculated based on the assumption that the null is true for the population, thus it cannot tell you the probability that the null is true or false.
- A low p-value indicates that your data are unlikely assuming a true null, but it cannot evaluate whether it is more likely that the low p-value comes from the null being true but having an unlikely sample or the null being false.

# Collective illusions

Think that you have made a t-test and get a p-value of 0.01

- You have *not* absolutely disproved the null hypothesis.
- You have *not* found the probability of the null hypothesis being true.
- You *do not* know if you decide to reject the null hypothesis the probability that you are making the wrong decision.

## An example

The MEAP93 data contains observations on 408 school districts on average teacher salary in thousands of dollars (*sal*) and the percentage of students passing the MEAP math. A regression gives the following output:

$$\widehat{math10} = \underset{(3.22)}{8.28} + \underset{(0.10)}{0.498}sal$$

- The constant term is 8.28 with a standard error of 3.22
- The slope parameter is 0.498 with a standard error of 0.1

## An example

$$\widehat{math10} = \underset{(3.22)}{8.28} + \underset{(0.10)}{0.498}sal$$

$$t = \frac{0.498 - 0}{0.10} = 4.98$$

$$p\text{-value} = 2\phi(-4.98) < 0.00001$$

- The 5% critical value is given by:  $t_{406}^c = 1.96$
- We can reject the null that salary does not affect the percentage of students passing the math10.

# An example

The stata output for the same regression shows the same conclusion.

```
1 . reg math10 sal
```

Source	SS	df	MS	Number of obs = 408		
Model	2562.57022	1	2562.57022	F( 1, 406) =	24.62	
Residual	42254.6103	406	104.075395	Prob > F =	0.0000	
Total	44817.1805	407	110.115923	R-squared =	0.0572	
				Adj R-squared =	0.0549	
				Root MSE =	10.202	

math10	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
sal	.4980309	.1003674	4.96	0.000	.3007264	.6953355
_cons	8.282175	3.228869	2.57	0.011	1.934787	14.62956



## One-sided vs two-sided test

- Two-sided tests, as described so far, are appropriate when we have no information about the alternative hypothesis
- When we are in a position to say that if the null hypothesis is not true the coefficient cannot be lower (greater) than that specified by it we can do a one sided test.
- Note: there is a trade-off between the significance level and the power of the test that comes into play when choosing a one-sided test.

## T-statistic one sided test

We remember that for the two sided test it was the absolute value of the t-statistic that was important. For the one sided test the sign of the t-statistic is also important.

Rejection rules for one sided tests:

- One sided:  $H_1 : \beta_1 > 0 : t^{act} > t^c$
- One sided:  $H_1 : \beta_1 < 0 : t^{act} < -t^c$

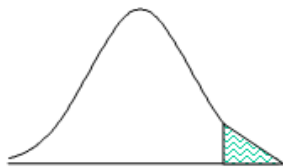
Note: The degrees of freedom of the test statistic is given by  $n - k - 1$  where  $k$  is the number of independent variables.

# One sided p-value

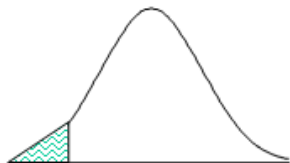
- One sided test:  $H_1 : \beta_1 > 0$
- If  $\hat{\beta}_1 < 0$  we know that the p-value is greater than 0.5 and there is no need to calculate it.
- If  $\hat{\beta}_1 > 0$  then  $t > 0$  and the p-value is half of the two-sided p-value.
- Since the t-distribution is symmetric around zero the reversed applied to the one sided test that  $\beta_1 < 0$

$$p - value = Pr_{H_0}(t < t^{act}) = Pr_{H_0}(t > |t^{act}|)$$

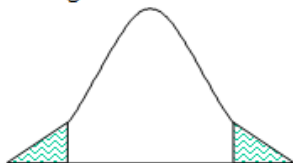
## One-sided vs two-sided test



Positive one-tailed test

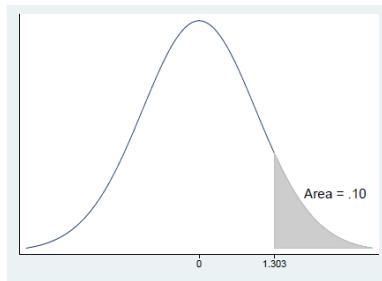
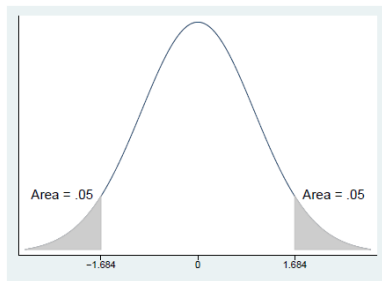


Negative one-tailed test



Two-tailed test

## One-sided vs two-sided test



An illustration of the difference between two-sided and one sided test with 40 degrees of freedom.

## Example one-sided test

Consider the following regression:

$$\hat{price} = 1.21 + 0.82 wage$$

(0.05)      (0.10)

$$t = \frac{0.82 - 1}{0.10} = -1.8$$

With  $n=20$  and  $t_{c,18,5\%} = 2.101$  (two-sided) we cannot reject the null hypothesis that  $\beta_1 = 1$  at the 5% level.

## Example one-sided test

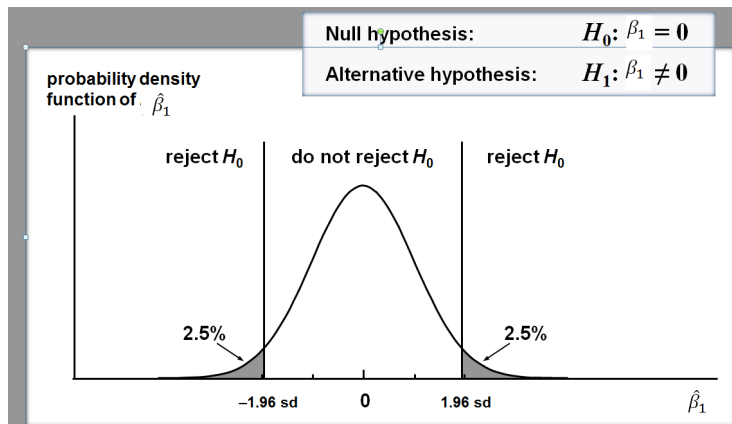
$$H_0 : \beta_2 = 1 \text{ vs } H_1 : \beta_1 < 1$$

The t-statistic is the same as long as the null is the same. But the critical t-value changes.

$$t_{c,18,5\%} = 1.734 \text{ one-sided test}$$

Now we can reject the null hypothesis and conclude that price inflation is significantly lower than wage inflation at the 5% level.

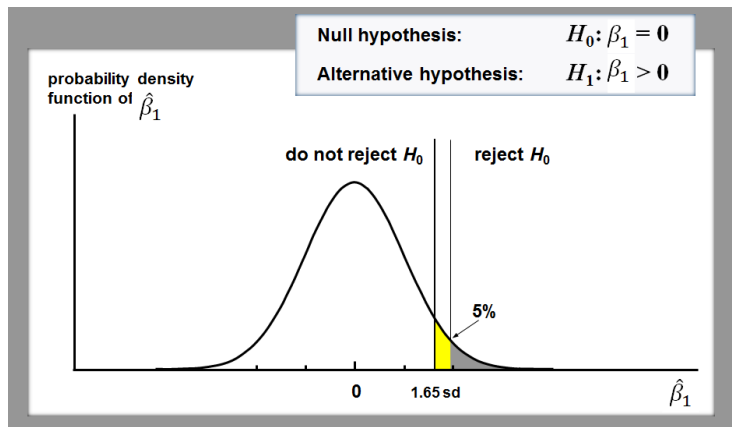
## One sided vs two sided



If you use a two-sided 5% significance test your estimate must be 1.96 standard deviations above or below 0 if you are to reject the null hypothesis.



## One sided vs two sided



If you can justify the use of a one-sided test, for example with  $H_0: \beta_1 > 0$ , the estimate has to be only 1.65 standard deviations above 0.

## Economic versus statistical significance

- The statistical significance of a variable is determined entirely by the size of the computed t-statistic.
- A coefficient can be statistically significant either because the coefficient is large, or because the standard error is small.
- With large samples parameters can be estimated very precisely which usually results in statistical significance.
- The economic significance is related to the size (and sign) of  $\hat{\beta}_1$ .
- Thus you should also discuss whether the coefficient is economically important (i.e. the magnitude of the coefficient)

# Example

```
1 . reg math10 sal
```

Source	SS	df	MS	Number of obs = 408		
Model	2562.57022	1	2562.57022	F( 1, 406) =	24.62	
Residual	42254.6103	406	104.075395	Prob > F =	0.0000	
Total	44817.1805	407	110.115923	R-squared =	0.0572	
				Adj R-squared =	0.0549	
				Root MSE =	10.202	

math10	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
sal	.4980309	.1003674	4.96	0.000	.3007264	.6953355
_cons	8.282175	3.228869	2.57	0.011	1.934787	14.62956

If you increase salary by \$1000 this is predicted to increase the percentage of students passing the math test by 0.5 percentage point. Is this an economically significant effect?

# Homoskedasticity

## Homoskedasticity assumption:

```
1 . reg ahe female
```

Source	SS	df	MS
Model	13091.0876	1	13091.0876
Residual	779560.368	7709	101.12341
Total	792651.456	7710	102.80823

```
Number of obs =      7711
F( 1, 7709) =     129.46
Prob > F       =     0.0000
R-squared      =     0.0165
Adj R-squared =     0.0164
Root MSE      =     10.056
```

ahe	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
female	-2.629912	.2311422	-11.38	0.000	-3.083013	-2.17681
_cons	20.11387	.1520326	132.30	0.000	19.81584	20.41189

## Heteroskedasticity robust:

```
1 . reg ahe female, robust
```

```
Linear regression
```

```
Number of obs =      7711
F( 1, 7709) =     134.80
Prob > F       =     0.0000
R-squared      =     0.0165
Root MSE      =     10.056
```

ahe	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
female	-2.629912	.2265122	-11.61	0.000	-3.073937	-2.185886
_cons	20.11387	.1614226	124.60	0.000	19.79744	20.4303

# Implication of heteroskedasticity

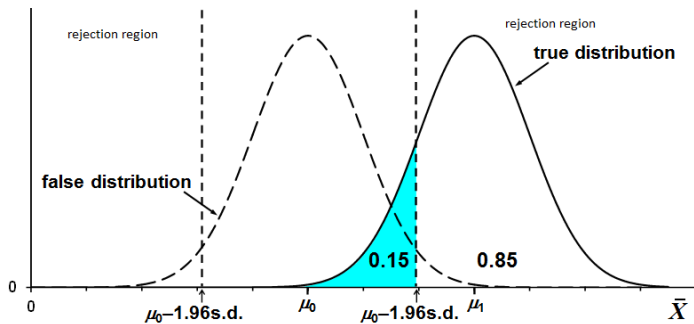
- If the regression errors are homoskedastic and normally distributed and if the homoskedasticity-only t-statistics is used, then critical values should be taken from the Student t distribution.
- In econometric applications the errors are rarely homoskedastic and normally distributed, but as long as  $n$  is large and we compute heteroskedasticity robust standard errors we can compute t-statistics and hence p-values and confidence intervals as normal.

## Note of caution:

- The test statistic: The t-value and hence the p-value and confidence interval is only as good as the underlying assumptions used to construct it.
- If any of the underlying assumptions are violated the test statistic is not reliable.
- Most often the violated assumption is the zero conditional mean assumption,  $X$  is often correlated with the error term.
- More about this in the next lecture when we talk about omitted variable bias.

# Power

A lot of emphasis is put on the significance of a test. It is important to remember that there is also another side to the story: power.



Power is the probability of rejecting the null hypothesis when it is false.