

ECON4150 - Introductory Econometrics

Lecture 12: Instrumental variables

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Stock and Watson Chapter 12

Lecture outline

- OLS assumptions and when they are violated
- Instrumental variable approach
 - 1 endogenous regressor & 1 instrument
 - IV assumptions:
 - instrument relevance
 - instrument exogeneity
 - 1 endogenous regressor, 1 instrument & control variables
 - 1 endogenous regressor & multiple instruments
 - multiple endogenous regressors & multiple instruments

Introduction

$$Y_i = \beta_0 + \beta_1 X_i + u_i$$

The 3 assumptions of an OLS regression model:

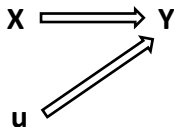
- 1 $E(u_i|X_i) = 0$
- 2 $(X_i, Y_i), i = 1, \dots, N$ are independently and identically distributed
- 3 Big outliers are unlikely.

Threats to internal validity (violation of 1st OLS assumption):

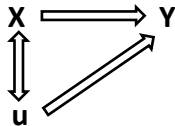
- Omitted variables
- Functional form misspecification
- Measurement error
- Sample selection
- Simultaneous causality

$$Y_i = \beta_0 + \beta_1 X_i + u_i$$

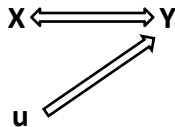
We **can** use OLS to obtain consistent estimate of the causal effect if



We **can't** use OLS to obtain consistent estimate of the causal effect if



and/or



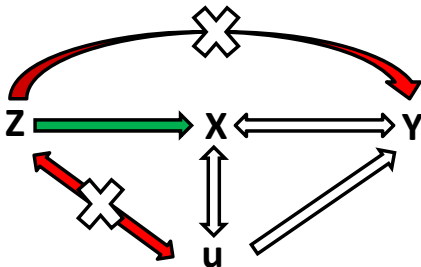
Instrumental variables: 1 endogenous regressor & 1 instrument

$$Y_i = \beta_0 + \beta_1 X_i + u_i$$

- Potential solution if $E[u_i|X_i] \neq 0$: use an instrumental variable (Z_i)
- We want to split X_i into two parts:
 - 1 part that is correlated with the error term (causing $E[u_i|x_i] \neq 0$)
 - 2 part that is uncorrelated with the error term
- If we can isolate the variation in X_i that is uncorrelated with u_i ...
- ...we can use this to obtain a consistent estimate of the causal effect of X_i on Y_i

Instrumental variables: 1 endogenous regressor & 1 instrument

- In order to isolate the variation in X_i that is uncorrelated with u_i we can use an **instrumental variable** Z_i with the following properties:
 - Instrument relevance:** Z_i is correlated with the endogenous regressor $Cov(Z_i, X_i) \neq 0$
 - Instrument exogeneity:** Z_i is uncorrelated with the error term $Cov(Z_i, u_i) = 0$ and has no direct effect on Y_i



Instrumental variables: 1 endogenous regressor & 1 instrument

We can extend the linear regression model

$$Y_i = \beta_0 + \beta_1 X_i + u_i \qquad X_i = \pi_0 + \pi_1 Z_i + v_i$$

We can estimate the causal effect of X_i on Y_i in two steps:

First stage: Regress X_i on Z_i & obtain predicted values $\hat{X}_i = \hat{\pi}_0 + \hat{\pi}_1 Z_i$

- If $Cov(Z_i, u_i) = 0$, \hat{X}_i contains variation in X_i that is uncorrelated with u_i

Second stage: Regress Y_i on \hat{X}_i to obtain the Two Stage Least Squares estimator $\hat{\beta}_{2SLS}$:

$$\hat{\beta}_{2SLS} = \frac{\sum_{i=1}^n (Y_i - \bar{Y}) (\hat{X}_i - \bar{\hat{X}})}{\sum_{i=1}^n (\hat{X}_i - \bar{\hat{X}})^2}$$

Application: estimating the returns to education

- Data from the NLS Young Men Cohort collected in 1976 on (among others) wages and years of education for 3010 men.
- Data are provided by Professor David Card, he used the data in his article "Using Geographic Variation in College Proximity to Estimate the Return to Schooling"

```
. regress ln_wage education, robust
```

Linear regression

```
Number of obs =      3010
F( 1, 3008) =    321.16
Prob > F      =    0.0000
R-squared     =    0.0987
Root MSE    =    .42139
```

ln_wage	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
education	.0520942	.0029069	17.92	0.000	.0463946	.0577939
_cons	5.570882	.0390935	142.50	0.000	5.49423	5.647535

- OLS estimate of the returns to education likely inconsistent due to omitted variables and measurement error.

Application: estimating the returns to education

- We want to isolate variation in years of education that is uncorrelated with the error term
- Card (1995) uses variation in college proximity as instrumental variable
- We have the following instrumental variable

near_college= 1 if individual grew up in area with a 4-year college
 0 if individual grew up in area without a 4-year college

Step 1: First stage regression

```
. regress education near_college, robust
```

Linear regression

```
Number of obs =      3010
F( 1, 3008) =      60.37
Prob > F       =      0.0000
R-squared      =      0.0208
Root MSE     =      2.6494
```

education	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
near_college	.829019	.1066941	7.77	0.000	.6198182	1.03822
_cons	12.69801	.0902199	140.75	0.000	12.52112	12.87491

Application: estimating the returns to education

Step 2: Obtain the predicted values and perform the second stage regression

```
1 . predict pr_education, xb
2 . regress ln_wage pr_education, robust
```

Linear regression

```
Number of obs =      3010
F( 1, 3008) =      83.79
Prob > F       =      0.0000
R-squared      =      0.0268
Root MSE      =      .43789
```

ln_wage	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
pr_education	.1880626	.0205454	9.15	0.000	.1477781	.2283472
_cons	3.767472	.2724927	13.83	0.000	3.233181	4.301763

Instrumental variables: 1 endogenous regressor & 1 instrument

Regression Y_i on \hat{X}_i gives the 2SLS estimator

$$\hat{\beta}_{2SLS} = \frac{\sum_{i=1}^n (Y_i - \bar{Y}) (\hat{X}_i - \bar{\hat{X}})}{\sum_{i=1}^n (\hat{X}_i - \bar{\hat{X}})^2}$$

If we substitute $\hat{X}_i - \bar{\hat{X}} = (\hat{\pi}_0 + \hat{\pi}_1 Z_i) - (\hat{\pi}_0 + \hat{\pi}_1 \bar{Z}) = \hat{\pi}_1 (Z_i - \bar{Z})$ we get

$$\hat{\beta}_{2SLS} = \frac{\sum_{i=1}^n (Y_i - \bar{Y}) \hat{\pi}_1 (Z_i - \bar{Z})}{\sum_{i=1}^n \hat{\pi}_1^2 (Z_i - \bar{Z})^2} = \frac{1}{\hat{\pi}_1} \times \frac{\sum_{i=1}^n (Y_i - \bar{Y}) (Z_i - \bar{Z})}{\sum_{i=1}^n (Z_i - \bar{Z})^2}$$

Since $\hat{\pi}_1$ is the first stage OLS estimator:

$$\hat{\beta}_{2SLS} = \frac{\sum_{i=1}^n (Z_i - \bar{Z})^2}{\sum_{i=1}^n (X_i - \bar{X}) (Z_i - \bar{Z})} \times \frac{\sum_{i=1}^n (Y_i - \bar{Y}) (Z_i - \bar{Z})}{\sum_{i=1}^n (Z_i - \bar{Z})^2}$$

Which gives the instrumental variable estimator

$$\hat{\beta}_{IV} = \frac{\sum_{i=1}^n (Y_i - \bar{Y}) (Z_i - \bar{Z})}{\sum_{i=1}^n (X_i - \bar{X}) (Z_i - \bar{Z})}$$

Application: estimating the returns to education

- We can obtain the 2SLS estimator in two steps as we have seen
- However the standard errors reported in the second stage regression are incorrect
- Stata does not recognize that it is a second stage of a two stage process, it fails to take into account the uncertainty in the first stage estimation.
- Instead obtain the 2SLS-estimator in 1 step:

```
. ivregress 2sls ln_wage (education=near_college), robust
```

```
Instrumental variables (2SLS) regression                Number of obs =          3010
                                                       Wald chi2(    1) =        51.78
                                                       Prob > chi2    =          0.0000
                                                       R-squared      =           .
                                                       Root MSE     =          .55667
```

ln_wage	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
education	.1880626	.0261339	7.20	0.000	.1368412	.2392841
_cons	3.767472	.3466268	10.87	0.000	3.088096	4.446848

```
Instrumented:  education
Instruments:   near_college
```

Instrumental variables: 1 endogenous regressor & 1 instrument

$$\hat{\beta}_{IV} = \frac{\sum_{i=1}^n (Y_i - \bar{Y}) (Z_i - \bar{Z})}{\sum_{i=1}^n (X_i - \bar{X}) (Z_i - \bar{Z})}$$

In large samples the IV-estimator converges to

$$plim(\hat{\beta}_{IV}) = \frac{Cov(Y_i, Z_i)}{Cov(X_i, Z_i)} = \frac{Cov(\beta_0 + \beta_1 X_i + u_i, Z_i)}{Cov(X_i, Z_i)} = \beta_1 + \frac{Cov(u_i, Z_i)}{Cov(X_i, Z_i)}$$

If the two IV-assumptions hold

- 1 **Instrument relevance:** $Cov(Z_i, X_i) \neq 0$
- 2 **Instrument exogeneity:** $Cov(Z_i, u_i) = 0$

The IV-estimator is consistent $plim(\hat{\beta}_{IV}) = \beta_1$, and is normally distributed in large samples

$$\hat{\beta}_{IV} \sim N \left(\beta_1, \frac{1}{n} \frac{Var [(Z_i - \mu_Z) u_i]}{[Cov(Z_i, X_i)]^2} \right)$$

Instrumental variables: 1 endogenous regressor & 1 instrument

The Instrumental Variables estimator is not unbiased

$$\begin{aligned}
 E[\hat{\beta}_{IV}] &= E\left[\frac{\sum_{i=1}^n (Y_i - \bar{Y})(Z_i - \bar{Z})}{\sum_{i=1}^n (X_i - \bar{X})(Z_i - \bar{Z})}\right] \\
 &= E\left[\frac{\sum_{i=1}^n ((\beta_0 + \beta_1 X_i + u_i) - (\beta_0 + \beta_1 \bar{X} + \bar{u}))(Z_i - \bar{Z})}{\sum_{i=1}^n (X_i - \bar{X})(Z_i - \bar{Z})}\right] \\
 &= E\left[\frac{\beta_1 \sum_{i=1}^n (X_i - \bar{X})(Z_i - \bar{Z}) + \sum_{i=1}^n (u_i - \bar{u})(Z_i - \bar{Z})}{\sum_{i=1}^n (X_i - \bar{X})(Z_i - \bar{Z})}\right] \\
 &= \beta_1 + E\left[\frac{\sum_{i=1}^n (u_i - \bar{u})(Z_i - \bar{Z})}{\sum_{i=1}^n (X_i - \bar{X})(Z_i - \bar{Z})}\right] = \beta_1 + E\left[\frac{\sum_{i=1}^n u_i (Z_i - \bar{Z})}{\sum_{i=1}^n (X_i - \bar{X})(Z_i - \bar{Z})}\right] \\
 &= \beta_1 + E_{X,Z}\left[\frac{\sum_{i=1}^n E[u_i | Z_i, X_i](Z_i - \bar{Z})}{\sum_{i=1}^n (X_i - \bar{X})(Z_i - \bar{Z})}\right] \\
 &\neq \beta_1
 \end{aligned}$$

Instrument exogeneity implies $E[u_i | Z_i] = 0$ but not $E[u_i | Z_i, X_i] = 0$ (this would mean that $E[u_i | X_i] = 0$ and we would not need an instrument!)

Instrumental variables: 1 endogenous regressor & 1 instrument

How can we know whether the IV assumptions hold?

1 Instrument relevance: $Cov(Z_i, X_i) \neq 0$

- We can check whether instrument relevance holds.
- Note that $\pi_1 = \frac{Cov(Z_i, X_i)}{Var(Z_i)}$
- We can therefore test $H_0 : \pi_1 = 0$ against $H_1 : \pi_1 \neq 0$

2 Instrument exogeneity: $Cov(Z_i, u_i) = 0$

- We can't check whether this assumption holds.
- We need to use economic theory, expert knowledge and intuition.

Instrument relevance & weak instruments

- Clearly, an irrelevant instrumental variable has problems, recall that

$$\hat{\beta}_{2SLS} \rightarrow \frac{\text{Cov}(Y_i, Z_i)}{\text{Cov}(X_i, Z_i)}$$

- In case of an irrelevant (but exogenous) instrumental variable both the denominator and numerator are 0.
- If instrument is not irrelevant but $\text{Cov}(X_i, Z_i)$ is close to zero
 - The sampling distribution of $\hat{\beta}_{2SLS}$ is not normal
 - $\hat{\beta}_{2SLS}$ can be severely biased, in the direction of the OLS estimator, even in relatively large samples!
- We should therefore always check whether an instrument is relevant enough.

Instrument relevance & weak instruments

- Let F_{first} be the F-statistic resulting from the test $H_0 : \pi_1 = 0$ against $H_1 : \pi_1 \neq 0$
- Staiger & Stock (Econometrica, 1997) show that in a simple model $\frac{1}{F_{first}}$ provides approximate estimate of finite sample bias of $\hat{\beta}_{2SLS}$ relative to $\hat{\beta}_{OLS}$
- Stock & Yogo (2005) argue that instruments are weak if the IV Bias is more than 10% of the OLS Bias.
- **Rule of thumb:** the F -statistic for (joint) significance of the instrument(s) in the first-stage should exceed 10.

Application: estimating the returns to education

Do the instrumental variable assumptions hold for college proximity as an instrument to estimate the returns to education?

1 Instrument relevance/weak instruments

education	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
near_college	.829019	.1066941	7.77	0.000	.6198182	1.03822
_cons	12.69801	.0902199	140.75	0.000	12.52112	12.87491

```
. test near_college
```

```
( 1) near_college = 0
```

```
      F( 1, 3008) =    60.37
      Prob > F =    0.0000
```

2 Instrument exogeneity:

- Is there a direct effect of living near a 4 year college on earnings?
- Is college proximity related to omitted variables that affect earnings?
 - What about area characteristics, such as living in a big city instead of a small village?

1 endogenous regressor, 1 instrument & control variables

- We can weaken the instrument exogeneity assumption by including area characteristics as control variables
- The Instrumental variables model is extended by including the control variables W_{1i}, \dots, W_{ri}

$$Y_i = \beta_0 + \beta_1 X_i + \delta_1 W_{1i} + \dots + \delta_r W_{ri} + u_i$$

$$X_i = \pi_0 + \pi_1 Z_i + \gamma_1 W_{1i} + \dots + \gamma_r W_{ri} + v_i$$

- The Instrument exogeneity condition is now conditional on the included regressors W_{1i}, \dots, W_{ri}

$$\text{Cov}(Z_i, u_i | W_{1i}, \dots, W_{ri}) = 0$$

- In the returns to education example we will include the following control variables:
 - age and age squared
 - *south* equals 1 if an individual lives in the southern part of the U.S.
 - *smsa* equals 1 if an individual lives in a Standard Metropolitan Statistical Area

Application: estimating the returns to education

Control variables must also be included in the first stage regression:

```
1 . regress education near_college age age2 south smsa, robust
```

Linear regression

Number of obs = 3010
 F(5, 3004) = 40.82
 Prob > F = 0.0000
 R-squared = 0.0710
 Root MSE = 2.5822

education	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
near_college	.3567396	.1117581	3.19	0.001	.1376095	.5758696
age	1.077846	.3044035	3.54	0.000	.4809854	1.674706
age2	-.0189181	.0052999	-3.57	0.000	-.0293099	-.0085264
south	-.8953645	.0987761	-9.06	0.000	-1.08904	-.7016888
smsa	.7962275	.1156382	6.89	0.000	.5694895	1.022965
_cons	-2.349802	4.329293	-0.54	0.587	-10.83848	6.138875

```
2 . test near_college
```

```
( 1) near_college = 0
```

```
F( 1, 3004) = 10.19  

  Prob > F = 0.0014
```

Don't use the overall F-statistic, this also tests whether the coefficients on the control variables equal zero!

Application: estimating the returns to education

IV estimates with control variables

```
. ivregress 2sls ln_wage (education=near_college) age age2 south smsa, robust
```

```
Instrumental variables (2SLS) regression
```

	Number of obs =	3010
	Wald chi2(5) =	757.69
	Prob > chi2 =	0.0000
	R-squared =	0.1510
	Root MSE =	.40884

ln_wage	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
education	.0954681	.0481396	1.98	0.047	.0011163	.1898199
age	.0815643	.0702011	1.16	0.245	-.0560274	.2191559
age2	-.0007088	.0012218	-0.58	0.562	-.0031034	.0016859
south	-.1277804	.0478661	-2.67	0.008	-.2215962	-.0339646
smsa	.1038856	.0472	2.20	0.028	.0113752	.1963959
_cons	3.246947	.7048721	4.61	0.000	1.865423	4.628471

```
Instrumented: education
Instruments: age age2 south smsa near_college
```

- Estimated return to an additional year of education is now 9.5%
- Do we believe that instrument exogeneity holds now that we have included control variables?

1 endogenous regressor, multiple instruments

- Instead of 1 instrument we can also use $M > 1$ instruments
- We could calculate M different IV-estimates of β
- Since any linear combination of the Z_{mi} is again a valid instrument:
 - combine the Z_{mi} to get a more efficient estimator of β_1

$$Y_i = \beta_0 + \beta_1 X_i + \delta_1 W_{1i} + \dots + \delta_r W_{ri} + u_i$$

$$X_i = \pi_0 + \pi_1 Z_{1i} + \dots + \pi_M Z_{Mi} + \gamma_1 W_{1i} + \dots + \gamma_r W_{ri} + v_i$$

- Instrumental variable assumptions:
 - 1 **Instrument relevance:** at least one of the instruments Z_{1i}, \dots, Z_{Mi} should have a nonzero coefficient in the population regression of X_i on Z_{1i}, \dots, Z_{Mi} .
 - 2 **Instrument exogeneity:**
 $Cov(Z_{1i}, u_i) = Cov(Z_{2i}, u_i) = \dots = Cov(Z_{Mi}, u_i) = 0$

Application: estimating the returns to education

- The data set contains two potential instruments for years of education:

near_2yrcollege= 1 if individual grew up in area with a 2-year college
 0 if individual grew up in area without a 2-year college

near_4yrcollege= 1 if individual grew up in area with a 4-year college
 0 if individual grew up in area without a 4-year college

- To check for instrument relevance we should estimate the first stage regression, including both instruments
- And use an F-test to test for the joint significance of the two instruments.

Application: estimating the returns to education

Linear regression

Number of obs = 3010
 F(6, 3003) = 34.03
 Prob > F = 0.0000
 R-squared = 0.0710
 Root MSE = 2.5827

education	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
near_4yrcollege	.3573365	.1121497	3.19	0.001	.1374385	.5772345
near_2yrcollege	-.0110908	.0976786	-0.11	0.910	-.2026145	.1804329
age	1.077147	.3045554	3.54	0.000	.4799884	1.674305
age2	-.0189051	.0053029	-3.57	0.000	-.0293028	-.0085074
south	-.8964387	.0991639	-9.04	0.000	-1.090875	-.7020027
smsa	.797801	.1167322	6.83	0.000	.5689179	1.026684
_cons	-2.336789	4.331927	-0.54	0.590	-10.83063	6.157055

```
2 . test near_4yrcollege=near_2yrcollege=0
```

```
( 1) near_4yrcollege - near_2yrcollege = 0
( 2) near_4yrcollege = 0
```

```
F( 2, 3003) = 5.09
Prob > F = 0.0062
```

- The first-stage F-statistic is well below 10, which indicates that we have weak instrument problems!
- It is better to drop the weakest instrument, *near_2yrcollege*, and use only 1 instrument *near_4yrcollege*

Overidentifying restrictions test (Sargan test, J-test)

- With more instruments than endogenous regressors we can test whether a subset of the instrument exogeneity conditions is valid.
- Suppose we have two instruments. Given our structural equation

$$Y_i = \beta_0 + \beta_1 X_i + \delta_1 W_{1i} + \dots + \delta_r W_{ri} + u_i$$

and assuming that $Cov(Z_{1i}, u_i) = 0$ we can test whether $Cov(Z_{2i}, u_i) = 0$ (or vice versa, but not both!)

- Intuition is as follows:
 - since $Cov(Z_{1i}, u_i) = 0 : \hat{\beta}_{2SLS}^{(Z_1)} \rightarrow \beta_1$
 - IF $Cov(Z_{2i}, u_i) = 0$ then also $\hat{\beta}_{2SLS}^{(Z_2)} \rightarrow \beta_1$
- Testing whether $Cov(Z_{2i}, u_i) = 0$ is equivalent to testing $\hat{\beta}_{2SLS}^{(Z_2)} = \hat{\beta}_{2SLS}^{(Z_1)}$

Overidentifying restrictions test (Sargan test, J-test)

We can implement the test as follows

- 1 Estimate $Y_i = \beta_0 + \beta_1 X_i + \delta_1 W_{1i} + \dots + \delta_r W_{ri} + u_i$ by 2SLS using Z_{1i} and Z_{2i} as instruments
- 2 Obtain the residuals $\hat{u}_i^{2SLS} = Y_i - \hat{\beta}_0 + \hat{\beta}_1 X_i + \hat{\delta}_1 W_{1i} + \dots + \hat{\delta}_r W_{ri}$
 - Note: use the true X_i and not the predicted value \hat{X}_i

- 3 Estimate the following regression

$$\hat{u}_i^{2SLS} = \eta_0 + \eta_1 \cdot Z_{1i} + \eta_2 \cdot Z_{2i} + \dots + \varphi_1 W_{1i} + \dots + \varphi_r W_{ri} + e_i$$

- 4 And obtain the F-statistic of the test

$$H_0 : \eta_1 = \eta_2 = 0 \quad \text{versus} \quad H_1 : \eta_1 \neq 0 \text{ and/or } \eta_2 \neq 0$$

- 5 Compute the J-test statistic

$$J = mF \sim \chi_q^2$$

where q is number of instruments minus number of endogenous regressors (in this case 1)

Application: estimating the returns to education

```
1 . ivregress 2sls ln_wage (education=near_4yrcollege near_2yrcollege) age age2 south smsa,
```

```
Instrumental variables (2SLS) regression
```

	Number of obs =	3010
	Wald chi2(5) =	766.83
	Prob > chi2 =	0.0000
	R-squared =	0.1609
	Root MSE =	.40646

ln_wage	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
education	.0927438	.0477741	1.94	0.052	-.0008916	.1863792
age	.0844422	.0696594	1.21	0.225	-.0520878	.2209722
age2	-.0007592	.0012123	-0.63	0.531	-.0031353	.0016169
south	-.1303678	.0475011	-2.74	0.006	-.2234683	-.0372672
smsa	.10638	.0468341	2.27	0.023	.0145869	.1981731
_cons	3.241778	.7006403	4.63	0.000	1.868548	4.615008

```
Instrumented: education
```

```
Instruments: age age2 south smsa near_4yrcollege near_2yrcollege
```

```
2 . predict residuals, resid
```

Application: estimating the returns to education

```
1 . regress residuals near_4yrcollege near_2yrcollege age age2 south smsa, robust
```

```
Linear regression                               Number of obs =      3010
                                                F( 6, 3003) =      0.42
                                                Prob > F       =     0.8684
                                                R-squared     =     0.0008
                                                Root MSE     =     .40676
```

residuals	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
near_4yrcollege	-.0003358	.0170653	-0.02	0.984	-.0337967	.0331252
near_2yrcollege	.0242942	.0154024	1.58	0.115	-.0059061	.0544946
age	.0015897	.0486995	0.03	0.974	-.093898	.0970775
age2	-.0000297	.0008437	-0.04	0.972	-.0016839	.0016245
south	.002501	.015634	0.16	0.873	-.0281535	.0331555
smsa	-.003772	.0174362	-0.22	0.829	-.0379601	.0304162
_cons	-.0297385	.6960319	-0.04	0.966	-1.394486	1.335009

```
2 . test near_4yrcollege=near_2yrcollege=0
```

```
( 1) near_4yrcollege - near_2yrcollege = 0
( 2) near_4yrcollege = 0
```

```
F( 2, 3003) =      1.24
Prob > F    =     0.2882
```

- $J = mF = 2 \cdot 1.24 = 2.48$
- $2.48 < 2.71$ (critical value of χ_1^2 at 10% significance level)
- So we do not reject the null hypothesis of instrument exogeneity.

Overidentifying restrictions test (Sargan test, J-test)

- Can we conclude that the two instruments satisfy instrument exogeneity? **NO!**
- Although the J-test seems a useful test there are 3 reasons to be very careful when using this test in practice
 - 1 When we don't reject the null hypothesis this does not mean that we can accept it!
 - 2 The power of the J-test can be low (probability of rejecting when H_0 does not hold)
 - 3 The J-test tests the joint hypothesis of instrument validity and correct functional form
 - 1 if the test rejects, the instruments might be valid but the functional form is wrong
 - 2 if the test rejects, the instruments might be valid but the effect of the regressor of interest is heterogeneous $\beta_{1i} \neq \beta_1$

The general IV regression model

- So far we considered the case with 1 endogenous variable, but we can extend the model to multiple endogenous variables

$$Y_i = \beta_0 + \beta_1 X_{1i} + \dots + \beta_K X_{Ki} + \delta_1 W_{1i} + \dots + \delta_r W_{ri} + u_i$$

$$\begin{aligned} X_{1i} &= \pi_0^1 + \pi_1^1 Z_{1i} + \dots + \pi_M^1 Z_{Mi} + \gamma_1^1 W_{1i} + \dots + \gamma_r^1 W_{ri} + v_i^1 \\ &\vdots \\ X_{Ki} &= \pi_0^K + \pi_1^K Z_{1i} + \dots + \pi_M^K Z_{Mi} + \gamma_1^K W_{1i} + \dots + \gamma_r^K W_{ri} + v_i^K \end{aligned}$$

- The general IV regression model has 4 types of variables
 - 1 The dependent variable Y_i
 - 2 K (possibly) endogenous regressors X_{1i}, \dots, X_{Ki}
 - 3 r control variables W_{1i}, \dots, W_{ri} (not the variables of interest)
 - 4 M instrumental variables Z_{1i}, \dots, Z_{Mi}

The general IV regression model

- When there are multiple endogenous regressors the 2SLS algorithm is similar except that each endogenous regressor requires its own first stage.
- For IV regression to be possible there should be at least as many instruments as endogenous regressors
- The model is said to be

Underidentified if $M < K$, we cannot estimate the model, the number of instruments is then smaller than the number of endogenous regressors

Exactly identified if $M = K$, the number of instruments equals the number of endogenous regressors

Overidentified if $M > K$, the number of instruments exceeds the number of endogenous regressors

The general IV regression model

Assumptions of the general IV-model

- 1 Instrument exogeneity:

$$\text{Cov}(Z_{1i}, u_i) = \text{Cov}(Z_{2i}, u_i) = \dots = \text{Cov}(Z_{Mi}, u_i) = 0$$

- 2 Instrument relevance:

- for each endogenous regressor X_{1i}, \dots, X_{Ki} , at least one of the instruments Z_{1i}, \dots, Z_{Mi} should have a nonzero coefficient in the population regression of the endogenous regressor on the instruments.

- The predicted values and the control variables $(\hat{X}_{1i}, \dots, \hat{X}_{Ki}, W_{1i}, \dots, W_{ri}, 1)$ should not be perfectly multicollinear.

- 3 $(X_{1i}, \dots, X_{Ki}, W_{1i}, \dots, W_{ri}, Z_{1i}, \dots, Z_{Mi}, Y_i)$ should be iid draws from their joint distribution.
- 4 Large outliers are unlikely: the X 's, W 's, Z 's and Y have finite fourth moments.

Application: estimating the returns to education

Summary of results using college proximity as instrument:

	OLS	1 IV without controls	1 IV with controls	2 IV's with controls
IV results, log(earnings) as dependent variable				
Education	0.052*** (0.003)	0.188*** (0.021)	0.095** (0.048)	0.093* (0.048)
First stage regression				
near 4yr college		0.829*** (0.107)	0.357*** (0.112)	0.357*** (0.112)
near 2yr college				-0.011 (0.098)
First stage F		60.37	10.19	5.09

* significant at 10%, ** significant at 5%, *** significant at 1%

- Is college proximity a valid instrument?

Application: estimating the returns to education

- Another possible instrument for education is compulsory schooling laws
- Between 1925 and 1970 there were quite some changes in the minimum school leaving age in the US
 - these changes varied between states
- Oreopoulos (AER,2006) uses variation in minimum school leaving age as instrument for years of schooling
- Main assumptions
 - Changes in minimum school leaving age uncorrelated with unobserved variables affecting education (such as ability)
 - No direct effect of changes in minimum school leaving age on wages
 - Minimum school leaving age has a nonzero impact of years of education

Estimating returns to education

- Oreopoulos estimates the following first stage and second stage equations:

$$Y_{ist} = \beta X_{ist} + \gamma_s + \gamma_t + V'_{ist}\theta + W'_{st}\lambda + \varepsilon_{ist}$$

$$X_{ist} = \pi Z_{st} + \delta_s + \delta_t + V'_{ist}\rho + W'_{st}\kappa + \mu_{ist}$$

- Y_{ist} is log wage of individual i living in state s in year t at age 14
- X_{ist} is years of schooling of individual i living in state s in year t at age 14
- Z_{st} is the minimum school leaving age in state s in year t
- γ_s and δ_s are state fixed effects, γ_t and δ_t are year fixed effects
- V'_{ist} are individual characteristics and W'_{st} are state characteristics

Estimating returns to education

Results from Oreopoulos (2006)

	OLS	First stage	IV
	ln(Earnings)	Education	ln(Earnings)
Years of education	0.078*** (0.0005)		0.142*** (0.012)
Minimum school leaving age		0.110*** (0.007)	

- First stage F-statistic: $F_{first} = t^2 = \left(\frac{0.110}{0.007}\right)^2 = 246.9$
- IV estimate almost twice as high as OLS estimate, not what we expect on basis of positive ability bias story
- Possible explanations:
 - downward bias in OLS due to measurement error
 - heterogeneity in the returns to education (IV estimates local average treatment effect)